

Lecture 13

(1)

Agenda: * Entropic compression review

- * Connection b/w compression and probability estimation/assignment
- * Add- α estimator and its Bayesian interpretation

A) Entropic compression review (See notes of IT)

Consider an iid sequence X_1, \dots, X_n with common distribution P .

If this distribution P is known,

- we can use Huffman codes to achieve avg. length

$$\mathbb{E}[L] \leq H(P) + 1,$$

even for $n=1$.

- in fact, arithmetic coding allows us to attain codewords with length $l(x) = \lceil -\log Q^{(n)}(x) \rceil$, $\forall x \in \mathcal{X}^n$.

* The advantage here is that we do not require the knowledge of $Q^{(n)}(x)$ at the outset, but only requires $Q(x_i | x^{(-i)})$ when encoding x_i , $1 \leq i \leq n$.

When P is unknown, and we used $Q^{(n)}$ to assign the lengths to arithmetic encoder, the average length achieved satisfies

$$\begin{aligned}(1) \quad \mathbb{E}[L] &\leq \mathbb{E}_{P^n} \left[\log \frac{1}{Q^{(n)}(X^n)} \right] + 1 \\ &= D(P^n || Q^{(n)}) + nH(P) + 1.\end{aligned}$$

On the other hand, for any prefix-free code (or even uniquely decodable code), the lengths $\{l(\underline{x}), \underline{x} \in \mathcal{X}^n\}$ satisfy Kraft's inequality (2)

$$k_n = \sum_{\underline{x}} 2^{-l(\underline{x})} \leq 1.$$

Consider $P_e(\underline{x}) = \frac{2^{-l(\underline{x})}}{k_n}$. Then,

$$\begin{aligned}\mathbb{E}[L(X^n)] &= \mathbb{E}\left[-\log k_n + \log \frac{1}{P_e(\underline{x})}\right] \\ &= -\log k_n + D(P^n \| P_e) + nH(P)\end{aligned}$$

$$(2) \Rightarrow \mathbb{E}[L(X^n)] - nH(P) \geq D(P^n \| P_e).$$

(1) and (2) yield an interesting duality b/w compression and probability assignment: the compression problem is equivalent (upto 1 bit) to assigning probabilities to sequences $\underline{x}^n \in \mathcal{X}^n$.

B Regret formulation

We now focus on the probability assignment problem, where the cost $D(P^n \| Q^{(n)})$ shows naturally.

Regret (instead of reward)

An individual sequence formulation:

$$l(\underline{x}, \theta) = \log \frac{1}{Q(\underline{x})} \equiv \text{log-loss function} \quad (3)$$

Σ = class of experts

Instead of seeking the absolute optimal scheme, we seek schemes that compete with the class Σ of experts.

Σ_0 = memoryless experts = $\{P^n, P \in \mathcal{P}_k\}$.

Then, the "regret" for using Q is given by

$$\begin{aligned} r(\underline{x}, Q) &= \log \frac{1}{Q(\underline{x})} - \min_{P \in \mathcal{P}_k} \log \frac{1}{P^n(\underline{x})} \\ &= \max_{P \in \mathcal{P}_k} \log \frac{P^n(\underline{x})}{Q(\underline{x})} \end{aligned}$$

The worst-case regret is given by

$$r_{k,n}(Q) = \max_{\underline{x} \in \mathcal{X}^n} r(\underline{x}, Q)$$

and the minimax regret is given by

$$r(k, n) = \min_{Q \in \mathcal{P}_{[k]^n}} r_{k,n}(Q)$$

Average regret formulation

$$\begin{aligned} \bar{r}_{k,n}(Q) &= \mathbb{E}_{P^n} \left[\log \frac{P(\underline{x}^n)}{Q(\underline{x}^n)} \right] \\ &= D(P^n \| Q) \end{aligned}$$

$$\bar{r}(k, n) = \min_{Q \in \mathcal{P}_{[k]^n}} \bar{r}_{k,n}(Q) \leq r(k, n)$$

In fact, we will see

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$$g_2(k, n) = \bar{g}_2(k, n) = \frac{k-1}{2} \log n + O_k(1).$$

C Bayesian Estimators: Uniform prior

$$\pi \equiv \text{unif}(P_k).$$

$Q \sim \pi$, denote by Q_π the overall measure.

$$\text{Then, } Q_\pi(X_{e+1} = x | X^l = x^l)$$

$$= \frac{Q_\pi(X^l = x^l, X_{e+1} = x)}{Q_\pi(X^l = x^l)}$$

$$Q_\pi(X^l = x^l) = \mathbb{E}_{Q \sim \pi} [Q^l(X^l = x^l)]$$

We first consider the case $\boxed{K=2}$. Then,

$$Q^l(X^l = x^l)$$

$$= \prod_{i=1}^l Q(X_i = x_i) = q^{n(1|x)} (1-q)^{n(0|x)}$$

To evaluate Q_π , we can evaluate

$$\mathbb{E}_{q \sim \text{unif}[0,1]} [q^i (1-q)^{l-i}] = \int_0^1 q^i (1-q)^{l-i} dq$$

$$\underbrace{\quad}_{I_{i,l}} = -\frac{q^i (1-q)^{l-i+1}}{(l-i+1)} \Big|_0^1 + \frac{i}{l-i+1} \int_0^1 (1-q)^{l-i+1} q^{i-1} dq$$

$$= \begin{cases} \frac{1}{l+1}, & i=0, l, \\ \frac{i}{l-i+1} \int_0^1 (1-q)^{l-i+1} q^{i-1} dq, & 1 \leq i \leq l-1. \end{cases}$$

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$$\text{Then, } I_{i,\ell} = \frac{i}{\ell-i+1} \cdot \frac{i-1}{\ell-i+2} \cdots \frac{1}{\ell+1} = \frac{i!(\ell-i)!}{(\ell+1)!}$$

and so, denoting $i = n(x/x^\ell)$,

$$\begin{aligned} \frac{Q_\pi(X^\ell = x^\ell, X_{\ell+1} = x)}{Q_\pi(X^\ell = x^\ell)} &= \begin{cases} \frac{1}{\ell+2} \cdot \frac{(i+1)!(\ell-i)!}{i!(\ell-i)!} & x=1 \\ \frac{1}{\ell+2} \cdot \frac{i!(\ell+1-i)!}{i!(\ell-i)!} & x=0 \end{cases} \\ &= \begin{cases} \frac{i+1}{(\ell+2)} & x=1, \\ \frac{\ell+1-i}{\ell+2} & x=0. \end{cases} \end{aligned}$$

$$Q_\pi(X_{\ell+1} = x | X^\ell = x^\ell) = \frac{n(x/x^\ell) + 1}{\ell+2}.$$

For $k=3$,

$$I_{i,j,\ell} = \int_0^1 q_1^i \int_0^{1-q_1} q_2^j (1-q_1-q_2)^{\ell-i-j} dq_2 dq_1$$

Note that the inner integral

$$\begin{aligned} \int_0^\theta q^j (\theta-q)^{\ell-j} dq &= \frac{j}{\ell-j+1} \cdot \frac{j-1}{\ell-j+2} \cdots \frac{1}{\ell} \cdot \int_0^\theta (\theta-q)^\ell dq \\ &= \frac{\theta^{\ell+1}}{\ell+1} \cdot \frac{1}{\binom{\ell}{j}} \end{aligned}$$

Thus,

$$\begin{aligned} I_{i,j,e} &= \frac{1}{(l-i+1)} \cdot \frac{1}{\binom{l-i}{j}} \int_0^1 q^i (1-q)^{l-i+1} dq \\ &= \frac{1}{(l-i+1)} \cdot \frac{1}{\binom{l-i}{j}} \cdot \frac{1}{(l+2) \binom{l+1}{i}} \\ &= \frac{(l-i-j)! j!}{(l-i)!} \cdot \frac{1}{(l+2)} \cdot \frac{(l-i)! i!}{(l+1)!} \\ &= \frac{1}{(l+2)(l+1)} \cdot \frac{1}{\binom{l}{i,j}} \end{aligned} \quad (6)$$

Thus, in general we get that a uniform prior yields an add-1 estimator.

D Jeffreys prior / Dirichlet prior

$k=2$ case

Consider $\pi(q) \propto \frac{1}{\sqrt{q(1-q)}}$

As before, we need to evaluate

$$\int_0^1 q^{i-\frac{1}{2}} (1-q)^{l-i-\frac{1}{2}} dq = \frac{\left(i-\frac{1}{2}\right)}{\binom{l-i-\frac{1}{2}+1}{l-i-\frac{1}{2}}} \dots \frac{\frac{3}{2}}{\binom{l-\frac{3}{2}}{l-\frac{3}{2}}} \int_0^1 \sqrt{q} (1-q)^{l-\frac{3}{2}} dq$$

Thus, $Q_\pi(X_{e+1}=0 | X^e = x^e)$ equals

$$\frac{(l-i+\frac{1}{2})}{(l-\frac{1}{2})} \left[\int_0^1 \sqrt{q} (1-q)^{l-\frac{1}{2}} dq / \int_0^1 \sqrt{q} (1-q)^{l-\frac{3}{2}} dq \right]$$

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Note that

$$I_{l+1} = \int_0^1 \sqrt{q} (1-q)^{l-\frac{1}{2}} dq = I_l - \frac{3}{2(l-\frac{1}{2})} I_{l+1}$$

$$\Leftrightarrow \left(\frac{2l+2}{2l-1} \right) \cdot I_{l+1} = I_l$$

$$\begin{aligned} \text{Therefore, } Q\pi(X_{l+1}=0 | X^l=x^l) &= \frac{l-i+\frac{1}{2}}{l+1} \\ &= \frac{n(0|x^l)+\frac{1}{2}}{n+1}. \end{aligned}$$

Similarly, can be extended to arbitrary k .

Finally, for any $\text{Beta}(a, b)$ prior,

$$\pi(q) \propto \frac{1}{q^{a-1}(1-q)^{b-1}}$$

add- α corresponds to $\text{Beta}(\alpha, \alpha)$.

→ In the next class, we will analyse the redundancy

corresponding to $\alpha = \frac{1}{2}$,

namely the Krichevsky-Trofimov (KT) estimator.