Dimensionality Estimation in Rank Deficient Noise using Profile Likelihood & Sparse Signal Recovery

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Outline



2 Rank-Deficient Noise

3 UCoNo-SBL

4 Different Approach

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5 Future Work

System Model



Figure: System Model

Goal: Recover **x** from measurements **y** when, **Q** and Φ are known

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Solution - CoNoSBL



- Modified EM-SBL to CoNo-SBL that recovers x from a mixture of noisy and noiseless measurements
- CRLB on MSE of x, when x is a compressible signal
- Assumes that **Q** is known
- Analysed time complexity

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 Solution when Q is unknown?

Non-uniform noise

- EM-SBL suffers from identifiability problem even when $\mathbf{Q}=\sigma^2\mathbb{I}$ is unknown
- Consider estimation of noise statistics when we have non-uniform noise i.e., **Q** = *D* using MMV setup



Figure: System Model

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Summary - SBL

Initialize
$$\Gamma \leftarrow I$$

 $\Gamma = \operatorname{diag}(\gamma_i), i = 1, \dots, N$
E - Step : $\operatorname{Pr}(X|Y; \Gamma, Q) \sim N(\mu, \Sigma)$
 $\Sigma = (\Phi Q^{-1} \Phi^T + \Gamma^{-1})^{-1}$
 $\mu = \Sigma \Phi^T (Q + \Phi \Gamma \Phi^T)^{-1} Y$
M - Step : $\gamma_i^{(t+1)} = |\mu_i|^2 + \Sigma_{(i,i)}$
 $\hat{x} = \mu$
Figure: Summary of SBL

Algorithm

1
$$\bar{\mathbf{D}} = \mathcal{I}_m$$

2 compute $\hat{\Gamma} = MSBL(\mathbf{Y}, \mathbf{A}, \hat{\mathbf{D}})$

 $\mathbf{8}$ Find S =

{indices corresponding to k maximum magnitude entries}

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$$P = \mathbf{A}_{S} (\mathbf{A}_{S}^{H} \mathbf{A}_{S})^{-1} \mathbf{A}_{S}^{H}$$

$$\hat{\mathbf{D}} = \frac{1}{L} \sum_{i=1}^{L} (\mathcal{I} - P) \mathbf{Y}$$

6 Repeat 2 to 5 till convergence

Convergence criteria:

 $||\Gamma^{t+1} - \Gamma^t||_F \le 10^{-4} \& ||\hat{\mathbf{D}}^{t+1} - \hat{\mathbf{D}}^t||_F \le 10^{-2}$

Simulation Results



Figure: MSE vs. L



Figure: Support recovery vs. L

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Figure: MSE vs. SNR

0.9 0.8 0.7 0.6 0.5 0.5 0.3 MUDSASBL 0 ٥ MBP MSBL SSP MOMP SNR = 10dB N = 50 0.2 0.1 m = 30 k = 5 0 15 10 L

Figure: Support Recovery vs. SNR

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- Let $\mathcal{S} = \{ \text{Support of } \mathbf{x} \}$
- $\mathbf{y}_{\mathcal{S}} = \Phi_{\mathcal{S}} \mathbf{x}_{\mathcal{S}} + \mathbf{n}_{\mathcal{S}}$
- $\mathbf{y}_{\mathcal{S}^c} = \mathbf{n}_{\mathcal{S}^c}$
- $\hat{D} = \frac{1}{L} \sum diag(\frac{1}{m} \sum \mathbf{y}_{\mathcal{S}^c}^H(:, i) \mathbf{y}_{\mathcal{S}}^c(:, i))$
- Estimation of $\hat{\mathbf{D}}$ involves estimating m parameters
- From simulations, it is seen that L = O(k) measurements are required for successful recovery

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Rank-Deficient noise

- Involves estimation of $m^2 + 1$ parameters(m^2 elements and p)
- MUDSASBL gives poor estimate of covariance matrix
- Follow PCA approach:
- $\Sigma_y = \Phi \Gamma \Phi^T + \mathbf{Q}$
- $\tilde{\mathbf{Q}} = \Sigma_y \Phi \Gamma \Phi^T$
- Use $\tilde{\mathbf{Q}}$ as an estimate of \mathbf{Q}
- To use CoNo-SBL, need to identify the dimensionality of underlying noise subspace ({p})

PCA

- Dimensionality reduction converts a set of observations of possibly correlated variables into a set of values of uncorrelated variables called as principal components
- $\{x_i\}_{i=1}^n \in R^m$ be the data matrix
- Goal: To find $\alpha_1, ... \alpha_p$ such that $Var(\alpha_1^T x_i) \ge Var(\alpha_2^T x_i) \ge \cdots Var(\alpha_p^T x_i)$ and $Cov(\alpha_k^T x_i, \alpha_l^T x_i) = 0 \forall k \ne l$
- Mathematically, if d_j = Var(α_j^Tx_i) and S is the sample covariance matrix of the data, then PCA: R^m → R^p and α₁, · · · α_p are eigen vectors of S and d₁, · · · , d_j are corresponding eigen values

What value of p to choose from the eigen values!!

• Percent Variance: Find q between 1 and m such that

$$\frac{d_1 + d_2 + \dots + d_q}{d_1 + d_2 + \dots + d_m} \ge \gamma$$

, where γ is a pre-determined proportion, say 80% or 90%

- Scree test: Plot the eigenvalues d₁, d₂, ..., d_p in descending order (scree plot) and look for a "big gap" or an "elbow" in such a graph.
- Sequential tests:

For $j = 1, 2, \cdots, m-1$, consider a series of null hypotheses:

$$H_{0,j}$$
: $d_m = d_{m1} = \cdots = d_{mj}$

Start by testing $H_{0,1}, H_{0,2}, \cdots$ until a null hypothesis is first rejected. Suppose $H_{0,q}$ is the first rejected null hypothesis, then the first m q components are retained.

Drawbacks:

- Threshold γ
- No objective function to determine gap or elbow
- Sequential detection assumes that data is generated from multivariate gaussian distribution

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• Computationally expensive

Objective: To propose a simple method to find gap in a objective and automated way

- Main idea: assume a distribution on d_j's and find p by maximising profile likelihood
- Profile Likelihood: Suppose *l*(θ, ψ; y) be the likelihood function, θ main parameter and ψ- nuisance parameter, then profile likelihood for θ is defined as

$$I_{ heta}(heta; y) = I(heta, \hat{\psi_{ heta}}; y)$$

 $\hat{\psi}_{\theta}$ is MLE of ψ for fixed θ

- Advantages of profile likelihood:
 - Always available
 - Maximum of profile likelihood is always same as MLE of $\boldsymbol{\theta}$

- $d_1 \geq d_2 \geq \cdots \leq d_m$ be the eigen values
- for a fixed number $1 \le p \le m$, $S_1 = \{d_1, \cdots d_p\}$ and $S_2 = \{d_{p+1}, \cdots d_m\}$
- If an elbow or gap exists at p, then $S_1 \in f(d; \theta_1)$ and $S_2 \in f(d; \theta_2)$. Assume S_1 and S_2 are independent
- p is the main parameter of interest

$$l(p,\theta_1,\theta_2) = \sum_{i=1}^{p} \log f(d;\theta_1) + \sum_{j=p+1}^{m} \log f(d;\theta_2)$$

 $heta_1$ and $heta_2$ are unknown and should be computed from data

Let $\hat{\theta_1}$ and $\hat{\theta_2}$ be MLEs of θ_1, θ_2 respectively

$$l_{p}(p) = \sum_{i=1}^{p} \log f(d; \theta_{1}(p)) + \sum_{j=p+1}^{m} \log f(d; \theta_{2}(p))$$
$$\hat{p} = \operatorname{argmax}_{p} l_{p}(k) \forall k = 1, \cdots, m$$

- Choose f to be Gaussian distribution
- The parameters will now be $\theta_1 = \{\mu_1, \sigma^2\}$ and $\theta_2 = \{\mu_2, \sigma^2\}$

• MLEs for μ_1 , μ_2 and σ^2 are computed using sample averages,

$$\hat{\mu_1} = \frac{\sum_{i \in \mathcal{S}_1 d_i}}{p}$$
$$\hat{\mu_2} = \frac{\sum_{j \in \mathcal{S}_2 d_j}}{m - p}$$

and

$$\hat{\sigma^2} = \frac{(p-1)s_1^2 + (m-p-1)s_2^2}{m-2}$$

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where s_j^2 is sample variance of S_j

Numerical results



Figure: Rank estimation using screeplot *vs.* percent variance



Figure: Phase transition diagram for estimation of rank using scree plot

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UCoNo-SBL



Figure: System model - Unknown noise covariance matrix

Proposed method

Algorithm:

1 Let $x_i \sim \mathcal{N}(0, \Gamma_i)$

2
$$\Gamma_i \leftarrow SBL(y_i, \Phi, \mathbb{I})$$

- **6** Compute $\hat{\mathbf{Q}} = \frac{1}{L} \sum_{i=1}^{L} (y_i y_i^T \Phi \Gamma_i \Phi^T)$
- **4** find \hat{p} using eigen values of **Q**, let V_p be the corresponding eigen vectors

3
$$Y_1 = V_p^H Y$$
, $Y_2 = V_{m-p}^H Y$, $\Phi_1 = V_p^H \Phi$, $\Phi_2 = V_{m-p}^H \Phi$ and $\hat{D} = \Lambda(1:p)$

$$(X_i, \Gamma_i) = CoNoSBL(Y_1(i), Y_2(i), \Phi_1, \Phi_2, \hat{D}) \forall i$$

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6 Repeat 3 to 5 until convergence

- Convergence criteria considered is convergence of **Q** in terms of frobenius norm on error of covariance matrix
- Since, we are using sample covariance matrix as an estimate of Covariance matrix, requires L ≥ m

- Consider some training samples to estimate the noise subspace
- Compute p̂ from Q̂
- Use CoNoSBL with this estimate of Q
- Identical to using an estimated ${\bf Q}$ instead of true ${\bf Q}$
- Relatively faster compared to previous approach, however naive in implementation

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• Use some other metrics for convergence like chordal distance, spectral distance etc

- Propose a method to recover Q using lesser number of measurements
- Theoretical guarantees on convergence of Γ and \boldsymbol{Q}
- Analysis on Phase transition of CoNo-SBL

Thank You

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