Property Testing on Distributions

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- Property testing broadly refers to testing structural properties of data
- Some examples
 - Testing whether a graph can be clustered
 - Testing whether a boolean function is monotone
 - Testing whether samples are generated from the uniform distribution
- Decide whether an object has a property or is far from having it

Testing uniformity

- Given i.i.d. samples $\{X_i\}_{i=1}^n$ from an unknown discrete distribution P, determine whether P is the uniform distribution or far from it
- Test $\mathcal{H}_0: P = U$ vs $\mathcal{H}_1: ||P - U|| \ge \epsilon$

where U denotes the uniform distribution on [k], and $\|\cdot\|$ is some notion of distance between distributions

• We will consider the ℓ_1 distance between discrete distributions

$$||P - Q||_1 := \sum_{i=1}^k |P_i - Q_i|$$

- We first consider testing in ℓ_2 and extend it to ℓ_1
- For testing in ℓ_2 , need a good estimate for $||P U||_2^2$

$$||P - U||_2^2 = \sum_{i=1}^k (P_i - \frac{1}{k})^2$$
$$= ||P||_2^2 - \frac{1}{k}$$

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So we need to estimate $||P||_2^2$

Additive vs multiplicative accuracy

- How much error can we allow in our estimate for $||P||_2^2$?
- First note that $||P||_2^2 = \frac{1}{k}$, when P = U
- If $||P U||_2 \ge \epsilon$, then

$$||P||_{2}^{2} = ||P - U||_{2}^{2} + \frac{1}{k}$$
$$\geq \epsilon^{2} + \frac{1}{k}$$

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• Under ℓ_2 distance, can allow additive error of $\frac{\epsilon^2}{2}$

Additive vs multiplicative accuracy

• How much error can we allow in our estimate for $||P||_2^2$?

• If $||P - U||_1 \ge \epsilon$, then

$$\begin{split} \|P\|_{2}^{2} &= \|P - U\|_{2}^{2} + \frac{1}{k} \\ &\geq \frac{1}{k} \|P - U\|_{1}^{2} + \frac{1}{k} \\ &\geq \frac{1}{k} \epsilon^{2} + \frac{1}{k} = \frac{1}{k} (1 + \epsilon^{2}) \end{split}$$

• Under ℓ_1 distance, can allow multiplicative error of $\frac{\epsilon^2}{3} \|P\|_2^2$

- How to estimate $||P||_2^2$ using the samples?
- Use collision probabilities

$$Y_{ij} := \begin{cases} 1, & \text{if } X_i = X_j \\ 0, & \text{otherwise} \end{cases}$$
$$T := \frac{1}{\binom{n}{2}} \sum_{i < j} Y_{ij}$$

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Note that

$$\mathbb{E}T = \frac{1}{\binom{n}{2}} \sum_{i < j} \mathbb{E}Y_{ij}$$
$$= \frac{1}{\binom{n}{2}} \sum_{i < j} (P_1^2 + \ldots + P_k^2)$$
$$= \|P\|_2^2$$

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T is an unbiased estimate of $||P||_2^2$

Also

$$\mathbb{E}\left(\sum_{i < j} Y_{ij}\right)^{2} = \sum_{i < j} Y_{ij}^{2} + \sum_{\substack{i < j, k < l \\ 3 \text{ distinct indices}}} Y_{ij}Y_{kl} + \sum_{\substack{i < j, k < l \\ \text{all indices distinct}}} Y_{ij}Y_{kl} = \binom{n}{2} \|P\|_{2}^{2} + 6\binom{n}{3}\|P\|_{3}^{3} + \binom{n}{2}\binom{n-2}{2}\|P\|_{2}^{4}$$

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■ Therefore,

$$Var(T) = \frac{1}{\binom{n}{2}^{2}} Var(\sum_{i < j} Y_{ij})$$
$$\leq \frac{2}{n(n-1)} \|P\|_{2}^{2} + \frac{4}{n} \|P\|_{3}^{3}$$

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• For testing in ℓ_2 ,

$$\mathbb{P}\left(T - \|P\|_2^2 \ge \frac{\epsilon^2}{2}\right) \le \frac{1}{(\epsilon^2/2)^2} \operatorname{Var}(T)$$

which gives $n \ge \frac{60}{\epsilon^4}$ for error probability $\le 1/3$

• For testing in ℓ_1 ,

$$\mathbb{P}\left(T - \|P\|_{2}^{2} \le \frac{\epsilon^{2}}{3} \|P\|_{2}^{2}\right) \le \frac{9}{\epsilon^{4} \|P\|_{2}^{4}} \operatorname{Var}(T)$$

Simplifying and using $||P||_2^2 \ge \frac{1}{k}$ and $||P||_3 \le ||P||_2$, we get $n \ge \frac{216}{\epsilon^4}\sqrt{k}$ for error probability $\le 1/3$

- \blacksquare The tester we saw doest not give optimal dependence on ϵ
- A $O(\frac{\sqrt{k}}{\epsilon^2})$ dependence was shown to be optimal ¹
- Some other problems in distribution testing
 - Independence testing
 - Identity testing
 - Testing closeness of distributions
 - Testing unimodality

¹Liam Paninski. A coincidence-based test for uniformity given very sparsely sampled discrete data. IEEE Trans. on Information Theory, 2008 + 2 + 4 = -9

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