

Sufficient conditions for sparse support recovery in M-SBL

Saurabh Khanna,

Signal Processing for Communication, ECE, IISc

Joint sparse signal recovery

MMV model: $\mathbf{y}_j = \Phi \mathbf{x}_j + \mathbf{w}_j \quad j = 1 \text{ to } L$

$\mathbf{x}_j \in \mathbb{R}^n$ are unknown k -sparse vectors following JSM-2 signal model

- ▶ \mathbf{x}_j have a common nonzero support
- ▶ Nonzero entries are uncorrelated

$\Phi \in \mathbb{R}^{m \times n}$ is the measurement matrix with $m \ll n$

$\mathbf{y}_j \in \mathbb{R}^m$ are the noisy linear measurements

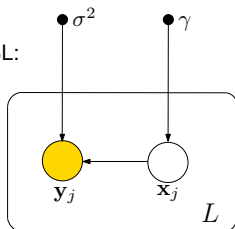
$\mathbf{w}_j \sim \mathcal{N}(0, \sigma^2 \mathbf{I}_m)$

Want to recover joint sparse vectors $\mathbf{x}_1, \mathbf{x}_2 \dots \mathbf{x}_L$ from noisy linear measurements $\mathbf{y}_1, \mathbf{y}_2 \dots \mathbf{y}_L$.

Examples of Joint sparse signal recovery algorithms: M-OMP, M-FOCUSS, Row-LASSO and M-SBL

M-SBL algorithm

Graphical model for MSBL:



$$\mathbf{x}_j \sim \mathcal{N}(0, \Gamma)$$

$$\Gamma = \text{diag}(\gamma)$$

$$\mathbf{y}_j \sim \mathcal{N}(0, \sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)$$

Use of common γ induces joint sparsity in $\mathbf{x}_1, \mathbf{x}_2 \dots \mathbf{x}_L$

For fixed γ , we have $p(\hat{\mathbf{x}}_j | \mathbf{y}_j; \gamma) \sim \mathcal{N}(\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)$, where

$$\boldsymbol{\Sigma} = \Gamma^{-1} + \frac{\Phi^T \Phi}{\sigma^2}, \quad \boldsymbol{\mu}_j = \frac{1}{\sigma^2} \boldsymbol{\Sigma}^{-1} \Phi^T \mathbf{y}_j$$

γ is found by solving Type-II ML estimation problem: $\hat{\gamma} = \arg \max_{\gamma \in \mathbb{R}_+^n} \log p(\mathbf{Y}; \gamma)$

M-SBL algorithm

MSBL cost function:

$$\begin{aligned}\log p(\mathbf{Y}; \boldsymbol{\gamma}) &= \log(\mathbf{y}_1, \mathbf{y}_2 \dots \mathbf{y}_L; \boldsymbol{\gamma}) \\ &= \sum_{j=1}^L \log p(\mathbf{y}_j; \boldsymbol{\gamma}) \\ &= \sum_{j=1}^L \log \mathcal{N}(\mathbf{y}_j; 0, \sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T) \\ &= -\frac{L}{2} \log 2\pi - \frac{L}{2} \log |\sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T| - \frac{1}{2} \sum_{j=1}^L \text{Tr} \left((\sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T)^{-1} \mathbf{y}_j \mathbf{y}_j^T \right) \\ &= -\frac{L}{2} \log 2\pi - \frac{L}{2} \log |\sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T| - \frac{1}{2} \text{Tr} \left((\sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T)^{-1} \mathbf{Y} \mathbf{Y}^T \right)\end{aligned}$$

M-SBL cost - interesting interpretation

MSBL cost function:

$$\log p(\mathbf{Y}; \gamma) = -\frac{L}{2} \log 2\pi - \frac{L}{2} \log |\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T| - \frac{1}{2} \text{Tr} \left((\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)^{-1} \mathbf{Y} \mathbf{Y}^T \right)$$

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Bregman matrix divergence with respect to $\phi(\cdot) = -\log|\cdot|$, is defined as:

$$\mathcal{D}(\mathbf{X}, \mathbf{Y}) = \text{trace}(\mathbf{X} \mathbf{Y}^{-1}) - \log |\mathbf{X} \mathbf{Y}^{-1}| - N$$

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MSBL cost can be interpreted as a matrix divergence:

$$\log p(\mathbf{Y}; \gamma) = -\frac{L}{2} \mathcal{D}_\phi \left(\frac{1}{L} \mathbf{Y} \mathbf{Y}^T, \sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T \right) - \underbrace{m - \frac{L}{2} \log 2\pi + \frac{L}{2} \log \left| \frac{1}{L} \mathbf{Y} \mathbf{Y}^T \right|}_{\text{constant}}$$

Thus, MSBL is trying to minimize the matrix divergence between

- (i) $\frac{1}{L} \mathbf{Y} \mathbf{Y}^T$ (empirical covariance matrix)
- (ii) $\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T$ (parameterized covariance matrix)

M-SBL cost - interesting interpretation

MSBL is trying to do “energy matching” or “second moment matching”

$$\log p(\mathbf{Y}|\boldsymbol{\gamma}) \propto \text{const} + \mathcal{D}_{-\log|\cdot|} \left(\frac{1}{L} \mathbf{Y}\mathbf{Y}^T, \sigma^2 \mathbf{I} + \boldsymbol{\Phi}\boldsymbol{\Gamma}\boldsymbol{\Phi}^T \right)$$

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[P. Pal and P.P. Vaidyanathan, '15] proposed the Cov-MMV problem:

$$\min_{\boldsymbol{\gamma} \in \mathbb{R}_+^n} \|\boldsymbol{\gamma}\|_1 \quad \text{subject to} \quad \frac{1}{L} \mathbf{Y}\mathbf{Y}^T = \sigma^2 \mathbf{I} + \boldsymbol{\Phi}\boldsymbol{\Gamma}\boldsymbol{\Phi}^T$$

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Cov-MMV can be rewritten in LASSO form as:

$$\min_{\boldsymbol{\gamma} \in \mathbb{R}_+^n} \|\boldsymbol{\gamma}\|_1 \quad \text{subject to} \quad (\boldsymbol{\Phi} \odot \boldsymbol{\Phi})\boldsymbol{\gamma} = \text{vec} \left(\frac{1}{L} \mathbf{Y}\mathbf{Y}^T - \sigma^2 \mathbf{I} \right)$$

MSBL vs. Cov-MMV

Concave $\log \|\cdot\|_1$ penalty is a better regularizer compared to convex $\|\cdot\|_1$ penalty

A better regularization of the feasible set translates to fewer measurements for guaranteeing a unique sparse solution.

We seek necessary and sufficient conditions for sparse support recovery for the M-SBL algorithm.

Support recovery in M-SBL

Constrained MSBL problem:

$$\hat{\gamma} = \arg \max_{\gamma \in \Theta_K} \mathcal{L}(\mathbf{Y}; \gamma)$$

$$\mathcal{L}(\mathbf{Y}; \gamma) = -\frac{L}{2} \log |\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T| - \frac{1}{2} \text{Tr} ((\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)^{-1} \mathbf{Y} \mathbf{Y}^T)$$

$$\Theta_K = \{\gamma \in \mathbb{R}_+^n : \|\gamma\|_0 \leq K, \|\gamma\|_\infty \leq \gamma_{\max}\}$$

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Problem setup:

Let $\mathbf{x}_1, \mathbf{x}_2 \dots \mathbf{x}_L$ be generated according to the true distribution $\mathcal{N}(0, \Gamma_0)$. Let $\mathcal{S}_0 = \text{supp}(\gamma_0)$ and, $|\mathcal{S}_0| \leq K$.

Question:

What is the (m, L) measurement rate region which guarantees that the solution of constrained MSBL problem, denoted by $\hat{\gamma}$ satisfies $\text{supp}(\hat{\gamma}) = \text{supp}(\gamma_0)$ with arbitrarily high probability.

Support recovery in M-SBL

Constrained MSBL:

$$\hat{\gamma} = \arg \max_{\gamma \in \Theta_K} \mathcal{L}(\mathbf{Y}; \gamma)$$

$$\mathcal{L}(\mathbf{Y}; \gamma) = \log p(\mathbf{Y}; \gamma)$$

$$\Theta_K = \{\gamma \in \mathbb{R}_+^n : \|\gamma\|_0 \leq K, \|\gamma\|_\infty \leq \gamma_{\max}\}$$

Success event \mathcal{A} :

$$\begin{aligned} \mathcal{A} &\triangleq \{ \text{supp}(\hat{\gamma}) = \text{supp}(\gamma_0) = \mathcal{S}_0 \} \\ &= \left\{ \sup_{\gamma \in \Theta_{\mathcal{S}_0}} \mathcal{L}(\mathbf{Y}, \gamma) \geq \sup_{\gamma \in \Theta_{\mathcal{S}}} \mathcal{L}(\mathbf{Y}, \gamma), \quad \forall \mathcal{S} \neq \mathcal{S}_0 \right\} \end{aligned}$$

Support recovery in M-SBL

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Question:

Identify the measurement rate region (m, L) such that $\mathbb{P}(\mathcal{A}) \geq 1 - \delta$.

Support recovery in M-SBL

We define **expected log likelihood**, $\bar{\mathcal{L}}(\mathbf{Y}, \gamma) = \mathbb{E}\mathcal{L}(\mathbf{Y}, \gamma)$

Expectation is w.r.t $\mathbf{Y} \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2 \mathbf{I} + \Phi \Gamma_0 \Phi^T)$

Then, we have

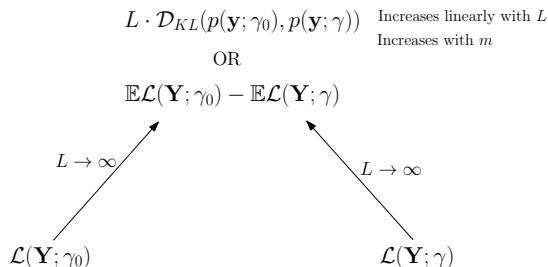
$$\begin{aligned}\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) &= \mathbb{E} \log p(\mathbf{Y}; \gamma_0) \\ \bar{\mathcal{L}}(\mathbf{Y}, \gamma) &= \mathbb{E} \log p(\mathbf{Y}; \gamma) \\ \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma) &= \mathbb{E}_{\mathbf{Y} \sim p(\mathbf{Y}; \gamma_0)} \log \frac{p(\mathbf{Y}; \gamma_0)}{p(\mathbf{Y}; \gamma)} \\ &= L \cdot \underbrace{\mathcal{D}_{\text{KL}}(p(\mathbf{y}; \gamma_0), p(\mathbf{y}; \gamma))}_{\text{non-negative !}}\end{aligned}$$

Tradeoff between m and L

$$\text{We want... } \mathbb{P} \left\{ \sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) \geq \sup_{\gamma \in \Theta_S} \mathcal{L}(\mathbf{Y}, \gamma), \quad \forall S \neq S_0 \right\} \geq 1 - \delta.$$

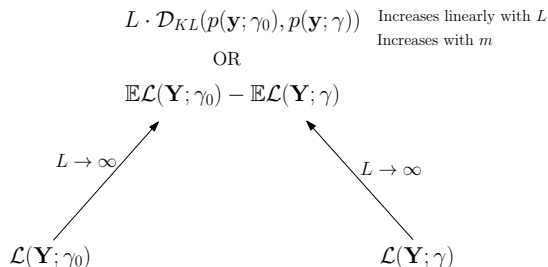
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$$\left[\sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) - \sup_{\gamma \in \Theta_S} \mathcal{L}(\mathbf{Y}, \gamma) \right] = L \cdot \mathcal{D}_{KL}(p(\mathbf{y}; \gamma_0), p(\mathbf{y}; \gamma)) - (\text{conc. terms})$$

$$\begin{aligned} \sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) &\geq \mathcal{L}(\mathbf{Y}, \gamma_0) = \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - (\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)) \\ &\geq \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - |\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)| \\ &\geq \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \sup_{\gamma \in \Theta_K} |\bar{\mathcal{L}}(\mathbf{Y}, \gamma) - \mathcal{L}(\mathbf{Y}, \gamma)| \end{aligned} \quad (1)$$

$$\begin{aligned}
\sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) &\geq \mathcal{L}(\mathbf{Y}, \gamma_0) = \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - (\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)) \\
&\geq \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - |\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)| \\
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\end{aligned}$$

$$\begin{aligned}
\sup_{\gamma \in \Theta_S} \mathcal{L}(\mathbf{Y}, \gamma) &= \sup_{\gamma \in \Theta_S} [\bar{\mathcal{L}}(\mathbf{Y}, \gamma) + (\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma))] \\
&\leq \sup_{\gamma \in \Theta_S} \bar{\mathcal{L}}(\mathbf{Y}, \gamma) + \sup_{\gamma \in \Theta_S} |\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma)| \\
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\end{aligned}$$

$$\begin{aligned}
\sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) &\geq \mathcal{L}(\mathbf{Y}, \gamma_0) = \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - (\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)) \\
&\geq \bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - |\bar{\mathcal{L}}(\mathbf{Y}, \gamma_0) - \mathcal{L}(\mathbf{Y}, \gamma_0)| \\
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\end{aligned}$$

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\sup_{\gamma \in \Theta_S} \mathcal{L}(\mathbf{Y}, \gamma) &= \sup_{\gamma \in \Theta_S} [\bar{\mathcal{L}}(\mathbf{Y}, \gamma) + (\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma))] \\
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&\leq \sup_{\gamma \in \Theta_S} \bar{\mathcal{L}}(\mathbf{Y}, \gamma) + \sup_{\gamma \in \Theta_K} |\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma)| \tag{2}
\end{aligned}$$

Subtracting (2) from (1), we get:

$$\begin{aligned}
&\left[\sup_{\gamma \in \Theta_{S_0}} \mathcal{L}(\mathbf{Y}, \gamma) - \sup_{\gamma \in \Theta_S} \mathcal{L}(\mathbf{Y}, \gamma) \right] \\
&\geq \inf_{\gamma \in \Theta_S} [L \cdot \mathcal{D}_{\text{KL}}(p(\mathbf{y}; \gamma_0), p(\mathbf{y}; \gamma))] - 2 \sup_{\gamma \in \Theta_K} |\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma)|
\end{aligned}$$

Lower bounding the KL divergence

KL divergence between multivariate Gaussians:

$$\mathcal{D}_{\text{KL}}(\mathcal{N}(0, \Sigma_1), \mathcal{N}(0, \Sigma_2)) \triangleq \frac{1}{2} \log \frac{|\Sigma_2|}{|\Sigma_1|} + \frac{1}{2} \text{trace}(\Sigma_2^{-1} \Sigma_1) - \frac{m}{2}.$$

- ▶ The ratio terms $\frac{|\Sigma_2|}{|\Sigma_1|}$ and $\text{trace}(\Sigma_2^{-1} \Sigma_1)$ can be difficult to handle.
- ▶ Norm based terms like $\|\Sigma_2 - \Sigma_1\|$ are preferred.

α -Renyi divergence lower bounds KL divergence for $\alpha \in (0, 1)$.

$$\mathcal{D}_\alpha(\mathcal{N}(0, \Sigma_1), \mathcal{N}(0, \Sigma_2)) = \frac{1}{2(1-\alpha)} \log \frac{|\Sigma_1|^{1-\alpha} |\Sigma_2|^\alpha}{|(1-\alpha)\Sigma_1 + \alpha\Sigma_2|}.$$

- ▶ $\mathcal{D}_{1/2}$ can be lower bounded in terms of $\|\Sigma_2 - \Sigma_1\|$ like difference terms.

Lower bound for α -Renyi divergence

Theorem

For $p_1 \sim \mathcal{N}(0, \Sigma_1)$ and $p_2 \sim \mathcal{N}(0, \Sigma_2)$, the α -Renyi divergence can be lower bounded as

$$D_{1/2}(p_1, p_2) \geq \frac{1}{2} \alpha(1 - \alpha) m^* \|\Sigma_2 - \Sigma_1\|_F^2$$

where m^* is the strong convexity constant of $-\log|\cdot|$ function and, is evaluated as:

$$m^* = \sup_{\Sigma \in \Theta} \frac{1}{(\lambda_{\max}(\Sigma))^2}$$

where $\Theta = \{t\Sigma_1 + (1 - t)\Sigma_2, t \in [0, 1]\}$.

Upper bound for $|\mathcal{L}(\mathbf{Y}; \gamma) - \bar{\mathcal{L}}(\mathbf{Y}; \gamma)|$

$$\begin{aligned} |\mathcal{L}(\mathbf{Y}; \gamma) - \bar{\mathcal{L}}(\mathbf{Y}; \gamma)| &= |\log p(\mathbf{Y}; \gamma) - \mathbb{E} \log p(\mathbf{Y}; \gamma)| \\ &= \frac{1}{2} \left| \text{trace} \left((\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)^{-1} (\mathbf{Y}^T - \mathbb{E} \mathbf{Y} \mathbf{Y}^T) \right) \right| \\ &= \frac{L}{2} \left| \text{trace} \left((\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)^{-1} (\hat{\mathbf{R}}_{\mathbf{y}} - \mathbf{R}_{\mathbf{y}}) \right) \right| \\ &\leq \frac{L}{2} \|(\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T)^{-1}\|_{\text{tr}} \cdot \|\hat{\mathbf{R}}_{\mathbf{y}} - \mathbf{R}_{\mathbf{y}}\|_2 \\ &\leq \frac{Lm}{2\sigma^2} \|\hat{\mathbf{R}}_{\mathbf{y}} - \mathbf{R}_{\mathbf{y}}\|_2 \\ &\leq \frac{Lm\epsilon}{2\sigma^2} \|\mathbf{R}_{\mathbf{y}}\|_2 \quad [\text{Vershyn, non-asymptotic RMT, '11}] \end{aligned}$$

Last inequality holds with probability at least $1 - \delta$, provided $L \geq \frac{C}{\epsilon^2} \log \frac{2}{\delta}$.

By eliminating ϵ , we get the following probabilistic bound:

$$\mathbb{P} \left(|\mathcal{L}(\mathbf{Y}; \gamma) - \bar{\mathcal{L}}(\mathbf{Y}; \gamma)| \leq \frac{1}{2\sigma^2} \sqrt{Lm^2 C \log \left(\frac{2}{\delta} \right)} \|(\sigma^2 \mathbf{I} + \Phi \Gamma_0 \Phi^T)\|_2 \right) \geq 1 - \delta.$$

Sufficient conditions

Recall, that for \mathcal{A} to hold, we need (m, L) such that

$$\inf_{\gamma \in \Theta_S} [L \cdot \mathcal{D}_{\text{KL}}(p(\mathbf{y}; \gamma_0), p(\mathbf{y}; \gamma))] \geq 2 \sup_{\gamma \in \Theta_K} |\mathcal{L}(\mathbf{Y}, \gamma) - \bar{\mathcal{L}}(\mathbf{Y}, \gamma)|$$

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Or equivalently,

$$\frac{L}{8} \left(\frac{\|\Sigma - \Sigma_0\|_F}{\|\Sigma(t)\|_2} \right)^2 \geq \frac{1}{\sigma^2} \sqrt{Lm^2 C \log \frac{2}{\delta}} \cdot \|\sigma^2 \mathbf{I} + \Phi \Gamma_0 \Phi^T\|_2$$

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Or equivalently,

$$L \geq \frac{64Cm^2}{\sigma^4} \cdot \log \frac{2}{\delta} \cdot \|\sigma^2 \mathbf{I} + \Phi \Gamma_0 \Phi^T\|_2^2 \cdot \left(\frac{\|\Sigma(t)\|_2}{\|\Sigma - \Sigma_0\|_F} \right)^4$$

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For $P(\mathcal{A}) \geq 1 - \delta$, we need (m, L) such that

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Sufficient conditions

For $P(\mathcal{A}) \geq 1 - \delta$, we need (m, L) such that

$$L \geq 64Cm^2 \log \frac{2}{\delta} \cdot \left(1 + \frac{\gamma_{\max}}{\sigma^2} (1 + \delta_K)\right)^2 \cdot \left(\frac{\sigma^2 + \gamma_{\max}(1 + \delta_{2K})}{\|(\Phi \odot \Phi)(\gamma - \gamma_0)\|_2}\right)^4$$

Or equivalently,

$$L \geq 64Cm^2 \log \frac{2}{\delta} \cdot \left(1 + \frac{\gamma_{\max}}{\sigma^2} (1 + \delta_K)\right)^2 \cdot \left(\frac{\sigma^2 + \gamma_{\max}(1 + \delta_{2K})}{\gamma_{\min} \sqrt{1 - \delta_{2K}(\Phi \odot \Phi)}}\right)^4$$

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Or equivalently,

$$L \geq 64Cm^2 \log \frac{2}{\delta} \cdot \left(1 + \frac{\gamma_{\max}}{\sigma^2}\right)^6 \left(\frac{\sigma^2}{\gamma_{\min}}\right)^4 \cdot \frac{(1 + \delta_K)^2 (1 + \delta_{2K})^4}{(1 - \delta_{2K}(\Phi \odot \Phi))^2}$$

Sparse support recovery in MSBL - sufficient conditions

Theorem

Let $\mathbf{x}_1, \mathbf{x}_2 \dots \mathbf{x}_L$ be generated according to the true distribution $\mathcal{N}(0, \Gamma_0)$. Let $\mathcal{S}_0 = \text{supp}(\gamma_0)$ and, $|\mathcal{S}_0| \leq K$. Let $\hat{\gamma}$ be the global maximizer of $\log p(\mathbf{Y}; \gamma)$ and, $\hat{\gamma} \in \Theta_K$, where $\Theta_K = \{\gamma \in \mathbb{R}_+^n : \|\gamma\|_0 \leq K, \|\gamma\|_\infty \leq \gamma_{\max}\}$. Then, the following holds.

$$\mathbb{P}(\text{supp}(\hat{\gamma}) = \mathcal{S}_0) \geq 1 - 2e^{-\eta m L}$$

where $\eta(m) = \left[64Cm^3 \cdot \left(1 + \frac{\gamma_{\max}}{\sigma^2}\right)^6 \left(\frac{\sigma^2}{\gamma_{\min}}\right)^4 \cdot \frac{(1+\delta_K)^2(1+\delta_{2K})^4}{(1-\delta_{2K}(\Phi \odot \Phi))^2} \right]^{-1}$.