

# The Role of MVUE and CRB in Composite Hypothesis Test

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# Outline

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# Introduction

- The relation between composite binary hypothesis test and the estimation of the unknown parameters is explored
- If an UMP test exists, it is identical to comparing the MVUE (Minimum variance Unbiased Estimator) for the unknown parameter to a threshold
- The relation between CRB (Cramer Rao Bound) in the estimation theory and the UMP performance bound in detection theory is analyzed
- If an UMP test does not exist, using a good estimator of the unknown parameter as a test statistic can improve the detection performance

## CRB

- If the pdf  $p(\mathbf{x}; \theta)$  satisfies the following regularity condition

$$E \left[ \frac{\ln(p(\mathbf{x}; \theta))}{\partial \theta} \right] = 0, \quad \text{for all } \theta$$

- Then, the variance of any unbiased estimator  $\hat{\theta}$  must satisfy

$$\text{var} \hat{\theta} \geq \frac{1}{E \left[ \frac{\partial^2 \ln(p(\mathbf{x}; \theta))}{\partial \theta^2} \right]}$$

- Further an unbiased estimate may be found that attains the bound for all  $\theta$  iff

$$\frac{\partial \ln(p(\mathbf{x}; \theta))}{\partial \theta} = I(\theta)(g(\mathbf{x}) - \theta) \quad (1)$$

- $\hat{\theta} = g(\mathbf{x})$  is the MVUE

- The decision rule divides the observation space into two partitions  $\Gamma_0$  and  $\Gamma_1$
- It can be shown that the decision rule for a binary hypothesis test can be stated as a comparison of a continuous test statistic with a threshold
- The observation space  $\chi$  is a normal space and  $\Gamma_0$  and  $\Gamma_1$  are connected sets

## Theorem

**Urysohn lemma:** *Let  $\chi$  be a normal space; Let  $A$  and  $B$  be disjoint closed subsets of  $\chi$ . Let  $[a, b]$  be a closed interval in the real line. Then, there exists a continuous map  $f : \chi \rightarrow [a, b]$ , such that  $f(x) = a$  for  $x \in A$  and  $f(x) = b$  for  $x \in B$*

## Theorem

Consider a decision rule  $d(\mathbf{x}) = 1$  if  $\mathbf{x} \in \Gamma_1$  and  $d(\mathbf{x}) = 0$  if  $x \in \Gamma_0$  in which  $x \in \chi$  and  $\Gamma_1 \cap \Gamma_0 = \emptyset$ ,  $\Gamma_1 \cup \Gamma_0 = \chi$ . If  $\Gamma_1$  and  $\Gamma_0$  are connected, then there exists a continuous statistic  $g(\mathbf{x})$  and a threshold  $\gamma$  such that

$$g : \chi \rightarrow \mathcal{R} \quad (2)$$

$$d(\mathbf{x}) = \begin{cases} 1, & g(\mathbf{x}) \geq \gamma \\ 0, & g(\mathbf{x}) < \gamma \end{cases} \quad (3)$$

- There exists an infinite number of functions  $g(\mathbf{x})$  that satisfy (2)

- Consider a one-sided hypothesis testing problem, i.e.,

$$H_0 : \mathbf{x} \sim f_x(\mathbf{x}; \theta), \theta < \theta_b$$

$$H_1 : \mathbf{x} \sim f_x(\mathbf{x}; \theta), \theta > \theta_b$$

- where  $\theta_b$  is known.

## Theorem

*Consider a one-sided hypothesis test  $H_0 : \mathbf{x} \sim f_x(\mathbf{x}; \theta), \theta < \theta_0$  against  $H_1 : \mathbf{x} \sim f_x(\mathbf{x}; \theta), \theta > \theta_b$ . Let  $\Omega_0 = \{\theta : \theta < \theta_b\}$ ,  $\Omega_1 = \{\theta : \theta > \theta_b\}$ ; and  $\Omega = \Omega_0 \cup \Omega_1$ . If the UMP test exists, then it can be stated as comparing the MVUE statistic of  $\theta \in \Omega$  with a threshold which is set to satisfy the probability of false alarm,  $P_{fa}$*

# Outline of the Proof

- The optimal N-P test is obtained from the likelihood ratio

$$LR(\mathbf{x}) = \frac{f_x(\mathbf{x}; \theta_1)}{f_x(\mathbf{x}; \theta_0)} > \gamma$$

- $\theta_1$  is  $\theta \in \Omega_1$  and  $\theta_0$  is  $\theta \in \Omega_0$
- The log-likelihood ratio is given by

$$\ln LR(\mathbf{x}) = \ln(f_x(\mathbf{x}; \theta_1)) - \ln(f_x(\mathbf{x}; \theta_0)) > \ln(\gamma)$$

- From the CRB theorem:

$$\frac{\partial \ln(f_x(\mathbf{x}; \theta))}{\partial \theta} = I(\theta)[g(\mathbf{x}) - \theta] \quad (4)$$



## Outline of the Proof

- $g(\mathbf{x})$  is the MVUE statistic;  $I(\theta)$  is the Fisher information matrix

$$\ln(f_x(\mathbf{x}; \theta)) = \int I(\theta)[g(\mathbf{x}) - \theta]d\theta + C(\mathbf{x}) \quad (5)$$

- $C(\mathbf{x})$  is a function of only  $\mathbf{x}$

$$\begin{aligned} \ln(LR(\mathbf{x})) &= \int I(\theta_1)[g(\mathbf{x}) - \theta_1]d\theta_1 \\ &= \int I(\theta_0)[g(\mathbf{x}) - \theta_0]d\theta_0 \end{aligned}$$

- Log-likelihood is increasing in

$$\left[ \int I(\theta_1)d\theta_1 - \int I(\theta_0)d\theta_0 \right] g(\mathbf{x}) \quad (6)$$

## Outline of the proof

- The Fisher information is positive, i.e.,  $I(\theta) > 0$ ,  $I(\theta)$  is an increasing function of  $\theta$
- From the one-sided hypothesis test,  $\theta_1 > \theta_0$ , therefore

$$\left[ \int I(\theta_1) d\theta_1 - \int I(\theta_0) d\theta_0 \right] > 0 \quad (7)$$

- Hence  $\left[ \int I(\theta_1) d\theta_1 - \int I(\theta_0) d\theta_0 \right] g(\mathbf{x})$  is increasing in  $g(\mathbf{x})$
- The UMP statistic rejects  $H_0$  if  $g(\mathbf{x}) > \gamma$ , where  $\gamma$  satisfies the  $P_{fa}$  constraint

# Vector Parameter

- $\Theta = [\theta_1, \theta_2, \dots, \theta_N]^T$
- From CRB theorem, for MVUE:

$$\ln(f_x(\mathbf{x}, \Theta)) = \int (I(\Theta)[g(\mathbf{x}) - \Theta])^H d\Theta \quad (8)$$

- The log-likelihood ratio is increasing in

$$g(\mathbf{x})^H \left[ \int I(\Theta_1)^H d\Theta_1 - \int I(\Theta_0)^H d\Theta_0 \right] \quad (9)$$

- UMP for the vector parameter is given by the linear combination of the elements of MVUE

# Vector Parameter

- The co-efficients in the linear combination do not simplify into constants
- In general, the co-efficients may be function of unknown parameters
- By properly choosing the co-efficients of the linear combination of the MVUE, UMP performance bound can be achieved
- In summary, if a good estimator for the function  $g(\mathbf{x})$  is chosen (reaching CRB), then UMP performance can be achieved

# Example-I

- Consider the following hypothesis testing problem

$$H_0 : \mathbf{x} = \mathbf{n}$$

$$H_1 : \mathbf{x} = \theta \mathbf{s} + \mathbf{n}$$

- $\mathbf{s}$  is the known signal with  $N$  samples and  $\theta > 0$
- Under  $H_0$ ,  $\mathbf{x} \sim N(0, C)$  and under  $H_1$ ,  $\mathbf{x} \sim N(\theta \mathbf{s}, C)$
- In the WGN case,  $C = \sigma^2 I$ .
- From the likelihood ratio, the test statistic for detection

$$T(\mathbf{x}) = \mathbf{x}^T C^{-1} \mathbf{s} > \gamma \quad (10)$$

# Example-I

- $\gamma$  is the detection threshold. This is an UMP test
- Generalized matched filter
- The MVUE for  $\theta$  is

$$\theta_{MVUE} = \frac{\mathbf{x}^T C^{-1} \mathbf{s}}{\mathbf{s}^T C^{-1} \mathbf{s}}$$

- The denominator is known positive constant;  $C$  is positive definite
- The UMP detection statistic is nothing but MVUE compared to a threshold

## Example-II

- Consider the hypothesis testing problem

$$H_0 : x_i = n_i$$

$$H_1 : x_i = x_0 + n_i$$

- The pdf of  $n_i$  is Cauchy distributed and  $x_0 > 0$
- No UMP exists and MVUE of center parameter also does not exist
- An unbiased estimation which tends to meet CRB as the number of observation increases

$$I_k(\mathbf{x}) = \frac{y_0}{\pi} \int_{-\infty}^{\infty} x_0^k \prod_{i=1}^N \frac{1}{y_0^2 + (x_i - x_0)^2} dx_0$$

- $0 \leq k \leq 2N - 1$ . Unbiased estimator of  $x_0$  is  $\hat{x}_0 = \frac{I_1(\mathbf{x})}{I_0(\mathbf{x})}$