Sparse PCA from incomplete data

Lekshmi Ramesh



Indian Institute of Science Bangalore

April 14, 2018

▲□▶ ▲□▶ ▲臣▶ ★臣▶ = 臣 = のへで

1/16

Inference from data with missing values

- Missing data occur frequently in practice
- How to design good estimators/tests in the presence of missing data?
- Commonly used fix
 - Discard samples with missing values—can lead to loss of large amount of data
 - Imputation–usually ad hoc
- This presentation: Two approaches to do SPCA using incomplete data, guarantees

Approximating SPCA from incomplete data¹

- Data matrix $\mathbf{X} \in \mathbb{R}^{d \times n} n$ samples in d dimensions
- PCA: find solution to the the following variance maximization problem

$$\mathbf{V}_k = \operatorname*{arg\,max}_{\mathbf{V} \in \mathbb{R}^{d \times k}, \mathbf{V}^\top \mathbf{V} = \mathbf{I}} \operatorname{Tr} \, \mathbf{V}^\top \mathbf{X} \mathbf{X}^\top \mathbf{V}$$

SPCA: additional sparsity constraint on columns of V

$$\mathbf{S}_k = \underset{\mathbf{V} \in \mathbb{R}^{d \times k}, \mathbf{V}^\top \mathbf{V} = \mathbf{I}, \|\mathbf{V}_i\|_0 \leq r}{\arg \max} \operatorname{Tr} \mathbf{V}^\top \mathbf{X} \mathbf{X}^\top \mathbf{V}$$

¹Abhisek Kundu, Petros Drineas, and Malik Magdon-Ismail. "Approximating Sparse PCA from Incomplete Data". In: Advances in Neural Information Processing Systems. 2015, pp. 388–396.

- Missing data case
 - \blacksquare Only a sparse sampling of entries of ${\bf X}$ available–use it to construct the "sketch" $\tilde{{\bf X}}$
 - \blacksquare Solve SPCA using the sketch–call the output $\tilde{\mathbf{S}}$
- How does $\tilde{\mathbf{S}}_k$ perform as an approximation to \mathbf{S}_k ?
- Quality of approximation measured in terms of the objective

Theorem

Let \mathbf{S}_k and $\tilde{\mathbf{S}}_k$ be solutions to the Sparse PCA problem with full data and with sketched data, respectively. Then,

 $\operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) \geq \operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) - 2k\|\mathbf{X}\mathbf{X}^{\top} - \tilde{\mathbf{X}}\tilde{\mathbf{X}}^{\top}\|_{op}$

- Doing SPCA using $\tilde{\mathbf{X}}$ is good if $\tilde{\mathbf{X}}\tilde{\mathbf{X}}^{\top}$ closely approximates $\mathbf{X}\mathbf{X}^{\top}$
- Will see: $\|\mathbf{X}\mathbf{X}^{\top} \tilde{\mathbf{X}}\tilde{\mathbf{X}}^{\top}\|_{op}$ small if larger data entries sampled with higher probability

Forming the sketch $\tilde{\mathbf{X}}$

Thresholding-based scheme

$$\tilde{\mathbf{X}}_{ij} = \begin{cases} \mathbf{X}_{ij}, & \text{if } |\mathbf{X}_{ij}| \ge \delta \\ 0, & \text{else} \end{cases}$$

• (ℓ_1, ℓ_2) element sampling: Sample index (i, j) w.p. p_{ij}

$$\tilde{\mathbf{X}}_{ij} = \begin{cases} \frac{1}{p_{ij}} \mathbf{X}_{ij}, & \text{w.p. } p_{ij} \\ 0, & \text{w.p. } 1 - p_{ij} \end{cases}$$

• Note that
$$\mathbb{E}\tilde{\mathbf{X}}_{ij} = p_{ij}\frac{\mathbf{X}_{ij}}{p_{ij}} + (1 - p_{ij})0 = \mathbf{X}_{ij}$$

 • Choose p_{ij} as follows $(\alpha \in (0,1))$

$$p_{ij} = \alpha \frac{|\mathbf{X}_{ij}|}{\|\mathbf{X}\|_1} + (1-\alpha) \frac{\mathbf{X}_{ij}^2}{\|\mathbf{X}\|_F^2}$$

- Biases the sampling scheme towards larger elements
- Reasonable way to model sampling in some cases
 Recommendation systems: users more likely to rate items they like/dislike a lot (large positive/large negative)

More details: thresholding scheme

• Let
$$\mathbf{X} = \tilde{\mathbf{X}} + \Delta$$
. This gives $\|\Delta\|_F^2 = \sum_{|\mathbf{X}_{ij}| < \delta} \mathbf{X}_{ij}^2$

• Let $\tilde{r} = \frac{\|\mathbf{X}\|_F^2}{\|\mathbf{X}\|_{op}^2}$ be the stable rank of \mathbf{X}

Then

$$\begin{aligned} \|\mathbf{X}\mathbf{X}^{\top} - \tilde{\mathbf{X}}\tilde{\mathbf{X}}^{\top}\|_{op} &= \|\mathbf{X}\Delta + \Delta^{\top}\mathbf{X}^{\top} + \Delta^{\top}\Delta\|_{op} \\ &\leq 2\|\mathbf{X}\|_{op}\|\Delta\|_{op} + \|\Delta\|_{op}^2 \end{aligned}$$

• Choose
$$\delta$$
 so that $\|\Delta\|_F \leq \frac{\epsilon}{\sqrt{\tilde{r}}} \|\mathbf{X}\|_F$

■ Then, previous theorem gives

$$\operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) \geq \operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) - 2k\epsilon(1+\epsilon)\|\mathbf{X}\|_{op}^{2}$$

Main result–II

Sample complexity for (ℓ_1, ℓ_2) scheme

Sample *s* entries from $\mathbf{X} \in \mathbb{R}^{d \times n}$ to form the sparse sketch \mathbf{X} using the (ℓ_1, ℓ_2) scheme. Let \mathbf{S}_k and $\tilde{\mathbf{S}}_k$ be solutions to the Sparse PCA problem using \mathbf{X} and $\tilde{\mathbf{X}}$, respectively. Then, if the number of samples satisfies

$$s \geq 2k^2\epsilon^{-2}(\rho^2 + \frac{\epsilon}{3k})\log(\frac{d+n}{\delta})$$

with $\rho^2 = \tilde{r} \max(d, n) f(\alpha, \mathbf{X})$, we have that

 $\operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) \geq \operatorname{Tr}(\tilde{\mathbf{S}}_{k}^{\top}\mathbf{X}\mathbf{X}^{\top}\tilde{\mathbf{S}}_{k}) - 2k\epsilon(1+\epsilon)\|\mathbf{X}\|_{op}^{2}$

w.p. at least $1 - \delta$.

Sparse PCA with Missing Observations²

- Samples $\mathbf{X}_1, \ldots, \mathbf{X}_n$ in \mathbb{R}^d with mean zeros and covariance Σ
- Goal is to estimate the first principal component in the high dimensional (d > n) and missing data regime
- Covariance matrix of the data represented as

$$\boldsymbol{\Sigma} = \sigma_1 \theta_1 \theta_1^\top + \sigma_2 \Gamma,$$

where θ_1 is the first principal component, $\sigma_1, \sigma_2 \ge 0$ and $\Gamma \succeq 0$

²Karim Lounici. "Sparse Principal Component Analysis with Missing Observations". In: *High Dimensional Probability VI*. ed. by Christian Houdré et al. Basel: Springer Basel, 2013, pp. 327–356.

- For each sample \mathbf{X}_i , we observe its j^{th} entry \mathbf{X}_{ij} , independently of other entries, w.p. δ
- Let $\mathbf{Y}_1, \ldots, \mathbf{Y}_n$ be the observations where

$$\mathbf{Y}_{ij} = \delta_{ij} \mathbf{X}_{ij}$$

and $\delta_{ij} \stackrel{iid}{\sim} \operatorname{Ber}(\delta)$

 \blacksquare In practice, δ can be approximated using the fraction of observed entries per sample

- Recovering the first principal component: Let Σ_n denote the sample covariance matrix
 - Sparsity level known

$$\hat{\theta}_1 = \underset{\theta \in S^d: \ \|\theta\|_0 \le s}{\arg \max} \theta^\top \Sigma_n \theta$$

Sparsity level unknown

$$\hat{\theta}_1 = \operatorname*{arg\,max}_{\theta \in S^d} \theta^\top \Sigma_n \theta - \lambda \|\theta\|_0$$

・ロト ・四ト ・ヨト ・ヨト ・ヨ

12/16

Estimation procedure

Missing data case

- Sample covariance $\Sigma_n^{\delta} = \frac{1}{n} \mathbf{Y} \mathbf{Y}^{\top}$ formed using incomplete samples is biased
- \blacksquare Can apply the following correction to get an unbiased estimate of Σ

$$\tilde{\Sigma}_n = \frac{1}{\delta^2} \Sigma_n^{\delta} + \left(\frac{1}{\delta} - \frac{1}{\delta^2}\right) \operatorname{diag}(\Sigma_n^{\delta})$$

Final estimator

$$\hat{\theta}_1 = \underset{\theta \in S^d}{\arg\max} \, \theta^\top \Sigma_n \theta - \lambda \|\theta\|_0$$

Guarantees

For $\mathbf{X}_1, \ldots, \mathbf{X}_n$ subgaussian and $\mathbf{Y}_1, \ldots, \mathbf{Y}_n$ defined as before, let

$$\lambda = \frac{\sigma_1^2}{\sigma_1 - \sigma_2} \frac{\log(ed)}{\delta^2 n}$$

Then the estimate $\hat{\theta}_1$ satisfies

$$\|\hat{\theta}_1\hat{\theta}_1^\top - \theta_1\theta_1^\top\|_F^2 \le c\|\theta\|_0\tilde{\sigma}^2 \frac{\log(ed)}{\delta^2 n}$$

w.p. at least $1 - \frac{1}{d}$, where $\tilde{\sigma} = \frac{\sigma_1}{\sigma_1 - \sigma_2}$

Bound increases as $\sigma_1 - \sigma_2$ decreases-problem harder when separation between first and second singular values decreases

- Choice of λ depends on (unknown) singular values of Σ ; a choice of λ based on singular values of $\tilde{\Sigma}_n$ is also given
- For fully observed case, bound shows that roughly $\|\theta\|_0 \log(d)$ samples suffice
- For missing data case, additional $\frac{1}{\delta^2}$ penalty
- The δ^{-2} penalty is shown to be tight in the lower bound result

Thank you