Journal Watch:

Transanctions on Signal Processing, May15, June-1, 2016

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Generalized Interference Alignment - Part 1: Theoretical Framework Authors: Lianzhong Ruan, Vincent Lau, Moe Win

- Generalized interference alignment (GIA) problem considers a very general setup comprising multiple transmitters, recievers, legitimate jammers and eavesdroppers.
 - Determine feasibility conditions for interference alignment
 - Design of transreceivers (precoder and decoder matrices)
- System model: Received signal at receiver-k:

$$\mathbf{y}_{k}^{(\iota)} = (\mathbf{U}_{k}^{(\iota)})^{\dagger} \left(\mathbf{H}_{kk}^{(\iota)} \mathbf{V}_{k} \mathbf{x}_{k} + \sum_{j=1, \neq k}^{K} \mathbf{H}_{kj}^{(\iota)} \mathbf{V}_{j} \mathbf{x}_{j} + \mathbf{z}_{k}^{(\iota)} \right)$$

► GIA transceiver design problem: Design $\{\mathbf{U}_k^{(\ell)}, \mathbf{V}_j\}$'s, such that

$$\operatorname{Rn}\left((\mathbf{U}_{k}^{(\ell)})^{\dagger}\mathbf{H}_{kk}^{(\ell)}\mathbf{V}_{k}\right)=d_{k},\ \forall k\in\{1,2,\ldots,K\},$$

 $\operatorname{Rn}(\mathbf{V}_j) = d_j, \ \forall j \in \{K+1, K+2, \dots, K\}, \qquad \text{and} \ (\mathbf{U}_k^{(\ell)})^{\dagger} \mathbf{H}_{kj}^{(\ell)} \mathbf{V}_j = \mathbf{0}, \ \forall (k, j) \in \mathcal{A}$

Feasibily constraints in GIA are nonlinear and non-convex.

- Feasibility of transceiver design problem = algebraic independence of GIA constraints = full rankness of associated Jacobian.
- ► Using algebraic geometry concepts, shows that local and global optimums have no performance gap.

Optimal Joint Detection and Estimation Based on Decision-Dependent Bayesian Cost

Authors: Shang-Li and Xiaodong Wang, Columbia Univ.

Joint detection and estimation problem:

 $\begin{aligned} \mathcal{H}_0 : & \mathbf{y} \sim f_0(\mathbf{y}|\theta_0), & \text{with } \theta_0 \sim \pi_0(\theta_0) \\ \mathcal{H}_1 : & \mathbf{y} \sim f_1(\mathbf{y}|\theta_1), & \text{with } \theta_1 \sim \pi_1(\theta_1) \end{aligned}$

y are observations and $\theta_{0/1}$ are unknown parameters under hypothesis $\mathcal{H}_{0/1}$.

- **Goal**: Decide between \mathcal{H}_0 and \mathcal{H}_1 , and at the same time, also estimate the unknown parameter θ .
- **Coupling** between the underlying estimation and detection problems:
 - The quality of parameter estimate depends upon the correctness of selected hypothesis
 - Estimate of parameter helps in deciding between the competing hypothesis

Naive approaches:

- Composite hypothesis testing followed by MAP estimation.
- Neyman-Pearson formulation: Minimize the estimation cost subject to constraint on detection performance.

Optimal Joint Detection and Estimation Based on Decision-Dependent Bayesian Cost

 Optimal joint detection and estimation: A Bayes estimation cost function is proposed:

$$\mathcal{C}(\hat{\theta}_0, \hat{\theta}_1, \delta) = \sum_{i,j \in \{0,1\}} c_{i,j} \mathcal{P}_{i,j}(\delta = j) \times \mathbb{E}_i \left(||\hat{\theta}_{c,j} - \theta_c||^2 + ||\hat{\theta}_{s,j} - \theta_{s,i}||^2 \cdot \mathbb{I}_{\{j=i\}} \mid \delta = j \right)$$

where δ is the detection output.

If prior on hypothesis is <u>not available</u>:

$$\begin{array}{l} \underset{\hat{\theta}_{0},\hat{\theta}_{1},\delta}{\text{minimize }} \mathcal{C}(\hat{\theta}_{0},\hat{\theta}_{1},\delta) \\ \text{subject to } \mathbb{P}_{0}(\delta=1) \leq \alpha, \quad \mathbb{P}_{1}(\delta=0) \leq \beta \end{array}$$

If prior on hypothesis is <u>available</u>:

$$\begin{array}{l} \underset{\hat{\theta}_0, \hat{\theta}_1, \delta}{\text{minimize }} \mathcal{C}(\hat{\theta}_0, \hat{\theta}_1, \delta) \\ \text{subject to } \mathbb{P}(H_0) \mathbb{P}_0(\delta = 1) + \mathbb{P}(H_1) \mathbb{P}_1(\delta = 0) \leq \gamma \end{array}$$

- Closed form solution for above problems is provided.
- Extension to multiple hypothesis is also discussed.

Design and Analysis of a Greedy Pursuit for Distributed Compressed Sensing

Authors: Dennis Sundman, Saikat Chatterjee, Mikael Skoglund, KTH Sweden

Distributed Compressed Sensing: Estimate x_i from y_i in distributed manner.

$$\mathbf{y}_j = \mathbf{A}_j \mathbf{x}_j + \mathbf{e}_j$$
 $j = 1, 2, \dots, L$

- Mixed support set model: Each \mathbf{x}_i can be decomposed as: $\mathbf{x}_i = \mathbf{x}_i^C + \mathbf{z}_i$.
 - $\mathbf{x}_1^C, \mathbf{x}_2^C, \dots, \mathbf{x}_l^C$ share a common sparse support.
 - **z**_{*i*} is sparse local innovation component.
- Distributed Parallel Pursuit (DIPP): Distributed Co-SAMP type algorithm with Co-SAMP inspired fusion of common support estimates across nodes.
- Main steps in single iteration of DIPP:
 - Generate local support estimate by CoSamp like update
 - Exchange current support estimates with neighboring nodes
 - Apply majority rule to obtain external support estimate
 - Refine local support by using external support estimate, once again using CoSamp like update
- ► RIP constant based stable signal recovery guarantees under measurement noise is provided.

Analytical Derivation of the Inverse Moments of One-Sided Correlated Gram Matrices With Aplications

Authors: K. Elkhalil, A. Kammoun, T. Y. Al Naffouri, and M. S Alouini

Moments of one-sided correlated gram matrices

- Let **H** be $n \times m$ random matrix with i.i.d zero-mean unit variance complex Gaussian random entries.
- Let Λ be deterministic PSD matrix with distinct eigenvalues.
- Then, consider the Gram matrix S as

$$S = H^* \wedge H.$$

• The *r*th moments $\mu_{\Lambda}(r)$ is defined as:

$$\mu_{\Lambda}(r) \triangleq \frac{1}{m} trace\left(\mathbb{E}_{\mathbf{H}}(\mathbf{S}^{r})\right).$$

Application-1: BLUE estimation error in (y = Hx + e) problem can be rewritten in terms of inverse moments.

$$\mathbb{E}_{\mathbf{H}}||\hat{\mathbf{x}}_{\mathsf{blue}} - \mathbf{x}||^2 = \mathbb{E}_{\mathbf{H}}tr(\mathbf{H}^*\boldsymbol{\Sigma}_{\mathsf{noise}}^{-1}\mathbf{H})^{-1} = m\mu_{\boldsymbol{\Sigma}_{\mathsf{noise}}^{-1}}(-1).$$

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▶ Application-2: Average estimation error in LMMSE for (y = Hx + e) problem:

$$\mathbb{E}_{\mathbf{H}}||\hat{\mathbf{x}}_{\text{Immse}} - \mathbf{x}||^{2} = \mathbb{E}_{\mathbf{H}} tr(\boldsymbol{\Sigma}_{\mathbf{x}}^{-1} + \mathbf{H}^{*}\boldsymbol{\Sigma}_{\text{noise}}^{-1}\mathbf{H})^{-1}.$$

High SNR regime

$$\mathbb{E}_{\mathbf{H}}\{\|\hat{\mathbf{x}}_{\text{immse}} - \mathbf{x}\|^2\} = m \sum_{k=0}^{l} \frac{(-1)^k}{\sigma_x^{2k}} \mu_{\Sigma_{\text{noise}}^{-1}}\left(-k-1\right) + o\left(\sigma_x^{-2r}\right)$$

where $l \leq p - 1$ with $p = \min(m, n - m)$.

Low SNR regime

$$\mathbb{E}_{\mathbf{H}}\{\|\mathbf{x}_{\text{immse}}-\mathbf{x}\|^2\} = m \sum_{k=0}^{\infty} (-1)^k \sigma_x^{2k+2} \mu_{\Sigma_{\text{noise}}(k)}$$

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Application-3: Accuracy of sample covariance matrix S:

$$\mathsf{Loss} = \mathbb{E} ||\mathbf{R}^{1/2} \hat{\mathbf{S}}^{-1} \mathbf{R}^{1/2} - \mathbf{I}||_F^2$$

- Once again, the Loss can be written in terms of inverse moments of sample covariance matrix Ŝ.
- Stieltjes Transform: For a Hermitian matrix A, its Stieltjes transform is given by:

$$m_{\mathbf{A}}(z) \triangleq \int \frac{1}{\lambda - z} dF^{\mathbf{A}}(\lambda) = \frac{1}{m} \operatorname{tr} (\mathbf{A} - z \mathbf{I}_m)^{-1}$$

where $F^{A}(.)$ is empirical spectral distribution of **A**.

Main result: The rth inverse moment can be written in terms of derivative of the Stieltjes transform.

Other Interesting Papers:

- Sequence Set Design With Good Correlation Properties Via Majorization-Minimization
- A General Design Framework for MIMO Wireless Energy Transfer With Limited Feedback
- Massive MIMO for Decentralized Estimation of a Correlated Source
- Bayes-Optimal Joint Channel-and-Data Estimation for Massive MIMO With Low-Precision ADCs
- Uplink Downlink Rate Balancing and Throughput Scaling in FDD Massive MIMO Systems
- Traffic Aware Resource Allocation Schemes for Multi-Cell MIMO-OFDM Systems
- Decentralized Sum Rate Maximization With QoS Constraints for Interfering Broadcast Channel Via Successive Convex Approximation
- SINR Constrained Beamforming for a MIMO Multi-User Downlink System: Algorithms and Convergence Analysis
- Robust Pilot Decontamination Based on Joint Angle and Power Domain Discrimination

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