## Target Self-Localization to an Area

#### Prabhasa K



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Introduction

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- Average Area Uncertainty

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- Motivation:
  - Localization in indoor environments is challenging.
  - Advances in WSN has enabled low-cost infrastructure deployment.
  - Algorithms that are computationally efficient.

- **Problem:** Self-localization (within an area) of a target node using RSS measurements from beacons transmitting from known locations.
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  - Algorithms that are computationally efficient.

### Applications:

- Tracking position of a target on a factory floor or in a hospital (intrusion detection, fire alarm).
- Enabling Coginitve Radio spectrum through geo-location of WSDs.

### Random node deployment strategies



#### Figure: Source:ResearchGate

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## Comparison of deployment strategies

	Coverage	Connectivity	coverage
Constant	++	++	++
diffusion			
Continuous	+	+	+
diffusion			
R-random	$\pm$	$\pm$	$\pm$
diffusion			
Simple	_	_	_
diffusion			
Exponential			
diffusion			

Figure: Source:ResearchGate

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Tools: Group Testing, Order Statistics, Stochastic Geometry (PPP).

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#### Notation:

- K Number of beacons
- *M* Number of power thresholds
- $\delta$  Required Degree of Accuracy/Size of Grid Cells
- L Number of grid points in each dimension  $(L \triangleq \lceil \frac{1}{\delta} \rceil)$
- $\lambda$  Beacon Density ( $K/L^2$ )

## Illustration



Figure: Measurement process for Target Self-Localization

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• Beacon node  $b_i$  transmits with a power  $P_0$ . RSS is observed at the target node  $P_{rx,i} \triangleq \min(P_0, P_0(d_0/d_i)^{\eta})$ .

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- Target node compares the RSS  $P_{rx,i}$  with M predetermined intervals,  $\left\{ \mathcal{I}^{(j)} \triangleq \left( P_{th}^{(j-1)}, P_{th}^{(j)} \right] : j = 1, \dots, M, P_{th}^{(0)} = P_0 \right\}.$

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- The reading corresponding to  $b_i$  and  $\mathcal{I}^{(j)}$  is set as follows:

$$y_i^{(j)} \triangleq \begin{cases} 1, \quad P_{th}^{(j-1)} > P_{rx,i} \ge P_{th}^{(j)} \\ 0, \qquad \text{else.} \end{cases}$$
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**Objective:** (i) Minimize the area uncertainty, or (ii) Minimize beacon density required to meet the desired localization accuracy (with high probability).

• K beacon nodes deployed uniformly at random in the area of interest.

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- Let  $\nu_i \triangleq \sum_{j=1}^{M} j y_i^{(j)}$ , which can take M + 1 possible values: {0,1,..., M}. So the the set of all possible readings is  $\mathcal{V} \triangleq \{0, 1, ..., M\}^{K}$ , with  $|\mathcal{V}| = (M+1)^{K}$ .

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- Let  $P_{\nu}$  be the probability that the target present at  $(x_t, y_t)$  has a reading  $\nu$ . Averaging over both target and beacon deployment, the the average area uncertainty at  $(x_t, y_t)$  is:

$$\Omega = \sum_{\nu \in \mathcal{V}} \mathbb{E}\left[P_{\nu}^{2}\right].$$
(2)

#### Theorem

When K beacon nodes, each with a power contour of radius r, are distributed uniformly at random in A, the average area uncertainty in localizing the target is given by

$$\Omega_{a}(q) \approx \left[q^{2} + (1-q)^{2}\right]^{K}$$
(3)

where  $q \triangleq \mathbb{E}[X]$  and X is the r.v. representing coverage area of a single beacon. Further,  $q^* = 1/2$  minimizes (3), and the corresponding beacon radius is  $r^* = 0.512$  and the average area uncertainty is  $\Omega_a(q^*) = (1/2)^K$ .

### Proof.

Suppose the first *I* entries of the reading  $\nu$  are '1' and the remaining (K - I) entries are '0'. Since beacons are i.i.d. uniformly over A, the probability of observing the reading  $\nu$  is  $P_{\nu} = X^{I}(1-X)^{K-I}$ .

### Proof.

Suppose the first *l* entries of the reading  $\nu$  are '1' and the remaining (K - l) entries are '0'. Since beacons are i.i.d. uniformly over A, the probability of observing the reading  $\nu$  is  $P_{\nu} = X^{l}(1-X)^{K-l}$ .

There are  $\binom{K}{I}$  combinations of readings with *I* ones and K - I zeros. Therefore, the expectation of  $\sum_{\nu \in \mathcal{V}} P_{\nu}^2$  over the target location, i.e., the average area uncertainty in localization is given by

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$$\Omega = \mathbb{E}\left[\sum_{l=0}^{K} \binom{K}{l} (X^2)^l ((1-X)^2)^{K-l}\right],$$
$$= \mathbb{E}\left[ (X^2 + (1-X)^2)^K \right].$$

(4)

### Proof.

Further, by Jensen's inequality, the lower bound on (4) is given by

$$egin{aligned} \Omega &\geq \left(\mathbb{E}\left[X^2
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where  $q \triangleq \mathbb{E}[X]$ . In comparison to  $q^2 + (1 - q)^2$ , the variance term is nearly flat across different values of r:

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$$\Omega_a(q) \approx \left(q^2 + (1-q)^2\right)^K. \tag{5}$$

 $q^* = 1/2$  minimizes (5) over  $q \in [0, 1]$ , and the corresponding beacon radius is  $r^* = 0.512$ , computed using

$$q = (1/2)r^4 - (8/3)r^3 + \pi r^2.$$
(6)

#### Theorem

When K beacon nodes, each with M power contours of radii  $r_1 < r_2 < \ldots < r_m < \ldots < r_M$ , are distributed uniformly at random in A, the average area uncertainty in localizing the target is given by

$$\Omega_a \approx \left[ q_1^2 + \sum_{m=2}^M (q_m - q_{m-1})^2 + (1 - q_M)^2 \right]^K$$
(7)

where  $q_m \triangleq \mathbb{E}[X_m]$ , m = 1, 2, ..., M, and  $X_m$  is an r.v. representing the area coverage of a single beacon with radius  $r_m$ . The quantities  $q_m^* = \frac{m}{M+1}$ , m = 1, 2, ..., M, minimize (5), and the corresponding average area uncertainty is  $\Omega_a^* = \left(\frac{1}{M+1}\right)^K$ . Note that, the beacon radii  $r_m^*$ , m = 1, 2, ..., M, is obtained by inverse-mapping the  $q_m^*$  using (6).





Figure: Outer loop Target, Inner loop Beacons

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Figure: Outer loop Beacons, Inner loop Target

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Figure: Joint deployment

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• For the test corresponding to the  $j^{th}$  threshold interval of the  $i^{th}$  beacon's signal, the grid points in the annulus  $\mathcal{A}_i^{(j)}$  are tested. Let it be represented by  $\mathbf{a}_i^{(j)} \in \{0,1\}^{1 \times C}$ , where  $C \triangleq L_1 L_2$
- For the test corresponding to the j<sup>th</sup> threshold interval of the i<sup>th</sup> beacon's signal, the grid points in the annulus A<sub>i</sub><sup>(j)</sup> are tested. Let it be represented by a<sub>i</sub><sup>(j)</sup> ∈ {0,1}<sup>1×C</sup>, where C ≜ L<sub>1</sub>L<sub>2</sub>
- The entries corresponding to the points being tested are set to 1 and the remaining entries are set to 0.

## Illustration



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- The entries corresponding to the points being tested are set to 1 and the remaining entries are set to 0.
- The measurement process:

$$\mathbf{y} = \mathbf{A}\mathbf{x},\tag{8}$$

 $\textbf{x} \in \{0,1\}^{\textit{C} \times 1}$  - true position of the target.

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 The Column Matching Algorithm attempts to match the columns of A with test result vector y:

$$\mathcal{K} = \sup \{ \max\{ \mathbf{y}^t \mathbf{A} - \mathbb{1}_{algo}(\mathbf{y}^c)^t \mathbf{A} \} \},$$
(9)

## Column Matching Algorithm (xnor)



Figure: Target Localization in a 10x10 grid. Target shown by a yellow star.

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When the beacon nodes are distributed as PPP with intensity  $\lambda$ , the number of beacon nodes with power contours of radius r intersecting any vertical/horizontal line segment S is Poisson distributed with mean  $\mu_1 = \lambda(2r)$ . The total number of such intersections N on the line segment S is approximately Poisson distributed with mean  $\lambda(4r - \pi r^2)$ .



Figure: Illustration of the beacon power contours intersecting a line

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## Proof.

Consider a region  $\mathcal{R}$  formed by a rectangular strip of size  $1 \times 2r$ . The average number of beacon nodes that intersect  $\mathcal{S}$  is

$$\mu_1 = \lambda(\text{Area of } \mathcal{R}) = \lambda(2r). \tag{10}$$

The mean of the number of intersections on  $\ensuremath{\mathcal{S}}$  is given by

$$u = 2\lambda(2r - \pi r^{2}) + \lambda(\pi r^{2}) = \lambda(4r - \pi r^{2}).$$
(11)

The cumulative distribution function (cdf) of the largest among the spacings between successive ordered uniform r.v.s in the range [0, 1] is given by

$$Pr(V_{(n+1)} \le \delta) = 1 - \sum_{k=1}^{\min(n+1,L-1)} (-1)^{k-1} \binom{n+1}{k} (1-k\delta)^n, \quad (12)$$

where  $n \ge 0$ ,  $\delta \in (0, 1)$  and  $L \triangleq \lceil \frac{1}{\delta} \rceil$ .

## Proof.

The probability of the occurrence of at least one of the events  $V_i > \delta$  can be expressed as (Boole's formula)

$$Pr\left\{\bigcup_{i=1}^{n+1} (V_i > \delta)\right\} = \sum_i Pr(V_i > \delta) - \sum_{i < j} Pr(V_i > \delta, V_j > \delta) + \dots + (-1)^n Pr(V_1 > \delta, V_2 > \delta, \dots, V_{n+1} > \delta).$$
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$$+ \dots + (-1)^n Pr(V_1 > \delta, V_2 > \delta, \dots, V_{n+1} > \delta).$$
(13)

The joint distribution of k events  $V_1 > \delta$ ,  $V_k > \delta$  is symmetrical in  $V_i$ . The union event  $\bigcup_{i=1}^{n+1} (V_i > \delta)$  is the same as  $(V_{(n+1)} > \delta)$ .

$$Pr(V_{(n+1)} > \delta) = \sum_{k=1}^{\min(n+1,L-1)} (-1)^{k-1} \binom{n+1}{k} (1-k\delta)^n, \qquad (14)$$

#### Theorem

The average probability of the largest spacing between successive intersections being less than or equal to the size of the grid cell, when the number of intersections N is Poisson distributed with mean  $\mu$ , is given by

$$\mathbb{E}\left[\Pr(V_{(N+1)} \le \delta)\right] = 1 - \sum_{k=1}^{L-1} \frac{e^{-k\delta\mu} \left[\mu(1-k\delta) + k\right] \left[-\mu(1-k\delta)\right]^{k-1}}{k!},$$
(15)

where  $\delta \triangleq \frac{1}{L}$  is the size of the grid cell.

## Proof.

$$\mathbb{E}\left[\Pr(V_{(N+1)} > \delta)\right] = \sum_{n=0}^{\infty} \Pr(V_{(n+1)} > \delta)\Pr(N = n)$$

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## Proof.

$$\mathbb{E}\left[Pr(V_{(N+1)} > \delta)\right] = \sum_{n=0}^{\infty} Pr(V_{(n+1)} > \delta)Pr(N = n)$$
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=  $\sum_{k=1}^{L-1} \sum_{n=k-1}^{\infty} (-1)^{k-1} {\binom{n+1}{k}} (1-k\delta)^n \frac{e^{-\mu}\mu^n}{n!}$   
=  $e^{-\mu} \sum_{k=1}^{L-1} \frac{(-1)^{k-1}}{k!} \sum_{n=k-1}^{\infty} \frac{(n+1)}{(n+1-k)!} [\mu(1-k\delta)]^n$ 

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## Proof.

$$=e^{-\mu}\sum_{k=1}^{L-1}\frac{(-1)^{k-1}}{k!}\left[\sum_{n=k-1}^{\infty}\frac{(n+1-k)}{(n+1-k)!}[\mu(1-k\delta)]^n+\sum_{n=k-1}^{\infty}\frac{k}{(n+1-k)!}[\mu(1-k\delta)]^n\right]$$
(16)

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#### Proof.

$$=e^{-\mu}\sum_{k=1}^{L-1}\frac{(-1)^{k-1}}{k!}\left[\sum_{n=k-1}^{\infty}\frac{(n+1-k)}{(n+1-k)!}[\mu(1-k\delta)]^n+\sum_{n=k-1}^{\infty}\frac{k}{(n+1-k)!}[\mu(1-k\delta)]^n\right]$$
(16)

The inner summation terms of (16) are Taylor series expansions of the scaled exponential function in  $\mu(1-k\delta)$ , so

$$\mathbb{E}\left[\Pr(V_{(N+1)} > \delta)\right] = e^{-\mu} \sum_{k=1}^{L-1} \frac{(-1)^{k-1}}{k!} \left[\left[\mu(1-k\delta)\right]^k + k\left[\mu(1-k\delta)\right]^{k-1}\right] e^{\mu(1-k\delta)}.$$
 (17)

# Evaluating $\mu$



Prabhasa K (IISc)

Target Self-Localization to an Area

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• For a given  $\delta$ ,  $\mathbb{E}\left[Pr(V_{(N+1)} > \delta)\right]$  can be upper bounded by the first term of the summation in (17), leading to the lower bound:

$$\mathbb{E}\left[\Pr(V_{(N+1)} \leq \delta)\right] \geq 1 - e^{-\delta\mu}[\mu(1-\delta) + 1].$$

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• For small  $\delta$  (< 0.2) and relatively large  $\mu$  (> 33):

$$\mathbb{E}\left[\Pr(V_{(N+1)} \le \delta)\right] \approx 1 - \mu e^{-\delta\mu} = 1 - (4\lambda \bar{r}M)e^{-\delta(4\lambda \bar{r}M)}$$
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Parameters λ, r̄ and M alone affect E [Pr(V<sub>(N+1)</sub> ≤ δ)] through their product.

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- Parameters λ, r̄ and M alone affect E [Pr(V<sub>(N+1)</sub> ≤ δ)] through their product.
- Best choice of Algorithm: CMA with 'Xnor-Centroid-Fine Grid' operations (simulation results...)
- **Practical Interest:** Choosing the optimal beacon density to meet a given localization accuracy with high probability.

- We consider a square area A of size (a, a), with a = 10.
- Area  ${\cal A}$  divided into grid cell fine-ness varying from 5 imes 5 to 100 imes 100
- Location of the target, beacon nodes are chosen uniformly at random over *A*.
- The free-space path loss model has path loss exponent  $\eta = 2$ .
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- Monte Carlo simulations of 10000 location instantiations.
- Goal 1: Verifying the minimum average area uncertainty.
- Goal 2: Selecting the 'best' localization algorithm.
- **Goal 3**: To compute beacon density required for achieving target localization to a desired accuracy for a specified number of the instantiations (say, 90%) while varying parameters.

# Performance comparison: Matrix vs Xnor, Centroid vs Random (Coarse Grid)



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# Performance comparison: Matrix vs Xnor, Centroid vs Random (Coarse Grid)



# Performance comparison: Matrix vs Xnor, Centroid vs Random



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## Coarse vs Fine grid



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Image: A match a ma

## Coarse vs Fine grid



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## Coarse vs Fine grid





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## Coarse vs Fine grid approach



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## Coarse vs Fine grid approach





# Performance Metric comparison



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# Performance Metric comparison



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## **Additional Plots**



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Image: A match a ma



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Results



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Results



#### Varying $P_{loc}$ with beacon radius, for LHS 0.9, Grid dict 50x50, Fraction 1





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	Constant diffusion	Simple diffusion	R-random
Delivery rate	+	+	+
Consumed energy per packet	+	±	±
End-to-end delay	±	+	±
Fault-tolerance related to detection errors	_	±	+
Fault-tolerance related to transient errors	_	+	±
Fault-tolerance related to global errors	—	+	+
Network lifespan based on coverage	±	+	+
Network lifespan based on connectivity	_	+	+
Network lifespan based on the quality of surveillance	_	±	+

#### Figure: Source:ResearchGate

• Is there a way to connect Hamming distance b/w readings and the Euclidean distance b/w locations? (For noisy group testing)

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- In a stochastic Energy Harvesting setting, a reading of "0" could arise for two reasons. Given this dilemma, what is a good algorithm for estimating the target's location?

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- In a stochastic Energy Harvesting setting, a reading of "0" could arise for two reasons. Given this dilemma, what is a good algorithm for estimating the target's location?
- What is the optimal trade-off between number of power thresholds, beacon energy consumption (transmission range) and required localization accuracy in the above setting?