

An Introduction to Approximate Inference with Applications to Massive MIMO Communication Systems

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- Evaluation of posterior distribution of latent variables given the observed data is a central task in statistical inference
- Computationally intractable due to the integrals over the latent variables
- Approximation techniques
 - Stochastic: Sampling methods like MCMC
 - Computationally demanding
 - Deterministic: Variational inference, Expectation propagation
 - Approximate solutions but faster
- Variational inference
- Expectation propagation

- Consider a Bayesian model
 - Observations $\mathbf{Z} = \{z_1, \dots, z_N\}$
 - Latent variables $\mathbf{X} = \{x_1, \dots, x_N\}$
- Goal is to find an approximate posterior distribution $p(\mathbf{X}|\mathbf{Z})$ and the model evidence $p(\mathbf{Z})$
 - Exact computations are computationally intractable

$$\ln p(\mathbf{Z}) = \mathcal{L}(q) + \text{KL}(q\|p)$$

where

$$\mathcal{L}(q) \triangleq \int q(\mathbf{X}) \ln \left\{ \frac{p(\mathbf{Z}, \mathbf{X})}{q(\mathbf{X})} \right\} d\mathbf{X}$$
$$\text{KL}(q\|p) = - \int q(\mathbf{X}) \ln \left\{ \frac{p(\mathbf{X}|\mathbf{Z})}{q(\mathbf{X})} \right\} d\mathbf{X} \geq 0$$

- Need to find a distribution $q(\mathbf{X})$ which will maximize the evidence lower bound (ELBO) $\mathcal{L}(q)$
 - Maximum occurs when $q(\mathbf{X}) = p(\mathbf{X}|\mathbf{Z}) \implies$ computational intractability
 - Impose structure on q and minimize the KL divergence

- Factorized distributions for q
 - Approximation framework developed in physics called *mean field theory*

$$q(\mathbf{X}) = \prod_{i=1}^M q_i(\mathbf{X}_i)$$

$$\begin{aligned} \mathcal{L}(q) &= \int \prod_i q_i \left\{ \ln p(\mathbf{Z}, \mathbf{X}) - \sum_i \ln q_i \right\} d\mathbf{X} \\ &= \int q_j \int \ln p(\mathbf{Z}, \mathbf{X}) \prod_{i \neq j} q_i d\mathbf{X}_i d\mathbf{X}_j - \int q_j \ln q_j d\mathbf{X}_j - \sum_{i \neq j} \int q_i \ln q_i d\mathbf{X}_i \\ &= \int q_j \ln \tilde{p}(\mathbf{Z}, \mathbf{X}_j) d\mathbf{X}_j - \int q_j \ln q_j d\mathbf{X}_j + \text{const.} \\ &= -\text{KL}(q_j \| \tilde{p}(\mathbf{Z}, \mathbf{X}_j)) + \text{const.} \end{aligned}$$

where

$$\ln \tilde{p}(\mathbf{Z}, \mathbf{X}_j) \triangleq \mathbb{E}_{i \neq j} [\ln p(\mathbf{Z}, \mathbf{X})] + \text{const.}$$

- To maximize $\mathcal{L}(q)$, need to minimize the KL divergence in (1)
 - Minimum occurs when $q_j(\mathbf{X}_j) = \tilde{p}(\mathbf{Z}, \mathbf{X}_j)$

- Optimal q_j is given by

$$\begin{aligned} q_j^*(\mathbf{X}_j) &= \text{const} \times \exp(\mathbb{E}_{i \neq j} [\ln p(\mathbf{Z}, \mathbf{X})]) \\ &= \frac{\exp(\mathbb{E}_{i \neq j} [\ln p(\mathbf{Z}, \mathbf{X})])}{\int \exp(\mathbb{E}_{i \neq j} [\ln p(\mathbf{Z}, \mathbf{X})]) d\mathbf{X}_j} \end{aligned}$$

- Can obtain the parameters of the distribution $q_j^*(\mathbf{Z}_j)$ by inspection
- Fix $q_{i \neq j}$ and obtain the parameters of q_j and iterate for all j

- The Exponential family

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp \left\{ \boldsymbol{\eta}^T \mathbf{u}(\mathbf{x}) \right\}$$

where $\boldsymbol{\eta}$ is the natural parameter vector

- Maximum likelihood and sufficient statistics:

$$-\nabla \ln g(\boldsymbol{\eta}_{ML}) = \frac{1}{N} \sum_{n=1}^N \mathbf{u}(\mathbf{x}_n)$$

- Differentiating $\int p(\mathbf{x}|\boldsymbol{\eta})d\mathbf{x} = 1$ gives

$$-\nabla \ln g(\boldsymbol{\eta}) = \mathbb{E}\{\mathbf{u}(\mathbf{x})\}$$

- To obtain an approximate posterior by minimizing $\text{KL}(p \parallel q)$ instead of $\text{KL}(q \parallel p)$

$$\text{KL}(p \parallel q) = -\ln g(\boldsymbol{\eta}) - \boldsymbol{\eta}^T \mathbb{E}_{p(\mathbf{z})} [\mathbf{u}(\mathbf{z})] + \text{const}$$

- Minimum occurs when

$$\mathbb{E}_{q(\mathbf{z})} [\mathbf{u}(\mathbf{z})] = \mathbb{E}_{p(\mathbf{z})} [\mathbf{u}(\mathbf{z})]$$

which is called moment matching

- Joint distribution of data \mathcal{D} and hidden variables $\boldsymbol{\theta}$

$$p(\mathcal{D}, \boldsymbol{\theta}) = \prod_i f_i(\boldsymbol{\theta})$$

- Approximation to the posterior distribution (from an exponential family)

$$q(\boldsymbol{\theta}) = \frac{1}{Z} \prod_i \tilde{f}_i(\boldsymbol{\theta})$$

- EP proceeds in a similar fashion as variational Bayes
 - Initialize the factors and cycle through them refining one at a time

- To obtain $\tilde{f}_j(\theta)$ by ensuring that the product

$$q^{\text{new}}(\theta) \propto \tilde{f}_j(\theta) \prod_{i \neq j} \tilde{f}_i(\theta)$$

is as close as possible to $f_j(\theta) \prod_{i \neq j} \tilde{f}_i(\theta)$

- Remove the factor $\tilde{f}_j(\theta)$ from the current approximation

$$q^{\setminus j}(\theta) = \frac{q(\theta)}{\tilde{f}_j(\theta)}$$

- Combine it with $f_j(\theta)$ to give a distribution $\frac{1}{Z_j} f_j(\theta) q^{\setminus j}(\theta)$
- Find $\tilde{f}_j(\theta)$ by minimizing

$$\text{KL} \left(\frac{f_j(\theta) q^{\setminus j}(\theta)}{Z_j} \parallel q^{\text{new}}(\theta) \right)$$

- Since q is from an exponential family, the minimization can be done using moment matching

$$\tilde{f}_j(\theta) = K \frac{q^{\text{new}}(\theta)}{q^{\setminus j}(\theta)}$$

- Disadvantage: No convergence guarantee in general. For exponential family, it may converge to a stationary point

Joint Channel Estimation and Data Detection for Uplink Massive MIMO Systems with Low Resolution ADCs

Motivation

- Large number of antennas results in an increased hardware complexity and circuit power consumption.
 - Power consumption of ADCs increases exponentially with respect to the number of bits per sample
- Non-linearity due to quantization necessitates novel signal processing algorithms
- Most of the existing literature on low resolution ADCs assume perfect CSI at the receiver, which is not realistic.

Goal

- Joint channel estimation and data detection in the uplink of a coded massive MIMO wireless communication system with low resolution ADCs

Contributions

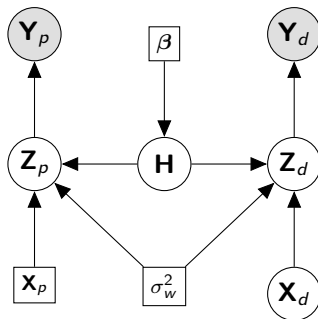
- Joint channel estimation and data detection as a statistical inference problem in a probabilistic graphical model.
- MIMO channel, data symbols and unquantized outputs as latent variables whose marginal distributions are inferred
- Variational Bayesian inference to compute the marginal distributions of the channel and data

Received Signal

$$\mathbf{Y}_p = \mathcal{Q}(\mathbf{Z}_p) = \mathcal{Q}(\mathbf{H}\mathbf{X}_p + \mathbf{W}_p), \quad (1)$$

$$\mathbf{Y}_d = \mathcal{Q}(\mathbf{Z}_d) = \mathcal{Q}(\mathbf{H}\mathbf{X}_d + \mathbf{W}_d), \quad (2)$$

Bayesian Network Model



Quantized Variational Bayesian Joint Channel Estimation and Data Detection

- Fully factorized approximation of the posterior distribution is shown below:

$$\begin{aligned} p(\mathbf{Z}_p, \mathbf{Z}_d, \mathbf{X}_d, \mathbf{H} | \mathbf{Y}_p, \mathbf{Y}_d, \mathbf{X}_p; \beta, \sigma_w^2) \\ \approx q(\mathbf{Z}_p) q(\mathbf{Z}_d) q(\mathbf{X}_d) q(\mathbf{H}), \end{aligned} \quad (3)$$

where

$$q(\mathbf{H}) = \prod_{n=1}^{N_{RX}} \prod_{k=1}^K q(h_{nk}), \quad q(\mathbf{X}_d) = \prod_{k=1}^K \prod_{t=1}^{\tau_d} q(x_{d,kt}) \quad (4)$$

$$q(\mathbf{Z}_d) = \prod_{n=1}^{N_{RX}} \prod_{t=1}^{\tau_d} q(z_{d,nt}), \quad q(\mathbf{Z}_p) = \prod_{n=1}^{N_{RX}} \prod_{t=1}^{\tau_p} q(z_{p,nt}). \quad (5)$$

- Conditional probability distributions of the observations and the latent variables are given as

$$p(\mathbf{Z}_p | \mathbf{X}_p, \mathbf{H}; \sigma_w^2) \propto \exp\left(-\frac{1}{\sigma_w^2} \sum_{t=1}^{\tau_p} \|\mathbf{z}_{p,t} - \mathbf{H}\mathbf{x}_{p,t}\|^2\right), \quad (6)$$

$$p(\mathbf{Z}_d | \mathbf{X}_d, \mathbf{H}; \sigma_w^2) \propto \exp\left(-\frac{1}{\sigma_w^2} \sum_{t=1}^{\tau_d} \|\mathbf{z}_{d,t} - \mathbf{H}\mathbf{x}_{d,t}\|^2\right), \quad (7)$$

$$p(\mathbf{H} | \boldsymbol{\beta}) \propto \exp\left(-\sum_{k=1}^K \frac{1}{\beta_k} \|\mathbf{h}_k\|^2\right), \quad (8)$$

$$p(\mathbf{Y}_d | \mathbf{Z}_d) = \mathbb{1}(\mathbf{Z}_d \in [\mathbf{Z}_d^{(lo)}, \mathbf{Z}_d^{(hi)}]), \quad (9)$$

$$p(\mathbf{Y}_p | \mathbf{Z}_p) = \mathbb{1}(\mathbf{Z}_p \in [\mathbf{Z}_p^{(lo)}, \mathbf{Z}_p^{(hi)}]), \quad (10)$$

- Computation of $q(h_{nk})$:

$$\ln q(h_{nk}) \propto \langle \ln p(\mathbf{Z}_p | \mathbf{X}_p, \mathbf{H}; \sigma_w^2) + \ln p(\mathbf{Z}_d | \mathbf{X}_d, \mathbf{H}; \sigma_w^2) + \ln p(\mathbf{H} | \boldsymbol{\beta}) \rangle, \quad (11)$$

$$\begin{aligned} &\propto -\frac{1}{\sigma_w^2} \left\{ \left(\sum_{t=1}^{\tau_p} |x_{p,kt}|^2 + \sum_{t=1}^{\tau_d} \langle |x_{d,kt}|^2 \rangle + \frac{\sigma_w^2}{\beta_k} \right) |h_{nk}|^2 \right. \\ &\quad \left. - 2\Re \left(\left(\sum_{t=1}^{\tau_p} \left[\langle z_{p,nt} \rangle^* x_{p,kt} - x_{p,kt} \sum_{\substack{k'=1 \\ k' \neq k}}^K x_{p,k't}^* \langle h_{nk'} \rangle^* \right] \right) \right. \right. \\ &\quad \left. \left. + \sum_{t=1}^{\tau_d} \left[\langle z_{d,nt} \rangle^* \langle x_{d,kt} \rangle - \langle x_{d,kt} \rangle \sum_{\substack{k'=1 \\ k' \neq k}}^K \langle x_{d,k't} \rangle^* \langle h_{nk'} \rangle^* \right] \right) h_{nk} \right\} \quad (12) \end{aligned}$$

- The structure of (12) resembles that of a complex normal distribution

- Computation of $q(x_{d,kt})$:

$$q(x_{d,kt} = s) = \frac{\exp\left(-\frac{1}{\sigma_w^2} f(s)\right)}{\sum_{s'} \exp\left(-\frac{1}{\sigma_w^2} f(s')\right)}, \quad (13)$$

where s belongs to a symbol from the M -QAM, and

$$f(s) = \langle \|\mathbf{h}_k\|^2 \rangle |s|^2 + 2\Re \left[\left(\sum_{\substack{k'=1 \\ k' \neq k}}^K \langle \mathbf{h}_{k'} \rangle^H \langle \mathbf{h}_k \rangle \langle x_{d,k't} \rangle^* - \langle \mathbf{z}_{d,t} \rangle^H \langle \mathbf{h}_k \rangle \right) s \right].$$

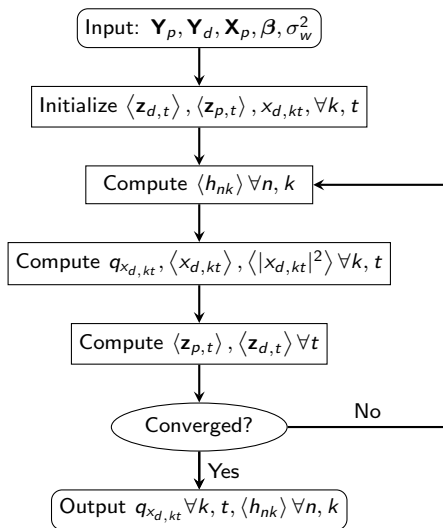
- Follows a Boltzmann distribution

- Computation of $q(\mathbf{z}_{d,t})$ and $q(\mathbf{z}_{p,t})$:

$$\ln q(\mathbf{z}_{d,t}) \propto \left\langle \ln \mathbb{1}(\mathbf{z}_{d,t} \in [\mathbf{z}_{d,t}^{(lo)}, \mathbf{z}_{d,t}^{(hi)}]) - \frac{1}{\sigma_w^2} \|\mathbf{z}_{d,t} - \mathbf{H}\mathbf{x}_{d,t}\|^2 \right\rangle \quad (14)$$

- The structure in (14) is that of a truncated complex normal distribution with mean given below.

$$\langle \mathbf{z}_{d,t} \rangle = \boldsymbol{\mu}_{\mathbf{z}_{d,t}} + \frac{\phi\left(\frac{\mathbf{z}_{d,t}^{(lo)} - \boldsymbol{\mu}_{\mathbf{z}_{d,t}}}{\sigma_w/\sqrt{2}}\right) - \phi\left(\frac{\mathbf{z}_{d,t}^{(hi)} - \boldsymbol{\mu}_{\mathbf{z}_{d,t}}}{\sigma_w/\sqrt{2}}\right)}{\Phi\left(\frac{\mathbf{z}_{d,t}^{(hi)} - \boldsymbol{\mu}_{\mathbf{z}_{d,t}}}{\sigma_w/\sqrt{2}}\right) - \Phi\left(\frac{\mathbf{z}_{d,t}^{(lo)} - \boldsymbol{\mu}_{\mathbf{z}_{d,t}}}{\sigma_w/\sqrt{2}}\right)} \frac{\sigma_w}{\sqrt{2}},$$



THANK YOU!