

Last time: SGL: will continue today!

$$\begin{aligned}
 y &= Ax + w, \quad x \in \mathbb{R}^n, y \in \mathbb{R}^m \\
 w &\sim N(0, \sigma^2 I_m); \quad x \sim N(0, \Gamma) \\
 \Rightarrow y &\sim N(0, \Sigma_y); \quad \Sigma_y = \sigma^2 I_m + A\Gamma A^T \\
 \text{Choose } \Gamma &\text{ to max. log } p(y; \Gamma, \sigma^2) \\
 \Rightarrow \min_{\Gamma} L(\Gamma) &\stackrel{\text{def}}{=} -\log p(y; \Gamma) = \log \det \Sigma_y + y^T \Sigma_y^{-1} y \\
 -L(\Gamma) &\geq \mathbb{E}_{x|y; \Gamma, \sigma^2} \left\{ \log \frac{p(y, x; \Gamma, \sigma^2)}{p(x|y; \Gamma, \sigma^2)} \right\} \\
 &\quad \uparrow \text{ Jensen's inequality} \quad \uparrow \text{ any value of } \log \det \Gamma - \sigma^2 \mathbb{E} \{ \|x\|^2 \} \\
 \text{max of the RHS led to: } &\quad \Gamma \leftarrow x|y; \Gamma, \sigma^2 \\
 \text{max } - \left\{ \sum_{i=1}^n \log \tau_i + \tau_i^{-1} \mathbb{E}_{x|y; \Gamma, \sigma^2} \{ x_i^2 \} \right\} & \\
 \text{where } \mu &= \sigma^{-2} \Sigma_w^{-1} A^T y; \quad \Sigma_w = (\sigma^{-2} A A^T + \Gamma^{-1})^{-1} \\
 \Rightarrow \tau_i^{-1} &= \mu_i^2 + \Sigma_{w,ii}^{-1}
 \end{aligned}$$

Thus, recurse for $\max_{\Gamma} \log p(y; \Gamma, \sigma^2)$:
Init: $\Gamma^{(1)} = I$; $k=0$.

- Repeat:
- $$\begin{aligned}
 q_i(x_i) &= p(x_i | y; \Gamma^{(k)}, \sigma^2) \\
 &= N(x_i; \mu_i, \Sigma_{w,ii}) \\
 \mu &= \sigma^{-2} \Sigma_w^{-1} A^T y \\
 \Sigma_w &= (\sigma^{-2} A A^T + \Gamma^{(k-1)})^{-1}
 \end{aligned}$$
E-STEP.
 - Optimize: $y_i^{(k+1)} = \mu_i^2 + \Sigma_{w,ii}^{-1}$, $i=1,2,\dots,n$. M-STEP.
- Until $L(\Gamma) = \log \det \Sigma_y + y^T \Sigma_y^{-1} y$, $\Sigma_y = (\sigma^2 I_m + A\Gamma A^T)$ converges
- Output \hat{x}

Remark 1: $\Sigma_w = (\sigma^{-2} A A^T + \Gamma^{(k-1)})^{-1}$: inversion of an $n \times n$ matrix.

Make the inversion lemma:
 $(A + UC)^{-1} = A^{-1} - A^{-1} U (C + V A^{-1} U)^{-1} V A^{-1}$

We get

$$\Sigma_w^{-1} = \Gamma - \Gamma A^T (\sigma^2 I + A \Gamma A^T)^{-1} A \Gamma$$

Remark 2: $\sigma^2 \rightarrow 0$

$$\begin{aligned}
 \mu &= \sigma^{-2} \Sigma_w^{-1} A^T y \\
 &= \sigma^{-2} (\sigma^{-2} A A^T + \Gamma^{-1})^{-1} A^T y \\
 &= (A A^T + \sigma^2 \Gamma^{-1})^{-1} A^T y \\
 &= \Gamma^{\frac{1}{2}} (\Gamma^{\frac{1}{2}} A A^T \Gamma^{\frac{1}{2}} + \sigma^2 I)^{-1} \Gamma^{\frac{1}{2}} A^T y \\
 \text{Limit relation (see Wikipedia: pseudoinverse)} \\
 B^{\dagger} &= \lim_{\delta \rightarrow 0} (B^T B + \delta I)^{-1} B^T = \lim_{\delta \rightarrow 0} B^T (B B^T + \delta I)^{-1} \\
 \mu &= \Gamma^{\frac{1}{2}} (A A^T)^{\dagger} y \\
 \sigma^2 \rightarrow 0 \\
 \Sigma_w &= \Gamma - \Gamma A^T (\sigma^2 I + A \Gamma A^T)^{-1} A \Gamma \\
 &= \Gamma - \Gamma^{\frac{1}{2}} \Gamma^{\frac{1}{2}} A^T (\sigma^2 I + A \Gamma A^T)^{-1} \Gamma^{\frac{1}{2}} A \Gamma \\
 &= \Gamma - \Gamma^{\frac{1}{2}} (A \Gamma A^T)^{\dagger} A \Gamma
 \end{aligned}$$

Then, let $\sigma^2=0$, and let x_0 be the maximally sparse soln. to $y = Ax$. Assume $|\text{supp}(x_0)| < m$, and $\|x_0\|_0 < \infty$. Let τ_0 (or Γ_0) be a vec. of prior variances s.t. $\tau_0 = \Gamma_0^{\frac{1}{2}} (A \Gamma_0 A^T)^{\dagger} y$ where the min nonzero entry of Γ_0 is $\delta > 0$, and $|\text{supp}(\tau_0)| = |\text{supp}(x_0)|$. Then, the global min. of $L = \log \det (\sigma^2 I + A \Gamma A^T) + y^T (\sigma^2 I + A \Gamma A^T)^{-1} y$ occurs at $\Gamma = \Gamma_0$ and $\sigma^2 = 0$.

Proof: We will s.t. $L \rightarrow -\infty$ at $\sigma^2=0$, $\Gamma = \Gamma_0$.
 $\log \det \Sigma_y \rightarrow -\infty$, $y^T \Sigma_y^{-1} y \rightarrow \text{bounded}$.

Now, when $\sigma^2=0$, $\Gamma = \Gamma_0$,
 $\text{rank}(\Sigma_y) = \text{rank}(A \Gamma_0 A^T) \leq \text{rank}(\Gamma_0) < m$.
 $\Rightarrow \det \Sigma_y = 0$, $\log \det \Sigma_y \rightarrow -\infty$.
With $\sigma^2=0$, $\Gamma = \Gamma_0$, $y^T (A \Gamma_0 A^T)^{-1} y = x_0^T \Gamma_0^{-1} x_0 \leq \frac{1}{\delta} \|x_0\|_0^2$
 $\Rightarrow L \rightarrow -\infty$ at $\sigma^2=0$, $\Gamma = \Gamma_0$. □

Facts: (See D. Wipf & B. Recht's paper)

- If A has the unique representation property (any subset of m cols of A are LI), then, the global min. of L can only occur at a degenerate γ , that produces a degenerate sparse soln., i.e., $|\text{supp}(x)| < m$. (x is obtained from the posterior mean computed using the γ that attains the global min. of L)
- Under mild assumptions on A , there will exist no other degenerate sparse solns.
- Can s.t. every local min. of L occurs at a sparse soln., i.e., $\|x\|_0 < m$, even in the noisy case.

Updating σ^2

In Step 2, must max. w.r.t σ^2

$$\mathbb{E}_{x|y; \Gamma^{(k)}, \sigma^2} \left\{ \log \frac{p(y, x; \Gamma^{(k)}, \sigma^2)}{p(x|y; \Gamma^{(k)}, \sigma^2)} \right\}$$

Isolating the terms that dep. only on σ^2 ,

$$\mathbb{E}_{x|y; \Gamma^{(k)}, \sigma^2} \left\{ \log p(y|x; \sigma^2) \right\} \\
 = \mathbb{E}_{\Theta} \left\{ -\frac{\|y - Ax\|_2^2}{2\sigma^2} - \frac{m}{2} \log \sigma^2 \right\}$$

Diff. w.r.t. σ^2 , set = 0:

$$\begin{aligned}
 \sigma^2 &= \frac{1}{m} \mathbb{E}_{\Theta} \left\{ \|y - Ax\|_2^2 \right\} \\
 &= \frac{1}{m} \mathbb{E}_{\Theta} \left\{ \|y - A\mu\|_2^2 + \text{tr} (A^T A (\mu - x)(\mu - x)^T) \right\} \\
 &= \frac{1}{m} \mathbb{E}_{\Theta} \left\{ \|y - A\mu\|_2^2 + \text{tr} (A^T A (\mu - x)(\mu - x)^T) \right\} \\
 &= \frac{1}{m} \left\{ \|y - A\mu\|_2^2 + \text{tr} (A^T A \Sigma_w) \right\}
 \end{aligned}$$

$$\begin{aligned} \Sigma_{\omega} &= (\sigma^2 \lambda^2 A^T A + \Gamma^{-1})^{-1} \\ \Rightarrow (\Sigma_{\omega}^{-1} - \Gamma^{-1}) &= \sigma^2 \lambda^2 A^T A \\ \Rightarrow A^T A &= \frac{\sigma^2 \lambda^2}{\sigma^2} (\Sigma_{\omega}^{-1} - \Gamma^{-1}) \\ &= \frac{1}{m} \{ \|y - A\mu\|_2^2 + \sigma^2 \lambda^2 \text{tr}((\Sigma_{\omega}^{-1} - \Gamma^{-1}) \Sigma_{\omega}) \} \\ &= \frac{1}{m} \{ \|y - A\mu\|_2^2 + \sigma^2 \lambda^2 \text{tr}(\mathbb{I} - \Gamma^{-1} \Sigma_{\omega}) \} \\ \sigma^2 &= \frac{1}{m} \{ \|y - A\mu\|_2^2 + \sigma^2 \lambda^2 \sum_{i=1}^M (1 - \gamma_i^{\lambda^2} \Sigma_{\omega, i}) \} \end{aligned}$$

Q. Why does SBL lead to sparse solns?

Suppose we place a "hyper prior" on γ :

Inverse gamma distⁿ:

$$p_{\gamma}(y) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} \gamma^{-\alpha-1} e^{-\beta/\gamma}, \quad \gamma > 0, \alpha, \beta > 0.$$

$$p_x(x) = \prod_{i=1}^M \int_0^{\infty} \frac{1}{\sqrt{\pi z_i}} e^{-\frac{x_i^2}{z_i}} \frac{\beta^{\alpha}}{\Gamma(\alpha)} \frac{1}{z_i^{\alpha+1}} e^{-\beta/z_i} dz_i$$

$$\approx \prod_{i=1}^M \int_0^{\infty} \frac{1}{\sqrt{\pi z_i}} e^{-\left(\frac{x_i^2}{z_i} + \beta\right) \frac{1}{z_i}} dz_i$$

$$\propto \prod_{i=1}^M \left(\frac{x_i^2}{z_i} + \beta \right)^{-\left(\alpha + \frac{1}{2}\right)}$$

As $\alpha, \beta \rightarrow 0$, this becomes

$$\propto \prod_{i=1}^M \frac{1}{|x_i|}. \quad \text{Clearly encourages sparsity.} \quad |\alpha|, \beta < 1$$