### On the Fundamental Limits of Adaptive Sensing

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December 15, 2018

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- It is possible to reliably recover sparse signals from very few linear measurements
- Conventional schemes are non-adaptive, the measurement matrix is fixed beforehand
- Can adaptive schemes provide any advantage? For example, can we reduce the MSE if the rows of the measurement matrix are chosen adaptively?

### Introduction

• Main result: For any adaptive sensing scheme and estimation procedure, it is impossible to significantly outperform random projection followed by  $\ell_1$  minimization

Model

$$y = Ax + z$$

with  $A \in \mathbb{R}^{m \times d}$  and  $z \in \mathcal{N}(0, \sigma^2 I)$ 

• Let A be a random projection matrix with unit norm rows and let  $\hat{x}$  be the output of the Dantzig selector. Then,

$$\mathbb{E}\|\hat{x} - x\|_2^2 \le c\frac{k}{m}d\sigma^2\log d$$

provided  $m \ge k \log \frac{d}{k}$ . This scaling is optimal.

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• Let  $a_i^{\top}$  be the rows of A and recall

$$y_i = a_i^\top x + z_i, \quad i \in [m]$$

Adaptive scheme:  $a_i$  is a (possibly random) function of  $(a_1, y_1, \ldots, a_{i-1}, y_{i-1})$ .

• Main result (formal). Let  $d \ge 2$ ,  $k < \frac{d}{2}$  and consider any m. Then,

$$\inf_{\hat{x}} \sup_{\|x\|_0 \le k} \mathbb{E} \|\hat{x} - x\|_2^2 \ge c \frac{k}{m} d\sigma^2.$$

• First lower bound the MSE for x drawn from the following prior:

$$x_i = \begin{cases} \mu, \text{ w.p. } k/d \\ 0, \text{ w.p. } 1 - k/d. \end{cases}$$

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- Show lower bound for support recovery
- Extend to MSE lower bound
- Extend to lower bound for arbitrary x

- Can restrict to  $a_i$  that are deterministic functions of  $(y_1, \ldots, y_{i-1})$
- Assumptions:  $||a_i||_2 \leq 1, \sigma = 1$
- We first look at error in support recovery when adaptive schemes are allowed
- Error metric: expected Hamming distance

$$\mathbb{E}|\hat{S}\Delta S| = \sum_{i=1}^{d} \mathbb{P}(\hat{S}_i \neq S_i)$$

**Result:** Suppose x is sampled from the Bernoulli prior with k < d/2. Then any estimate  $\hat{S}$  obeys

$$\mathbb{E}|\hat{S}\Delta S| \ge k \left(1 - \frac{\mu}{2}\sqrt{\frac{m}{d}}\right)$$

• If signal amplitude is low (say  $\mu \leq \sqrt{d/m}$ ), then large number of errors  $(\mathbb{E}|\hat{S}\Delta S| \geq k/2)$ 

• Let 
$$P_{1,j} = P(\cdot|x_j \neq 0)$$
 and  $P_{0,j} = P(\cdot|x_j = 0)$  for any  $j \in [d]$ 

• Let 
$$\pi_1 = k/d$$
 and  $\pi_0 = 1 - k/d$ . Then  
 $\mathbb{P}(\hat{S}_j \neq S_j) = \pi_1 P_{1,j}(\hat{S}_j = 0) + \pi_0 P_{0,j}(\hat{S}_j \neq 0)$ 

• Optimizing over all tests, the Bayes risk is

$$B \ge \min(\pi_0, \pi_1)(1 - d_{TV}(P_{0,j}, P_{1,j}))$$

where  $d_{TV}$  denotes the total variation distance

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Expected Hamming distance

 $\mathbb{E}$ 

$$\hat{S}\Delta S = \sum_{j=1}^{d} P(\hat{S}_{j} \neq S_{j})$$

$$\geq \sum_{j=1}^{d} B_{j}$$

$$\geq \pi_{1} \sum_{j=1}^{d} (1 - d_{TV}(P_{0,j}, P_{1,j}))$$

$$\geq k \left(1 - \frac{1}{\sqrt{d}} \sqrt{\sum_{j=1}^{d} d_{TV}^{2}(P_{0,j}, P_{1,j})}\right)$$

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#### Using

$$\mathbb{E}|\hat{S}\Delta S| \ge k \left(1 - \frac{1}{\sqrt{d}} \sqrt{\sum_{j=1}^{d} d_{TV}^2(P_{0,j}, P_{1,j})}\right)$$

and

$$\sum_{j=1}^d d_{TV}^2(P_{0,j},P_{1,j}) \leq \frac{m\mu^2}{4}$$

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we get the final result. We now prove the second inequality.

• We first upper bound  $d_{TV}(P_{0,j}, P_{1,j})$ . Using Pinsker's inequality

$$d_{TV}^2(P_{0,j}, P_{1,j}) \le \frac{\pi_0}{2} D(P_{0,j} \| P_{1,j}) + \frac{\pi_1}{2} D(P_{1,j} \| P_{0,j})$$

• Let  $P_{0,j} = P_0$  and  $P_{1,j} = P_1$  and note

$$P_0 = \sum_{x'} P(x') P(y_1, \dots, y_m | x', x_j = 0)$$
  
=:  $\sum_{x'} P(x') P_{0,x'}$ 

where  $x' = (x_1, \dots, x_{j-1}, x_{j+1}, x_d)$ 

• Similar expression for  $P_1$ 

#### Thus

$$D(P_0 || P_1) = D(\sum_{x'} P(x') P_{0,x'} || \sum_{x'} P(x') P_{1,x'})$$
  
$$\leq \sum_{x'} P(x') D(P_{0,x'} || P_{1,x'})$$

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using joint convexity of KL divergence

- What are  $P_{0,x'}$  and  $P_{1,x'}$ ?
- Recall that

$$y_i = a_i^\top x + z_i$$
$$= \sum_{l=1}^d a_{il} x_l + z_i$$

• Let 
$$j \in [d]$$
. For  $x_j = 0$ ,  
 $y_i = c_i + z_i$ 

and for  $x_j = \mu$ 

$$y_i = c_i + \mu a_{ij} + z_i$$

where  $c_i = \sum_{l \neq j} a_{il} x_l$ 

#### ■ Thus,

$$P(y_i|x', x_j = 0) \equiv \mathcal{N}(c_i, \sigma^2)$$

#### and

$$P(y_i|x', x_j = 1) \equiv \mathcal{N}(c_i + \mu a_{ij}, \sigma^2)$$

#### This gives

$$D(P_{0,x'} || P_{1,x'}) = D(P(\underline{y} | x', x_j = 0) || P(\underline{y} | x', x_j = 1))$$
  
$$= \mathbb{E}_{P_{0,x'}} \sum_{i=1}^{m} \log \frac{P(y_i | x', x_j = 0)}{P(y_i | x', x_j = 1)}$$
  
$$= \mathbb{E}_{P_{0,x'}} \sum_{i=1}^{m} \frac{1}{2} \left( (y_i - c_i - \mu a_{ij})^2 - (y_i - c_i)^2 \right)$$
  
$$= \frac{\mu^2}{2} \sum_{i=1}^{m} \mathbb{E}_{P_{0,x'}} a_{ij}^2$$

■ Finally,

$$D(P_0 || P_1) \le \sum_{x'} P(x') D(P_{0,x'} || P_{1,x'})$$
$$\le \sum_{x'} P(x') \frac{\mu^2}{2} \sum_{i=1}^m \mathbb{E}(a_{ij}^2 | x_j = 0)$$

■ Similarly,

$$D(P_1 || P_0) \le \sum_{x'} P(x') \frac{\mu^2}{2} \sum_{i=1}^m \mathbb{E}(a_{ij}^2 | x_j \neq 0)$$

Relating to the total variation distance,

$$\sum_{j=1}^{d} d_{TV}^2(P_{1,j}, P_{0,j}) \le \frac{\mu^2}{4} \sum_{i,j} \mathbb{E}a_{ij}^2$$
$$= \frac{m\mu^2}{4}$$

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#### ■ Lower bound on expected Hamming distance

$$\mathbb{E}|\hat{S}\Delta S| \ge k \left(1 - \frac{1}{\sqrt{d}} \sqrt{\sum_{j=1}^{d} d_{TV}^2(P_{1,j}, P_{0,j})}\right)$$
$$\ge k \left(1 - \frac{\mu}{2} \sqrt{\frac{m}{d}}\right)$$

# Connecting to MSE

• Let 
$$S = \operatorname{supp}(x)$$
 and  $\hat{S} = \{j : \hat{x}_j \ge \mu/2\}$ . Then,  
 $\|\hat{x} - x\|_2^2 = \sum_{j \in S} (\hat{x}_j - x_j)^2 + \sum_{j \in S^C} \hat{x}J^2$   
 $\ge \frac{\mu^4}{2} |S \setminus \hat{S}| + \frac{\mu^2}{4} |\hat{S} \setminus S|$   
 $= \frac{\mu^2}{4} |\hat{S} \Delta S|$ 

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#### ■ Taking expectation,

$$\mathbb{E}\|\hat{x} - x\|_2^2 \ge \frac{\mu^2}{4} \mathbb{E}|\hat{S}\Delta S|$$
$$\ge \frac{\mu^2}{4} k \left(1 - \frac{\mu}{2} \sqrt{\frac{m}{d}}\right)$$

• Final step: bound for arbitrary x

■ E. A.-Castro, E. J. Candès and M. A. Davenport. On the Fundamental Limits of Adaptive Sensing. In IEEE Transactions on Information Theory, January 2013.