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Outline

- Background and motivation
- Sparse Bayesian Learning
- Joint-sparse recovery
 - Support recovery guarantees
- Extensions and new algorithms
- Applications in communication systems

Part 1: Setting the Stage



Motivation and background

Basic results

Sparse Signal Recovery \widetilde{m} V y Φ X $M \times 1$ $M \times 1$ M x N noise measurements Measurement matrix $N \times 1$ M < Na.k.a. Dictionary sparse signal k nonzero entries, k << N

Goal: Recover x from y

M << N: infinitely many solutions</p>



- Wireless channels exhibit multipath
 - Naturally sparse in the lag-domain
 - Need to estimate both support & channel
- Channel equalization & data detection

Compressed Sensing Deals with three main questions: Design of sensing matrices Sparsifying Basis $\Phi_{M\times N} = \mathbf{A}_{M\times N} \Psi_N^{\bullet}$

- Guarantees for recovery
- Computationally efficient algorithms

This talk: New algorithms and guarantees for sparse signal recovery!

Robust Linear Regression: Underdetermined Case



Robust Linear Regression: Overdetermined Case

Measurement model:

 $\mathbf{y} = \mathbf{A}\mathbf{x} + \mathbf{E} + \mathbf{e}$

 $M \times N;$ Outliers; Noise $M \ge N$ sparse

• Use SVD: $\mathbf{A} = \mathbf{U}_1 \Sigma \mathbf{V}_1^T$; $\mathbf{U}_2^T \mathbf{A} = \mathbf{0}$

Processed measurements:

 $\tilde{\mathbf{y}} = \mathbf{U}_2^T \mathbf{y} = \mathbf{U}_2^T \mathbf{E} + \mathbf{U}_2^T \mathbf{e}$

Can now directly apply sparse signal recovery algorithms to estimate and remove outliers!

The Problem

- Noiseless case: Given y and Φ , solve $\min \|\mathbf{x}\|_0$ subject to $\mathbf{y} = \Phi \mathbf{x}$
- Noisy case: solve $\min \|\mathbf{x}\|_0$ subject to $\|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_2 \leq \beta$
- Lo norm minimization
 - Combinatorial complexity
 - Not robust to noise

Breakthrough 1: The Null Space Property

- Underdetermined systems: y = Φx; Φ is M x N, M < N,
 x is k-sparse:
 - Infinitely many solutions, but ...
 - Unique soln. if <u>nullspace</u> of Φ has no "sparse" vectors [Donoho, Elad '02]
 - Recovery of all k-sparse x: M ≥ 2k is nec. & suff.
 - Given x, unique soln. with high probability, if M ≥ k+1 [Bresler; Wakin etc]
- Thus: Sub-Nyquist sampling (compression) possible:
 - When we restrict to sparse signals
 - And sample in an "appropriate" basis Φ

Breakthrough 2: Just Relax!

- l_1 min. instead of l_0 min. min $\|\mathbf{x}\|_1$ subject to $\mathbf{y} = \mathbf{\Phi}\mathbf{x}$
 - Convex optimization problem; linear program
- Same solution as Lo minimization!
 - If the measurement matrix is random
 - Use slightly larger number of measurements
 - Robust to measurement noise $M \approx k \log\left(\frac{N}{k}\right) \ll N$

See [Donoho; Candes, Romberg, Tao etc]

Breakthrough 3: Recovery Guarantees

- Noisy measurements: y = Φx + ν
 We solve(P_{1,η}): min ||z||₁ subject to ||Φz y||₂ ≤ η
 Robust NSP: Φ satisfies RNSP(k) if, ∀ S of cardinality k, ||w_S||₁ ≤ ρ||w_{S^c}||₁ + τ ||Φw||₁ ∀ w
- Result: If Φ satisfies RNSP(k), a sol. x* of $(P_{1,\eta})$ with $\|\mathbf{v}\|_2 \leq \eta$ and $\mathbf{y} = \mathbf{A}\mathbf{x} + \mathbf{v}$ satisfies by a k-sparse vector

$$\|\mathbf{x} - \mathbf{x}^*\|_1 \le 2\left(\frac{1+\rho}{1-\rho}\right) \sigma_k(\mathbf{x})_1 + \frac{4\tau}{1-\rho} \eta$$
 Noise

Analysis of BP: stable and robust recovery

• Theorem: Suppose $2k^{th}$ Restricted Isometry <u>Constant</u> of Φ satisfies $\delta_{2k} < \frac{4}{\sqrt{41}} \approx 0.6246$ then for x,y with $||Ax-y||_2 \le \eta$, the sol x* of $\min ||\mathbf{z}||_1$ subject to $||\Phi \mathbf{z} - \mathbf{y}||_2 \le \eta$ satisfies \mathbf{z}

$$\|\mathbf{x} - \mathbf{x}^*\|_1 \le C\sigma_k(\mathbf{x})_1 + D\sqrt{k\eta}$$
$$\|\mathbf{x} - \mathbf{x}^*\|_2 \le \frac{C}{\sqrt{k}}\sigma_k(\mathbf{x})_1 + D\eta$$

where C, D >0 depend only on δ_{2k}

Breakthrough 4: Recovery Algorithms

Greedy algorithms:

- Matching pursuit [Mallat, Zhang; Cotter, Rao]
- Orthogonal matching pursuit [Tropp 03]
- COSAMP [Needell, Tropp]

Relaxation based methods (minimize diversity meas.):

- Basis pursuit (1-p, with p=1) [Chen et al.]
- Lasso (BPDN) [Tibshirani]
- Dantzig selector [Candes, Tao]
- Homotopy based methods (e.g., LARS) [Garrigues et al. 09]
- FOCUSS (1-p, with p < 1) [Gordonitsky et al.]</p>
- Iterative methods:
 - Basic/Iterative hard thresholding
 - Hard thresholding pursuit

Recovery guarantees exist for most of these algorithms! See [Rauhut & Foucart]

Limitations of Greed & Relaxation

Performance of BP and OMP depend on Φ

- Poor performance when conditions are violated
- Hard to relate estimation error to the dictionary
- Correlated dictionary: disrupts lo-l1 equivalence
- BP: performance independent of nonzero coeffs [Malibutov et al. 2004]
 - Cannot improve when situation is favorable
- OMP: performance highly sensitive to magnitudes of nonzero coefficients

Poor performance with unit magnitudes

Other issues:

Estimating embedded parameters, exploiting additional structure when available

To Recap

Sparse signal recovery

- Basic problem, breakthroughs in CS
- Algorithms
- Guarantees
- Limitations
 - Scaling/shrinkage
 - Correlated dictionary
 - Embedded parameters

Part 2: Don't Relax!



A time and place for nonconvex methods?

Bayesian Methods

MAP estimation (Type I):

- Also a regression problem with sparsity promoting penalties (e.g., lp-norm)
- L₁-min (BP/LASSO) is a special case
- Hierarchical Bayesian methods (Type II):
 - Iterative reweighted L1 [Candes et al. 2008]
 - Iterative reweighted L2 [Chartrand \$ Yin 2008]
 - EM-based SBL [Tipping, 2001], [Wipf, Rao 2007]
 - AMP [schniter 2008], [Rangan 2011]

$$\widehat{\mathbf{x}} = \arg \max_{\mathbf{x}} p(\mathbf{x}|\mathbf{y})$$

$$= \arg \min_{\mathbf{x}} -\log p(\mathbf{y}|\mathbf{x}) - \log p(\mathbf{x}) \text{ (Bayes' rule)}$$

$$= \arg \min_{\mathbf{x}} \|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_{2}^{2} + \lambda \sum_{i=1}^{N} g(|x_{i}|)$$
Separable prior

- For sparse solutions, g(|x_i|) should be a concave, nondecreasing function
 - Example: $g(|x_i|) = |x_i|^p, p \le 1$
 - Lasso is a special case: p=1
- Any Local min. of the MAP estmn problem has at most M nonzeros [Rao et al., 99]

The Optimization Problem

To solve

$$\arg\min_{\mathbf{x}} G(\mathbf{x}) \triangleq \|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_{2}^{2} + \lambda \sum_{i=1}^{N} g(|x_{i}|)$$

- g(x) concave, monotonically \uparrow in |x|
- G(x) convex + concave
- Many options for g(x) to promote sparsity
 Many options for solving the optz. problem

Sparsity-Promoting Penalties

 $f(\theta)$

 θ_n

- Concave penalty fns. promote sparsity
 - $g(x) = log(x^2 + \varepsilon), \varepsilon > 0$ [Chartrand \$ Yin 2008]
 - $g(x) = log(|x| + \epsilon), \epsilon > 0$ [candes et al. 2008]
 - g(x) = |x|P, 0
- A general approach: majorize-minimize

 $\mathbf{g}(\mathbf{\theta})$ $\Theta_{n-1} g(\Theta_n) \leq g(\Theta_{n-1})$

Majorization-Minimization Approach

- Find an upper bound $g(x) \leq f(x|x^{(m)})$ • Equality at $x = x^{(m)}$, convenient for opt.
- Step 1: Optimize $\arg \min_{\mathbf{x}} F\left(\mathbf{x} | \mathbf{x}^{(m)}\right) \triangleq \|\mathbf{y} - \mathbf{\Phi} \mathbf{x}\|_{2}^{2} + \lambda \sum_{i=1}^{N} f\left(|x_{i}| | x_{i}^{(m)}\right)$ • Step 2: Set m <- m+1, update $f(\mathbf{x} | \mathbf{x}^{(m)})$, iterate
- Works because $G(x^{(m+1)}) \leq F(x^{(m+1)}|x^{(m)}) \leq F(x^{(m)}|x^{(m)}) = G(x^{(m)})$

Iterative Reweighted L1

- Concavity in |x|: $g(x) \leq g'(x(m))(x-x(m)) + g(x(m))$ • Equality at x = x(m), linear in x
- Iterative reweighted L1: [Candes et al. 08]
 - Init: m = 0, x^(m) = something convenient
 - Iterate:
 - Optimize $\mathbf{x}^{(m+1)} = \arg\min_{\mathbf{x}} \|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_{2}^{2} + \lambda \sum_{i=1}^{N} g'(x_{i}^{(m)}) |x_{i}|$ • $m \leftarrow m+1$, update $g'(x_{i}^{(m)})$ • Until convergence Weighted l_{1} minimization

Iterative Reweighted L2

- g(x) concave in x^2 : $g(x) \le \left(\left. \frac{\partial g(\sqrt{x^2})}{\partial (x^2)} \right|_{x=x_0} \right) (x^2 x_0^2) + g(x_0)$
- Optimization problem $\mathbf{x}^{(m+1)} = \arg\min_{\mathbf{x}} \|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_{2}^{2} + \lambda \sum_{i=1}^{N} ||\mathbf{x}_{i}|^{2}$

- S Iterative reweighted 12 [Chartrand et al. 08]
 - Init: m = 0, $x^{(m)} =$ something convenient
 - Iterate:
 Compute $\mathbf{x}^{(m+1)} = \mathbf{W}_m \Phi^T \left(\lambda \mathbf{I} + \Phi \mathbf{W}_m \Phi^T \right)^{-1} \mathbf{y}$ m <- m+1, update W_m
 - Until convergence

An Example

- Suppose $g(x) = \log(|x| + \varepsilon)$, $\varepsilon > 0$ • Concave in |x|, x^2
- Iterative reweighted l1 $g'\left(x_{i}^{(m)}\right) = \left[\left|x_{i}^{(m)}\right| + \epsilon\right]^{-1}$

Iterative reweighted 12

$$w_i^{(m)} = \left[\left(x_i^{(m)} \right)^2 + \epsilon \left| x_i^{(m)} \right| \right]^{-1}$$

Limitations of MAP

Many local minima O(NCM)

May get stuck at a local minimum

- MAP only guarantees max $p(x = x_0|y)$
 - Probability mass, rather than mode, may be more relevant for continuous random vars
 - Perhaps posterior mean E(x|y)?
- Even with the true prior, MAP estimators do not minimize MSE: so MSE may be high!
 - In fact, using "true" statistics often does not lead to the lowest MSE!

To Recap

- Bayesian estimation
 - Basic MAP estimation
 - Majorization-minimization approach
 - Iterative reweighted algorithms
- Limitations
 - Many Local minima
 - Posterior mean vs. posterior mode

Part 3: Sparse Bayesian Learning



Use lots of priors and pick the best one!

Point of Departure: Alternative Prior

Gaussian Scaled Mixtures (GSM) $\mathbf{x} = \gamma G; \ G \sim \mathcal{N}(\mathbf{g}; 0, 1)$ $p(\mathbf{x}) = \int p(\mathbf{x}|\gamma) p(\gamma) d\gamma = \int \mathcal{N}(\mathbf{x}; 0, \gamma) p(\gamma) d\gamma$ Y: +ve random variable, indep. of G • Spike-and-slab model if $\gamma = Bern(0,1)$ Most priors can be expressed using GSM (incl. ones with concave g) [Palmer et al., 2006]

Examples

- We use: $p(\gamma) = \frac{a^2}{2} \exp\left(-\frac{a^2}{2}\gamma\right), \gamma \ge 0$ • And get: $p(x_i; a) = \frac{a}{2} \exp(-a|x_i|)$
- Which leads to the familiar LASSO problem

Student's t distribution

- We use: gamma distribution
- And get: $p(x_i; a, b) = \frac{b^a \Gamma(a + 1/2)}{\sqrt{2\pi} \Gamma(a)} \frac{1}{(b + x_i^2/2)^{a+1/2}}$

Examples

Generalized Gaussian

- We use: positive alpha-stable density of order p/2• And get: $p(x_i; p) = \frac{1}{2\Gamma\left(1 + \frac{1}{p}\right)} \exp(-|x_i|^p)$
- Generalized Logistic distribution
 - We use: A scale mixing density related to the Kolmogorov Smirnoff distance

• And get:

$$p(x_i; \alpha) = \frac{\Gamma(2\alpha)}{\Gamma(\alpha)^2} \frac{\exp(-\alpha |x_i|)}{(1 + \exp(-|x_i|))^{2\alpha}}$$

Sparse Bayesian Learning

Canonical model



Parameterized Gaussian prior:

$$p(x_i; \gamma_i) = \frac{1}{\sqrt{2\pi\gamma_i}} \exp\left(-\frac{x_i^2}{2\gamma_i}\right), \ \gamma_i \ge 0$$

Graphical Model

- Markov chain (graphical model): Y -> X -> y
- p(x;y) Gaussian leads to tractable algorithms
- Given Y, p(x|y;Y) is
 Gaussian: easy to find
 point estimates
- But we don't know y

• When in doubt, approximate! Find $p(\mathbf{x}|\mathbf{y};\hat{\gamma})$ instead



 $\mathsf{y} = \Phi \mathsf{x} + \mathsf{v}$

Approach

- First, estimate hyperparameters: $\hat{\gamma} = \arg \max_{\gamma} p(\gamma | \mathbf{y})$
 - γ: deterministic and unknown, or random with hyperprior distbn.
- Then, find posterior distribution $p(\mathbf{x}|\mathbf{y};\hat{\mathbf{y}})$ $p(\mathbf{x}|\mathbf{y};\hat{\mathbf{y}}) = \mathcal{N}(\mu_x, \Sigma_x)$ $\mu_x = \hat{\Gamma}\Phi^T \left(\Phi\hat{\Gamma}\Phi^T + \lambda \mathbf{I}\right)^{-1} \mathbf{y}$ $\Sigma_x = \hat{\Gamma} - \hat{\Gamma}\Phi^T \left(\Phi\Gamma\Phi^T + \lambda \mathbf{I}\right)^{-1} \Phi\hat{\Gamma}$

• For point estimates: e.g., posterior mean: $\mathbb{E}\left(\mathbf{x}|\mathbf{y};\hat{\gamma}
ight)$

Estimating the Hyperparameters

• Estimate γ_i from the data: Type-II ML $\mathcal{L}(\Gamma) = \log p(\mathbf{y}; \Gamma) = \log \int p(\mathbf{y} | \mathbf{x}; \Gamma) p(\mathbf{x}; \Gamma) d\mathbf{x}$ $\mathbf{y} = \mathbf{\Phi} \mathbf{x} + \mathbf{v}$ $p(\mathbf{y}; \Gamma) = \mathcal{N} \left(0, \underbrace{\sigma^2 \mathbf{I} + \mathbf{\Phi} \Gamma \Phi^T}_{\Sigma_{\mathbf{y}}} \right)$

SBL cost function:

$$\mathcal{L}(\Gamma) \propto -\log \det(\Sigma_{\mathbf{y}}) - \mathbf{y}^T \Sigma_{\mathbf{y}}^{-1} \mathbf{y}$$

Optimization via EM

Log likelihood of the complete data $-\log p(\mathbf{y}, \mathbf{x}; \gamma) = \frac{\|\mathbf{y} - \Phi \mathbf{x}\|_2^2}{2\sigma^2} + \frac{1}{2} \left| \sum_{i=1}^N \frac{x_i^2}{\gamma_i} + \log \gamma_i \right|$ $-\log p(\mathbf{y}|\mathbf{x}; \gamma)$ $-\log p(\mathbf{x}; \gamma)$ indep. of γ func. of γ E-Step: compute "Q-function" $Q\left(\Gamma|\Gamma^{(t)}\right) = \mathbb{E}_{\mathbf{x}|\mathbf{y};\Gamma^{(t)}}\left[-\log p(\mathbf{y},\mathbf{x};\Gamma)\right]$ from previous iteration $\doteq \sum_{i=1}^{N} \frac{\mathbb{E}(x_i^2 | \mathbf{y}; \Gamma^{(t)})}{\gamma_i} + \log \gamma_i$ • Easy to compute: $p(x_i|\mathbf{y};\Gamma^{(t)})$ is Gaussian
The EM Ilterations

• E-step (continued):
$$p(\mathbf{x}|\mathbf{y};\Gamma^{(t)}) = \mathcal{N}(\mu,\Sigma)$$

 $\Sigma = \left(\sigma^{-2}\Phi^{T}\Phi + \left(\Gamma^{(t)}\right)^{-1}\right)^{-1} \quad \mu = \sigma^{-2}\Sigma\Phi^{T}\mathbf{y}$

• M-step: maximize Q($\Gamma | \Gamma^{(t)}$) given posteriors gathered in the E-step: $\Gamma^{(t+1)} = \arg \max_{\gamma_i \ge 0} Q\left(\Gamma | \Gamma^{(t)}\right) = \operatorname{diag}\left(\mu_i^2 + \Sigma_{ii}\right)$

Component-wise updates

The SBL Algorithm

1. Initialize $\Gamma = I$

2. Compute $\Sigma = \left(\sigma^{-2}\Phi^T\Phi + \left(\Gamma^{(t)}\right)^{-1}\right)^{-1}$ $\mu = \sigma^{-2}\Sigma\Phi^T\mathbf{y}$

- 3. Update $\Gamma^{(t+1)} = \operatorname{diag}\left(\mu_i^2 + \Sigma_{ii}\right)$
- 4. Repeat steps 2 and 3
- 5. Output µ after convergence

Empirical Example

- Generate random 50 x 100 matrix A
- Generate sparse vector x_o
- Compute y = Axo
- Solve for x_o, average over 1000 trials
- Repeat for different sparsity values



Highly scaled nonzero entries

Convergence

- Convergence guaranteed to a fixed pt. of L from any initialization (property of EM)
- The global min of L occurs at the sparsest solution in the noiseless case [wipf et al. 04]
- All local minima occur at sparse solutions in the noisy case [wipf et al. 04]
- Strictly better than MAP estmn. with a factorial prior [Wipf and Nagarajan 09]
 - Always has fewer local minima
 - Global min. at global optimum of lo min.

Other Options

- McKay updates [Tipping, 2001]
 - Set gradient of SBL cost = 0
 - Faster convergence than EM
- Greedy approach:
 - Update hyperparams one at a time [Tipping & Faul, 2003]
 - Closed-form update for each hyperparam
 - Fast, but can get trapped in a local min.
 - S Fast Bayesian matching pursuit [schniter et al., 08]

Other Options

Use dual-form of SBL. Cost function:

$$\mathbf{x}_{\text{opt}} = \arg\min_{\mathbf{x}} \|\mathbf{y} - \mathbf{\Phi}\mathbf{x}\|_{2}^{2} + \sigma^{2}g_{\text{SBL}}(\mathbf{x})$$

 $g_{\rm SBL}(\mathbf{x}) \triangleq \min_{\gamma \ge 0} \mathbf{x}^T \Gamma^{-1} \mathbf{x} + \log \det \left(\sigma^2 \mathbf{I} + \Phi \Gamma \Phi^T \right)$

Facilitates iterative reweighted L₁ and L₂ algorithms [Wipf and Nagarajan, 09]

Replace E-step with an approx. posterior computation: AMP-SBL [Al-shoukairi and Rao 14]

Approximate Message Passing

AMP [Donoho, Maleki, Montanari 09]:

- Uses Loopy belief propagation + Gaussian approximations to solve LASSO
- Key advantage: Low complexity

In SBL:

- All Gaussian PDFs: approximation is not necessary
- Only need to track means and variances
- Can replace computationally expensive E-step with the AMP based iterations

Empirical Example

 \mathfrak{M}

N = 200, M = 100, K = 20, Gaussian measurement matrix



[Al-Shakouri et al., 14]

Advantages of SBL

- Averaging over x: fewer minima in p(y;y)
 Get an estimate of the error in recovery
- Allows for "exact inference"
- Versatile: γ can also be used to tie several params. together - easier to estimate
- Useful extensions: incorporate structure
 - Intra/inter-vector correlation
 - SBL allows the use of Kalman framework
 - Block/cluster sparsity
 - Colored noise (rank-deficient cov.)

To Recap

Sparse Bayesian Learning

- Sparse vector recovery via estimating hyperparameters
- Expectation-maximization iterations
- Convergence properties
- Alternative implementations

Limitations

- Computational complexity
 - More recent algos overcome this
- Slow convergence

Fast versions exist, but without the same convergence guarantees

Part 4: Joint Sparse Signal Recovery

One important variant of the sparse recovery problem

Outline

- The joint sparse signal recovery problem
- Lobound in support recovery
- Sparse Bayesian Learning (the MMV version)
 Performance guarantees and new insights
- New algorithms
 - Covariance matching framework
- New theoretical results
 - Restricted isometry of Khatri-Rao product

Joint Sparse Recovery Problem



Let k = number of nonzero rows in X.

Want to recover X or support(X) from the Multiple Measurement Vectors (MMVs) Y



Magnitude spectrum across secondary cell users is approximately jointly sparse.

Exploit structure to improve accuracy.

Multi-sensor Data



sensor-3 Signals acquired by different sensors have overlapping subspaces

> Approximate as different linear combinations of the same elementary signals

Why do subspaces overlap? Commonality of physical process being sensed Overlapping sensing regions

Generative Model for Multi-sensor Data

Simultaneous sparse approximation (SSA) [Tropp 04]



Overlapping subspace = $column-space(D_s)$.

Compression of Multisensor Data

 ∞

Linear encoder



high dimensional data from L sensors

S1 S2 SI



Two-Stage Decoder



Stage 1: Recover joint-sparse X_{est} from sketch Y Stage 2: $S_{est} = D X_{est}$

Support Recovery is also Important







Subspace filtering by projecting to common signal subspace

L Bound

Canonical Lo problem:

 (L_0) : $\min_{\mathbf{X}} \mathcal{R}(\mathbf{X})$ subject to $\mathbf{Y} = \Phi \mathbf{X}$.

 $\mathcal{R}(\mathbf{X}) = No. \text{ of nonzero rows in } \mathbf{X}$

• Unique k-sparse solution if [Chen & Huo, 06] $k < \frac{\text{Spark}(\Phi) - 1 + \text{Rank}(Y)}{2}$ (ℓ_0 - bound)

Spark(\$\Phi\$) = min. num. of lin. dep. cols in \$\Phi\$.
\$\mathbf{L}_0\$ bound on num. of nonzero rows: \$\mathbf{k} < m\$.

Support Recovery Beyond Lo Bound

Supports of size k > m are recoverable!
Key idea: Use correlation-aware priors

Support recovery phase transition



Sparse Bayesian Learning (M–SBL)





Correlation Awareness

Latent structure within joint sparse vectors!

Zero intra-vector correlation $\mathbb{E}\left[\mathbf{x}_{j}(i)\mathbf{x}_{j}(k)\right] \approx 0$



Correlation-aware prior [Pal & Vaidyanathan, 15] $\mathbf{x}_{i} \stackrel{i.i.d.}{\sim} \mathcal{N}(\mathbf{0}, \Gamma), \quad \Gamma = \operatorname{diag}(\gamma).$

MSBL-Sparse Bayesian Learning using MMVs

• Observation model: $\mathbf{Y} = \mathbf{\Phi}\mathbf{X} + \mathbf{W}$

Correlation-aware prior: x_j ^{i.i.d.} N(0, Γ), Γ = diag(γ)
 Common Γenforces same support in columns of X.
 Gaussian MMVs:

$$\mathbf{y}_j \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I} + \mathbf{\Phi} \mathbf{\Gamma} \mathbf{\Phi}^T)$$

- M-SBL algorithm: $\hat{\gamma} = \operatorname*{argmax}_{\gamma \in \mathbb{R}^n_+} \log p(\mathbf{Y}; \gamma)$
 - Nonconvex objective
 - Solved via Expectation Maximization (EM)
 - Estimated support = support($\hat{\gamma}$).

The M-SBL Algo

- Cost function $p(\mathbf{Y}; \gamma) = \int p(\mathbf{Y}, \mathbf{X}; \gamma) d\mathbf{X} = \prod_{j=1}^{L} \int p(\mathbf{y}_j | \mathbf{x}_j) p(\mathbf{x}_j; \gamma) d\mathbf{x}_j$
- Key point: y couples the sparsity pattern across xj
 - Fewer parameters to estimate: N << (N x L)</p>
- EM iterations E-step: $Q(\gamma|\gamma^k) = \mathbb{E}_{\mathbf{X}|\mathbf{Y},\gamma^k} [\log p(\mathbf{Y}, \mathbf{X}; \gamma)]$ M-step: $\gamma^{k+1} = \arg \max_{\gamma \in \mathbb{R}^N_+} Q(\gamma|\gamma^k)$ • Posterior distbn.: $p(\mathbf{x}_j|\mathbf{y}_j;\gamma^k) \sim \mathcal{N}(\mu_j^{k+1}, \Sigma_j^{k+1})$

E & M Steps

E Step: $\Sigma_{j}^{k+1} = \Gamma^{k} - \Gamma^{k} \Phi_{j}^{T} \left(\sigma_{j}^{2} \mathbf{I}_{M} + \Phi_{j} \Gamma^{k} \Phi_{j}^{T} \right)^{-1} \Phi_{j} \Gamma^{k}$ $\mu_i^{k+1} = \sigma_i^{-2} \Sigma_i^{k+1} \Phi_i^T \mathbf{y}_j$ M Step: $\gamma^{k+1}(i) = \frac{1}{L} \sum_{i=1}^{L} \mu_j^{k+1}(i)^2 + \Sigma_j^{k+1}(i,i)$ i=1Average of the individual estimates of yi across measurements

Performance of MSBL

Support recovery phase transition

 \mathfrak{M}



n=200, L=400, SNR=20 dB

Recoverable support size k grows as O(m²)!

Part 5: Performance Guarantees for Sparse Bayesian Learning

Sufficient conditions for support recovery by M-SBL

Sufficient Conditions for Support Recovery in SBL

Single measurement vector (L = 1)
Noiseless observations
Result: SBL correctly recovers the support for all 1 ≤ k < spark(Φ) - 1
spark: min. number of lin. dep. cols
Usually, in CS, spark(Φ) = m + 1
For L1 recovery, 1 ≤ k ≤ 0(m / log N)

Sufficient Conditions for Support Recovery in MSBL

Suppose X1, X2, ..., XL have common support S* of size k. Nonzero entries i.i.d. zero mean Gaussian with variance in [γ_{min}, γ_{max}]. Then,

 $\mathbb{P}\left(\operatorname{supp}(\hat{\gamma}) \neq \mathcal{S}^*\right) \le \exp\left(-\frac{\eta L}{8}\right)$

Under Conditions 1 & 2 (next slide)

[Khanna and M., CoRR abs/1703.04930 (2017)]

Conditions 1 & 2

• Condition 1: The self Khatri-Rao product matrix $\Phi \odot \Phi$ has a strictly positive minimum singular value

Condition 2:

$$\eta \geq \inf_{\gamma \in \Theta_K} \frac{\|(\boldsymbol{\Phi} \odot \boldsymbol{\Phi})(\gamma - \gamma^*)\|_2^2}{\left(\sigma^2 + \gamma_{\max} \|\boldsymbol{\Phi}_{\mathcal{S} \cup \mathcal{S}^*}^T \boldsymbol{\Phi}_{\mathcal{S} \cup \mathcal{S}^*}\|_2\right)^2}$$

Proof Outline

- Error event \$\mathcal{E}_{S^*} = \bigcup_{S \in S_K \setminus \{S^*\}} \biggl\{ \mathbf{max}_{\gamma \in \Theta(S)} \mathcal{L}(Y; \gamma) \ge \mathbf{max}_{\gamma \in \Theta(S^*)} \mathcal{L}(Y; \gamma) \biggr\}\$
 Apply union bound, replace first max. by a finite sized union over an epsilon net
- Use a large deviation property on the likelihood function to bound error prob.

$$\mathbb{P}(\mathcal{E}_{\mathcal{S}^*}) \leq \sum_{\mathcal{S} \in \mathcal{S}_K \setminus \mathcal{S}^*} \sum_{\gamma \in \Theta^{\epsilon}(\mathcal{S})} \exp\left(-L\mathcal{D}_{1/2}^*/4\right)$$

- Lower bound the worst case exponent $\mathcal{D}_{1/2}^* \geq \eta$
- Hence

 𝒫(𝔅_{𝔅*}) ≤ exp (−L (η/4 − log |𝔅_𝑘| − log(max_{𝔅∈𝔅_𝑘} |Θ^ϵ(𝔅)|))/L)

 Bound the cardinality of the epsilon net

 Lipschitz const. of the log likelihood

when will Condition 1 hold?

• Theorem (paraphrased): Suppose Φ has real i.i.d. Gaussian(0,1) entries. Then,

 $\mathbb{P}\left(\delta_k\left(\frac{\Phi}{\sqrt{m}}\odot\frac{\Phi}{\sqrt{m}}\right) \ge \delta\right) \le 5n^{-2(\beta-1)}$

for all $\beta > 1$, provided $m \ge 4c_3\beta\left(\frac{k\log n}{\delta}\right)$ If A has i.i.d. Gaussian entries, $\mathbb{P}\left(\delta_k\left(\frac{\mathbf{A}}{\sqrt{m}}\right) > \delta\right) \leq \frac{1}{n^{\alpha}}$ provided $m \geq \frac{c}{\delta^2}(k+\alpha)\log n$ [Foucart & Rauhut, Thm. 9.27]

Proof Outline

Point of departure
$$\delta_k \left(\frac{\Phi}{\sqrt{m}} \odot \frac{\Phi}{\sqrt{m}} \right) = \sup_{\mathbf{z} \in \mathbb{R}^n, \|\mathbf{z}\|_2 = 1, \|\mathbf{z}\|_0 \le k} \left\| \left\| \left(\frac{\Phi}{\sqrt{m}} \odot \frac{\Phi}{\sqrt{m}} \right) \mathbf{z} \right\|_2^2 - 1 \right\|$$

- Union bound $\Pr(\mathsf{RHS} \ge \delta)$. Use $(\Phi \odot \Phi)^T (\Phi \odot \Phi) = \Phi^T \Phi \circ \Phi^T \Phi$
- Then use Gaussian tail bounds and the Hanson Wright inequality: $\mathbb{P}\left\{ |\mathbf{x}^T \mathbf{A} \mathbf{x} - \mathbb{E} \mathbf{x}^T \mathbf{A} \mathbf{x}| > t \right\}$ $\leq 2 \exp \left[-c \min \left(\frac{t^2}{K^4 ||\mathbf{A}||_{HS}^2}, \frac{t}{K^2 ||\mathbf{A}||} \right) \right]$

Implications

- In MSBL, for a fixed n, we get perfect support recovery w.h.p., with noisy meas.
 \$\overline{k}\$ even if \$k > m\$
 - For $1 \le k \le \operatorname{krank}(\Phi \odot \Phi)/2$
 - * krank($\Phi \odot \Phi$): largest p s.t. any p cols of $\Phi \odot \Phi$ are linearly independent
 - For suitable Φ , $\operatorname{krank}(\Phi \odot \Phi) = \mathcal{O}(m^2)$
 - Sparsity level up to O(m²) is potentially recoverable!

¹P. Pal and P. P. Vaidyanathan, "Correlation-Aware Techniques for Sparse Support Recovery", SSP Workshop, 2012.

New Interpretation of M-SBL Cost Function

$$-\log p(\mathbf{Y}; \gamma) = -\sum_{j=1}^{L} \log \mathcal{N} \left(\mathbf{y}_{j}; 0, \sigma^{2} \mathbf{I}_{m} + \mathbf{\Phi} \Gamma \mathbf{\Phi}^{T} \right)$$
$$\propto \log |\sigma^{2} \mathbf{I}_{m} + \mathbf{\Phi} \Gamma \mathbf{\Phi}^{T}| + \operatorname{tr} \left(\left(\sigma^{2} \mathbf{I}_{m} + \mathbf{\Phi} \Gamma \mathbf{\Phi}^{T} \right)^{-1} \left(\frac{1}{L} \mathbf{Y} \mathbf{Y}^{T} \right) \right)$$
$$\propto \mathcal{D}_{-\log \det}^{\operatorname{Bregman}} \left(\frac{1}{L} \mathbf{Y} \mathbf{Y}^{T}, \sigma^{2} \mathbf{I}_{m} + \mathbf{\Phi} \Gamma \mathbf{\Phi}^{T} \right) + \operatorname{const. terms}$$

 Motivates covariance matching based approaches to sparse recovery
 Can we use some other divergence?


Also, $\gamma \geq 0$. Want a non-negative null space property

Empirical Study

 ∞

 $\sim N(0, 1/m) \notin m = 0.5n$ Aij, Bij

2-RIC





RIP improved by taking Khatri-Rao product!

Deterministic RIC bound for Khatri-Rao product

Theorem: [Khanna & M., TSP 2018]
For m × n sized matrices A and B with unit norm columns,

 $\delta_k (\mathbf{A} \odot \mathbf{B}) \leq [\max (\delta_k (\mathbf{A}), \delta_k (\mathbf{B}))]^2$

Proof technique:
 (A \circ B)^T(A \circ B) = A^TA \circ B^TB
 Kantorovitch matrix inequalities

• Useful corollary: $\delta_k(\mathbf{A} \odot \mathbf{A}) \leq (\delta_k(\mathbf{A}))^2 < \delta_k(\mathbf{A}) \lim_{\text{improves}} \delta_k(\mathbf{A}) \leq (\delta_k(\mathbf{A}))^2 < \delta_k(\mathbf{A})$

Part 6: New Algorithms



Covariance matching is the key!

Covariance Matching framework

 ∞

• Observation
Model:

$$\mathbf{Y} = \mathbf{\Phi}\mathbf{X} + \mathbf{W}$$

• Principle:
 $\hat{\gamma} = \arg\min_{\gamma \in \mathbb{R}^n_+} \operatorname{dist}\left(\begin{bmatrix} \mathbf{1}\\ L \mathbf{Y}\mathbf{Y}^T \\ L \end{bmatrix}, \sigma^2 \mathbf{I}_m + \mathbf{\Phi}\Gamma \mathbf{\Phi}^T \right)$
Empirical covariance matrix

Support estimate = Support(
$$\hat{\gamma}$$
)

Covariance Matching Algorithm 1

- Distance = Frobenius norm
- Algorithm = CO-LASSO [Pal & Vaidyanathan, 15]

$$\hat{\gamma} = \arg\min_{\boldsymbol{\gamma}\in\mathbb{R}^n_+} \left\| \left\| \frac{1}{L} \mathbf{Y} \mathbf{Y}^T - (\sigma^2 \mathbf{I} + \mathbf{\Phi} \mathbf{\Gamma} \mathbf{\Phi}^T) \right\| \right\|_F^2 + \lambda \left\| \boldsymbol{\gamma} \right\|_1$$

- Convex objective
- High memory requirements (to store $\Phi \odot \Phi$)

Covariance Matching Algorithm 2

- Distance = Log-Det Bregman Matrix Divergence
- Algorithm = M-SBL [Wipf & Rao, 07]

$$\hat{\gamma} = \underset{\boldsymbol{\gamma} \in \mathbb{R}^{n}_{+}}{\arg\min} \log \left| \sigma^{2} \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^{\mathcal{T}} \right| + \operatorname{tr} \left(\left(\sigma^{2} \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^{\mathcal{T}} \right)^{-1} \left(\frac{1}{L} \mathbf{Y} \mathbf{Y}^{\mathcal{T}} \right) \right)$$

- Nonconvex objective (optimize using EM)
- Slow convergence
- Best performance

Covariance Matching Algorithm 3

 Distance = Log-Det Jensen Difference
 Algorithm = Rényi Divergence based Covariance Matching Pursuit (RD-CMP) [Khanna & M., 17]
 A new correlation-aware prior x_j ~ N(0, γdiag(1_S))
 Covariance matrix parameterized by support S
 Induces Gaussian distributed MMVs
 y_i ~ N(0, σ²I_m + γΦ_SΦ^T_S)

Covariance matching with α-Rényi divergence

$$\hat{\mathcal{S}} = \operatorname*{argmin}_{\mathcal{S}\subseteq[n]} \mathcal{D}_{\alpha} \left(\mathcal{N} \left(\mathbf{0}, \frac{1}{L} \mathbf{Y} \mathbf{Y}^{T} \right), \mathcal{N} \left(\mathbf{0}, \sigma^{2} \mathbf{I}_{m} + \gamma \Phi_{\mathcal{S}} \Phi_{\mathcal{S}}^{T} \right) \right)$$

Rényi Divergence based Covariance Matching Pursuit

Covariance matching using α-Rényi Divergence

$$\hat{\mathcal{S}} = \operatorname*{argmin}_{\mathcal{S}\subseteq[n]} \mathcal{D}_{\alpha} \left(\mathcal{N} \left(\mathbf{0}, \frac{1}{L} \mathbf{Y} \mathbf{Y}^{T} \right), \mathcal{N} \left(\mathbf{0}, \sigma^{2} \mathbf{I}_{m} + \gamma \Phi_{\mathcal{S}} \Phi_{\mathcal{S}}^{T} \right) \right)$$

 $\hat{\mathcal{S}} = \underset{\mathcal{S}\subseteq[n]}{\operatorname{argmin}} \underbrace{\log \left| (1-\alpha) \frac{1}{L} \mathbf{Y} \mathbf{Y}^{T} + \alpha \left(\sigma^{2} \mathbf{I} + \gamma \Phi_{\mathcal{S}} \Phi_{\mathcal{S}}^{T} \right) \right|}_{f(\mathcal{S}), \text{ submodular in } \mathcal{S}} - \underbrace{\alpha \log \left| \sigma^{2} \mathbf{I} + \gamma \Phi_{\mathcal{S}} \Phi_{\mathcal{S}}^{T} \right|}_{g(\mathcal{S}), \text{ submodular in } \mathcal{S}}$

Generalizes M-SBL cost
 RD-CMP objective = difference of submodular funcs.

Submodular functions: a primer

- V = ground set of elements
- Set func. $f: \mathcal{V} \to \mathbb{R}_+$ submodular if, for $S \subseteq \mathcal{T} \subseteq \mathcal{V}_+$
 - f is Monotonic: $f(S) \leq f(T)$ f follows Law of Diminishing Returns $f(\mathcal{T} \cup \{a\}) - f(\mathcal{T}) \leq f(\mathcal{S} \cup \{a\}) - f(\mathcal{S}) \quad \forall a \in \mathcal{V} \setminus \mathcal{T}$
- Examples:
 - Rank of a matrix

 - Joint entropy $f(S) = \log |\mathbf{A} + \gamma \mathbf{B}_S \mathbf{B}_S^T|$ is submodular in S for A, γ >0

Submodular functions: a primer

- Maximizing a submodular function subject to cardinality constraints is easy!
 - A Greedy algorithm will maximize a submodular function f to within $\left(1-\frac{1}{e}\right) f_{opt}$
- Tight modular upper bnd for submodular f

$$f(\mathcal{S}) \leq h_{\mathcal{X}}^{f}(\mathcal{S}) \triangleq f(\mathcal{X}) - \sum_{j \in \mathcal{X} \setminus \mathcal{S}} \left[f(\mathcal{X}) - f(\mathcal{X} \setminus \{j\}) \right] + \sum_{j \in \mathcal{S} \setminus \mathcal{X}} \left[f(j) - f(\phi) \right]$$

• RD-CMP objective $\hat{S} = \underset{S \subseteq [n]}{\operatorname{argmin}} \underbrace{\log \left| (1 - \alpha) \frac{1}{L} \mathbf{Y} \mathbf{Y}^{T} + \alpha \left(\sigma^{2} \mathbf{I} + \gamma \Phi_{S} \Phi_{S}^{T} \right) \right|}_{f(S), \text{ submodular in } S} - \underbrace{\alpha \log \left| \sigma^{2} \mathbf{I} + \gamma \Phi_{S} \Phi_{S}^{T} \right|}_{g(S), \text{ submodular in } S}$

• Find \hat{S} via majorization-minimization

- Majorization
 - Replace 1st log-det term f(S) by its modular upper bound h_{st}(S): tight at St
- Minimization

$$\mathcal{S}_{t+1} = \underset{\mathcal{S}\subseteq[n]}{\operatorname{arg\,min}} \quad \underbrace{h_{\mathcal{S}_t}(\mathcal{S}) - \alpha \log \left| \sigma^2 \mathbf{I} + \gamma \Phi_{\mathcal{S}} \Phi_{\mathcal{S}}^T \right|}_{\mathcal{S}_t}$$

Supermodular func., minimized by greedy search

RD-CMP Performance

Co-LASSO









SNR = 10 dB; n = 200; L = 200

SNR = 10 dB; k = 50 log n m = 0.75 k, mL = 50 k log n

RD-CMP currently the FASTEST covariance matching based MMV solver! 85



RD-CMP Performance

SNR=20 dB, n=500, k=200, m =100

 \mathfrak{M}

Support detection rate





RD-CMP performs better than Co-LASSO but slightly worse than M-SBL

Support Recovery via Nonnegative Parameter Estimation

Covariance matching principle:

find sparse, nonnegative diagonal Γ s.t.

$$\frac{1}{L}\mathbf{Y}\mathbf{Y}^{T}\approx\sigma^{2}\mathbf{I}_{m}+\mathbf{\Phi}\mathbf{\Gamma}\mathbf{\Phi}^{T}$$

Example: Recover Γ via nonnegative LASSO

$$\hat{\gamma} = \arg\min_{\boldsymbol{\gamma} \in \mathbb{R}^n_+} \left\| \left\| \frac{1}{L} \mathbf{Y} \mathbf{Y}^T - (\sigma^2 \mathbf{I} + \boldsymbol{\Phi} \boldsymbol{\Gamma} \boldsymbol{\Phi}^T) \right\| \right\|_F^2 + \lambda \left\| \boldsymbol{\gamma} \right\|_1$$

Non-negative Least Squares Based Approach

- We write $\hat{\Sigma} = \frac{1}{L} \mathbf{Y} \mathbf{Y}^T = \Sigma + \mathbf{E}$
- In the noiseless case, $\Sigma = \mathbf{\Phi} \Gamma \mathbf{\Phi}^T$
- Noisy case: L. Ramesh and M. ICASSP 18
 Vectorize: r ≜ vec(Σ̂) = (Φ ⊙ Φ)γ + e
 Model e as Gaussian distributed with mean zero and covariance W ≜ cov(e) = 1/L(Φ ⊗ Φ)(Γ^{1/2} ⊗ Γ^{1/2})B(Γ^{1/2} ⊗ Γ^{1/2})(Φ ⊗ Φ)^T
 Here, B = cov(zz^T) where z is N(0, I_n). Can be found in closed form

ML estimation of γ

• We seek to solve: $\gamma_{\mathrm{ML}} = rg \max_{\gamma>0} p(\mathbf{r};\gamma)$ Optimization problem: $\gamma_{\text{ML}} = \arg \min_{\gamma \ge 0} \log |\mathbf{W}| + (\mathbf{r} - \mathbf{A}\gamma)^{\top} \mathbf{W}^{-1} (\mathbf{r} - \mathbf{A}\gamma)$ For a given W, it is a non-negative weighted least-squares problem Can solve using non-negative quadratic programming (NNQP) Reweighted minimization procedure: (1) Compute W, (2) solve an NNQP; iterate

Performance

 \mathfrak{M}



N = 40, M = 20, k = 25 N = 70, M = 20, L = 50, 1000 N = 20, L = 200

Non-negativity Encourages Sparsity!

- $\min_{\mathbf{x}} \|\mathbf{x}\|_{0} \qquad (P_{0^{+}}) \qquad \min_{\mathbf{x}} \|\mathbf{y} \Phi \mathbf{x}\|_{2}^{2} \qquad (NNQP)$ s. t. $\mathbf{y} = \Phi \mathbf{x}, \ \mathbf{x} \ge 0 \qquad$ s. t. $\mathbf{x} \ge 0$
- When do the two yield the same soln?
- [Uniqueness]: x₀ k-sparse & y = Φ x₀. Then, x₀ is unique sol. to (P₀+) iff every v ∈ ker(Φ)\{0} has at least (k+1) +ve or (k+1) -ve entries
- [Recoverability via NNQP]: x₀ is exactly recovered by NNQP if every $v \in ker(\Phi) \setminus \{0\}$ has at least $(k+1) + v_{k}$ and $(k+1) v_{k}$ entries

Non-Negative Parameter Estimation Framework

Minimize a convex loss $L: S^n_+ \to \mathbb{R}_+$

[Tsuda, Rätsch and Warmuth, Matrix Exponentiated Gradient
 Updates for On-Line Learning and Bregman Projection, 2005]

$$\begin{split} \mathbf{W}_{t+1} &= \arg\min_{\mathbf{W}\in S^n_+} \mathcal{D}_F(\mathbf{W},\mathbf{W}_t) + \eta L(\mathbf{W}) \\ \mathbf{W}\in S^n_+ \mathcal{D}_F(\mathbf{W},\mathbf{W}_t) &= \mathbf{M}_F(\mathbf{W},\mathbf{W}_t) \\ \mathbf{W}\in S^n_+ \mathcal{D}_F(\mathbf{W},\mathbf{W}_t) \\ \mathbf{W}\in S^n_+ \mathcal{D$$

Bregman Matrix Divergence

Bregman matrix divergence

$$\mathcal{D}_{F}(\mathbf{W},\mathbf{W}_{t})=F(\mathbf{W})-F(\mathbf{W}_{t})-\mathrm{tr}\left(f(\mathbf{W}_{t})^{T}(\mathbf{W}-\mathbf{W}_{t})\right)$$

first order approx. of $F(\mathbf{W})$ around \mathbf{W}_t

- $f = \nabla F$
- Seed function F: strictly convex and differentiable
- Choice of F
 F(W) = -log det (W)
 F(W) = tr(W logW)



MEG Updates

• $F(\mathbf{W}) = -\log \det \mathbf{W}$ $\mathcal{D}_F(\mathbf{W}, \mathbf{W}_t) = \log \frac{|\mathbf{W}_t|}{|\mathbf{W}|} + \operatorname{tr} \left(\mathbf{W}_t^{-1}\mathbf{W}\right) - n$ Log-Det Bregman Matrix divergence

MEG update:
$$\mathbf{W}_{t+1} = \left((\mathbf{W}_t)^{-1} + \eta \nabla L(\mathbf{W}_t) \right)^{-1}$$

 $F(\mathbf{W}) = \operatorname{tr}(\mathbf{W} \log \mathbf{W})$

 $\mathcal{D}_{\mathsf{F}}(\mathbf{W}, \mathbf{W}_{t}) = \mathsf{tr}(\mathbf{W} \log \mathbf{W} - \mathbf{W} \log \mathbf{W}_{t} - \mathbf{W} + \mathbf{W}_{t})$ Von-Neumann Matrix divergence

MEG update: $\mathbf{W}_{t+1} = \exp(\log \mathbf{W}_t - \eta (\nabla L(\mathbf{W}_t)))$

• Loss $L(\Gamma): \left| \left| \left| \mathbf{R}_{yy} - \mathbf{\Phi} \mathbf{\Gamma} \mathbf{\Phi}^{T} \right| \right| \right|_{F}^{2} + \lambda \left| |\gamma| \right|_{1}$

- Parameter space: all p.s.d. diagonal matrices
- Log-Det divergence based MEG update $\gamma_{t+1}(i) = \gamma_t(i) \left(\frac{1}{1 + 2\eta\gamma_t(i) \left[\phi_i^T (\Phi \Gamma \Phi^T - \mathbf{R}_{yy}) \phi_i + \lambda\right]} \right), \quad \forall i \in [n].$
 - Von-Neumann divergence based MEG update $\gamma_{t+1}(i) = \gamma_t(i) \cdot e^{-2\eta \left[\phi_i^T \left(\Phi \Gamma \Phi^T - \mathsf{R}_{yy}\right)\phi_i + \lambda\right]}, \quad \forall i \in [n].$

Performance

 \mathfrak{M}

support recovery phase transition

 $(n = 200, L = 400, SNR=20dB, \lambda = 0.25, \eta = 0.5)$

Log-Det divergence based MEG Von-Neumann divergence based MEG





Part 7: Other Extensions



- 1. Cluster-sparsity, inter-vector correlation
- 2. Online sparse signal recovery
- 3. Distributed sparse signal recovery

Block Sparsity & Intra-Block Correlation

Intra-vector correlation is often present, and is important to model & exploit v Φ $\begin{array}{c} x_{1} \\ x_{2} \\ \\ noise \\ \mathcal{N}(0, \sigma^{2}\mathbf{I}_{M}) \end{array}$

 χ_g

sparse

signal

- g blocks; few nonzero
- Intra-block correlation

• Measurement model: $\mathbf{y} = \Phi \mathbf{x} + \mathbf{v}$ $\mathbf{x} = [\underbrace{x_1, \dots, x_{d_1}}_{\mathbf{x}_1^T}, \dots, \underbrace{x_{d_{g-1}+1}, \dots, x_{d_g}}_{\mathbf{x}_g^T}]^T$

Parameterized prior

 $p(\mathbf{x}_i; \gamma_i, \mathbf{B}_i) \sim \mathcal{N}(0, \gamma_i \mathbf{B}_i), i = 1, 2, \dots, g$

• γ_i controls sparsity

B_i controls intra-block correlation

Optimization Problem

• Posterior distribution $p\left(\mathbf{x}|\mathbf{y};\sigma^{2},(\gamma_{i}\mathbf{B}_{i})_{i=1}^{g}\right) \sim \mathcal{N}(\mu_{x},\Sigma_{x})$

• where
$$\mu_x = \Sigma_0 \Phi^T (\sigma^2 \mathbf{I} + \Phi \Sigma_0 \Phi^T)^{-1} \mathbf{y}$$

 $\Sigma_x = \Sigma_0 - \Sigma_0 \Phi^T (\sigma^2 \mathbf{I} + \Phi \Sigma_0 \Phi^T)^{-1} \Phi \Sigma_0$
 $\Sigma_0 = \operatorname{diag}(\gamma_1 \mathbf{B}_1, \dots, \gamma_g \mathbf{B}_g)$

• All params. can be estimated by maximizing: $\mathcal{L}(\Theta) = -2\log \int p(\mathbf{y}|\mathbf{x}; \sigma^2) p(\mathbf{x}; \Sigma_0) d\mathbf{x}$ $= \log \det \left(\sigma^2 \mathbf{I} + \Phi \Sigma_0 \Phi^T\right) + \mathbf{y}^T \left(\sigma^2 \mathbf{I} + \Phi \Sigma_0 \Phi^T\right)^{-1} \mathbf{y}$

Several Options for Optimization

- BSBL-EM: Use expectation-maximization
- BSBL-BO: Use bounded optimization, i.e.,
 a form of majorization-minimization
- BSBL-11: Use a reweighted 11 procedure (special case of BSBL-B0)
- Different strategies offer a variety of performance-complexity tradeoffs

Phase Transition



Correlation = 0

Correlation = 0.95



N = 1000, M = δ N, g = 40, block size = 25 Curves indicate > 99% success

[Zhang et al. 2013]

Delay-Constrained Sparse Vector Recovery

 Temporally correlated sparse vectors with a

common support

- Goal: Given y1:k estimate xk-A
- Max. delay constraint between measurement and estimation



Offline approach:

M-step: Compute $\Gamma^{(r+1)}$ as a closed-form function of the state statistics

Mean: $\hat{\mathbf{x}}_{t|K} = \mathbb{E}\{\mathbf{x}_t | \mathbf{y}_{1:K}\}$

Autocov.: $\mathbf{P}_{t|K} = \operatorname{cov}\{\mathbf{x}_t, \mathbf{x}_t | \mathbf{y}_{1:K}\}$

Cross-cov: $\mathbf{P}_{t,t-1|K} = \operatorname{cov}\{\mathbf{x}_t, \mathbf{x}_{t-1} | \mathbf{y}_{1:K}\}\$ E-step: Compute state statistics using fixed interval Kalman smoothing

Online Version

 Key idea: Update γ once as each input arrives and compute sparse vector
 Online recursion:

Non-iterative!

 $\gamma_k = \gamma_{k-1} + \frac{1}{k} \operatorname{Diag} \left\{ \left(\mathbf{I} - \mathbf{D}^2 \right)^{-1} \left(\mathbf{T}_{k|k+\Delta} - \Gamma_{k-1} \right) \right\}$

fcn. of $\hat{\mathbf{x}}_{t|t+\Delta}, \mathbf{P}_{t|t+\Delta}, \mathbf{P}_{t,t-1|t+\Delta}$

- Two implementations of the Kalman filter:
 - fixed lag smoothing
 - sawtooth lag smoothing
- Main advantage: reduced computational and memory costs [G. Joseph, M., TSP 2017]

Convergence Analysis

Assume D=0 (uncorrelated sparse vecs)
Simplified algorithm: γ_k = γ_{k-1} + ¹/_kdiag{P(γ_{k-1}) + x̂(γ_{k-1})x̂(γ_{k-1})^T - Γ_{k-1}} P(γ) = Γ - ΓΦ^T (ΦΓΦ^T + σ²I)⁻¹ΦΓ x̂(γ) = P(γ)Φ^Ty_k/σ²
Stochastic approximation recursion γ_k = γ_{k-1} + ¹/_kf(γ_{k-1}) + ¹/_ke_k

 $\begin{array}{l} \text{Mean field function} & \mathbf{e}_{k} = \operatorname{diag} \left\{ \mathbf{P}(\gamma_{k-1}) + \hat{\mathbf{x}}(\gamma_{k-1}) \hat{\mathbf{x}}(\gamma_{k-1})^{T} \right\} - \gamma_{k-1} - \mathbf{f}(\gamma_{k-1}) \\ \mathbf{f}(\gamma) = \mathbb{E} \left\{ \operatorname{diag} \left(\mathbf{P}(\gamma) + \hat{\mathbf{x}}(\gamma) \hat{\mathbf{x}}(\gamma)^{T} - \Gamma \right) \right\} \\ \end{array}$

Convergence Result



If Rank $\{ \boldsymbol{\Phi} \odot \boldsymbol{\Phi} \} = N$, then $\gamma_k \to \gamma \in \{ \boldsymbol{0}, \gamma_{\text{opt}} \} a$. s.

- Result independent of
 - Sparsity level
 - Initialization
 - Distribution of sparse vecs
 - Restricted isometry properties of Φ
 - Noise level

Performance

 \mathfrak{m}



m = 20, N = 60, sparsity = 6, SNR = 20 dB
Dictionary Learning

 \mathfrak{M}

Matrix factorization problem:



SBL framework for DL

- Type-II ML: solve $\max_{\mathbf{\Lambda}=(\mathbf{\Phi},\mathbf{\Gamma})} -\log p\left(\mathbf{Y};\mathbf{\Lambda}\right)$
- EM procedure:
 - E-step: update statistics of X, as before
 - M-step: separable in variables Φ, Γ
 - Closed-form update for Γ
 - Non-convex in Φ
 - Alternating minimization (AM):
 update one column of Φ at a time

M-Step Analysis

• Cost function for updating Φ : $g(\Phi) = -\operatorname{Tr} \{ \mathbf{M} \mathbf{Y}^T \Phi \} + \frac{1}{2} \operatorname{Tr} \{ \Phi(\mathbf{\Sigma} - \mathcal{D}\{\mathbf{\Sigma}\}) \Phi^T \}$

• Proposition: The sequence of matrices generated by AM converges to a fixed point of $g(\Phi)$, and every fixed point is a Nash equilibrium, i.e., $g(\Phi_1^*, \dots, \Phi_{i-1}^*, \Phi_i^*, \Phi_{i+1}^*, \dots, \Phi_N^*) \leq g(\Phi_1^*, \dots, \Phi_{i-1}^*, \mathbf{a}, \Phi_{i+1}^*, \dots, \Phi_N^*), \forall i \in [N]$ for any unit norm a

Image Denoising Example



(a) Original image

(d) SimCO, PSNR = 28.64 dB,

(g) SGK, PSNR = 27.44 dB,

run time = 82.5 s

run time = 58.7 s



(b) Corrupted image, PSNR = 20 dB



(c) DL-SBL, PSNR = 28.96 dB, run time = 105.7 s



(e) DL-MM, PSNR = 28.54 dB, run time = 98.7 s



(h) PAU, PSNR = 27.44 dB, run time = 84.5 s



(f) KSVD, PSNR = 28.34 dB,

run time = 76.7 s

(i) MOD, PSNR = 27.42 dB, run time = 79.2 s

- 512 x 512 image "Barbara"
- Goal: remove AWGN
- Learn dictionary using 1000
 8 x 8 blocks, randomly chosen
- N = 256
- Learn dictionary
- Reconstruct image using OMP

Sparsity and Linear Dynamical Systems

• System Model: $\mathbf{x}_{k+1} = \mathbf{D}\mathbf{x}_k + \mathbf{H}\mathbf{h}_k$ $\mathbf{y}_k = \mathbf{A}_k\mathbf{x}_k$

- Want to observe, control and stabilize linear dynamical systems under sparsity constraints
 Examples:
 - Known inputs: recover sparse state
 - Unknown sparse inputs: recover state
 and inputs

Applications

- Diffusion processes with sparse initialization
 - Disease/epidemic spreading
 - Pollution
 - Computer/mobile nwk virus spreading
 Info. propagation in social networks
- Identifying the initial state of the system is critical to control the spreading
- Goal: Observability of the system when initialized with a sparse xo

Compressed Sensing Formulation



- Challenges:
 - Non-identically distributed rows
 - Columns not independent
 - D can be arbitrary

Independent Subgaussian Ak

RIP condition: The RIC δ_s of $\tilde{A}_{(K)}$ satisfies $\delta_s < \delta$ for all $\delta > 1 - \lambda^{2(K-1)}$ with probability at least $1-\epsilon$ if $Km\left(\delta - 1 + \lambda^{2(K-1)}\right)^2 \ge \tilde{c} \left|9\log\left(\frac{eN}{s}\right) + 2\log(2\epsilon^{-1})\right|$ • $\lambda \leq 1$ is the ratio of the smallest to largest singular value of D System observable if $\delta_s < \text{threshold}$ Total number of meas. Km = O(s log N)!

Joint Recovery of Initial State and Sparse Inputs



Key Observations

- Num. meas. sufficient for observability reduced from O(N) to O(s log N)
- Fewer measurements if
 - Transfer matrix D is well conditioned
 - Sparsity s is small
- Systems unobservable in the classical setting are observable under sparsity constraints!

Distributed Recovery: Learning Over a Network

- Network of L data centers
 Node j has observation y;
- Want to learn x;
 - Statistically related to y;
- Centralized processing:
 - Optimal, but
 - Computationally demanding
- Distributed (in-network) processing:
 - Secure
 - Robust to node failures



Recap: SBL for Joint Sparse Recovery

EM Iterations:

• E-step:
$$\begin{split} \Sigma_j^{k+1} &= \Gamma^k - \Gamma^k \Phi_j^T \left(\sigma_j^2 \mathbf{I}_M + \Phi_j \Gamma^k \Phi_j^T \right)^{-1} \Phi_j \Gamma^k \\ \mu_j^{k+1} &= \sigma_j^{-2} \Sigma_j^{k+1} \Phi_j^T \mathbf{y}_j \end{split}$$

• Separable: x_j are independent given Γ • Can be computed locally at each node • M-step: not separable $\Gamma^{k+1} = \frac{1}{L} \sum_{i=1}^{L} \mathbf{a}_j^{(k+1)}$

A Simple Trick

Equivalent problems

$$\gamma^* = \frac{1}{L} \sum_{j=1}^{L} a_j \qquad \gamma^* = \arg\min_{\gamma} \sum_{j=1}^{L} |\gamma - a_j|^2$$

For distributed implementation

- Can be computed locally at each node! Objective fn. separable
- Can now use, e.g., ADMM to solve

$$\underset{j=1}{\operatorname{arg min}} \gamma_j, j \in [L] \sum_{j=1}^L |\gamma_j - a_j|^2$$

subject to $\gamma_j = \gamma_b, b \in \mathcal{B}_j, j \in [L]$

Bridge nodes Linear constraints

Simulation Result: NMSE Phase Transition



L = 5 nodes, n = 50, m = 10, 10% sparsity, SNR = 30 dB

[S. Khanna, C. R. Murthy, TSIPN 2017]

Part 8: Applications



Wireless channel estimation & data detection



- Wireless channels exhibit multipath
 Naturally sparse in the lag-domain
- Channel equalization & data detection
 Need to estimate both support & channel

Channel Models

Block fading channel:

Channel constant for the duration of a block (say, K symbols), changes i.i.d. from block-toblock (classic SMV-SBL)

Time-varying channel:

Channel varies from symbol-to-symbol

Want to exploit temporal correlation and groupsparsity (MMV-SBL)

Outline

- 1. Block fading case:
 - 1. Known channel support: Joint channel estimation & data detection
 - 2. Unknown channel support: Channel and support estimation using pilot symbols
 - 3. Unknown data & support: Joint support, channel estimation & data detection
- 2. Time-varying case:
 - 1. AR model: Kalman-EM algo for joint support, channel estimation & data detn



Sparse Channel Estimation from Pilot Symbols



- h sparse in time (lag) domain
- Hierarchical prior: $h(i) = \mathcal{CN}(0, \gamma_i)$ Y_i deterministic, unknown hyperparams
- Goal:
 Given y, X, estimate h & sparsity profile

Joint Channel, Support Estmn. & Data Detn.

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Get h as a by-product of the E-step

Simulation Result

 \mathfrak{M}

- OFDM system
- N=256 subcarriers,
- max delay spread
 L=64
- K=7 symbols/slot
- PedB PDP:
 6 nonzero taps
- 44 pilot subcarriers
- Data: rate ½ turbo code, QPSK



BER Performance



Time-Varying Channels

- Channel correlated from symbol-tosymbol
- AR model: $\mathbf{h}_k = \rho \mathbf{h}_{k-1} + \mathbf{u}_k$
- The factor p depends on the normalized doppler freq, which in turn depends on the speed of the mobile
- SBL framework can be extended to incorporate the temporal correlation

Joint Kalman SBL (JK-SBL)

 Complexity O(KL³): smaller than block-based methods O(K³L³) [Zhang et al. 10]
 (K = num. OFDM symbols used in joint estimation)

In the block-fading case: get recursive, more computationally efficient versions of our algos

E-Step $j = 1, 2, \dots, K$ **Predict:** $\hat{\mathbf{h}}_{j|j-1}, \mathbf{P}_{j|j-1}$ **Update:** $\hat{\mathbf{h}}_{j|j}, \mathbf{P}_{j|j}$ Smooth: $\hat{\mathbf{h}}_{j-1|K}, \mathbf{P}_{j-1|K}$ **M**-step $oldsymbol{\gamma}^{(r)}, \mathbf{X}_1^{(r)}, \dots, \mathbf{X}_K^{(r)}$

 $\mathcal{O}(KL^3)$

Simulation Result \mathcal{H} 10 10¹ Solid: Uncoded Dashed: Coded 10^{0} FDI 10^{-2} SBL J-SBL K-SBI 10 **FM-OFDM** 10^{-3} JK-SBL MIP-aware Kalmar BER MSE 10 10^{-4} 10⁻³ SBL О J-SBL 10^{-5} Δ K-SBL 10-4 EM-OFDM JK-SBL Genie 10⁻⁵ 10^{-6} 10 15 20 25 30 10 15 30 5 20 25 SNR E_b/N_0

• $f_d T_s = 0.001$ (slowly time-varying)

MIMO-OFDM $\mathcal{H}\mathcal{H}$ OFDM MODULATOR TURBO ENCODER-INPUT BITS SYMBOL MIMO MAPPING ENCODER OFDM MODULATO OFDM DEMODULATO TURBO DECODER-OUTPUT BITS LLR ➡ {b̂} AND DATA DEINTERLEAVE OFDM DEMODULATOR YN, $\hat{\mathbf{h}}_{11},\ldots,\hat{\mathbf{h}}_{N_tN_t}$ $\mathbf{y}_{n_r} = \sum_{n_t=1}^{N_t} \mathbf{X}_{n_t} \mathbf{F} \mathbf{h}_{n_t n_r} + \mathbf{v}_{n_r}, \ n_r = 1, \dots, N_r$ Goal: Recover h1, ..., hNr from y1 ... yNr

[Prasad & M., NCC 2014]

MMV Framework

 \widetilde{m}

Measurement model



Pilot subcarriers



MSE Performance

 \widetilde{m}



- 256 subcarriers
- CP Length 64
 - 44 pilot
 subcarriers
- PedB PDP
- QPSK constellation





BER Performance

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Wideband spectrum sensing using compressive measurements



Experimental setup No. primary users = 5 No. secondary users = 10 11 of total 128 frequency subbands are in use SNR range: -2 4 to 28 dB

- SNR range: -2.4 to 7.8 dB.
- Compression ratio = 12.5



Summary

- Bayesian methods:
 - Simple updates
 - Promising performance
- Challenges:
 - Theoretical analysis
 - New algorithms
 - Novel applications
- Plenty of opportunities!

References

J. M. Adler, B. Rao, and K. Kreutz-Delgado, Comparison of basis selection methods, Asilomar 1999

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- S. F. Cotter, B. Rao, K. Engan, and K. Kreutz-Delgado, Sparse solutions to linear inverse problems with multiple measurement vectors, IEEE Trans. Sig. Proc., 2005
- D. Wipf and B. Rao, Sparse Bayesian learning for basis selection, IEEE Trans. on Sig. Proc., 2004
- D. Wipf and B. Rao, An empirical bayesian strategy for solving the simultaneous sparse approximation problem, IEEE Trans. Sig. Proc., 2007
- Z. Zhang and B. Rao, Sparse signal recovery with temporally correlated source vectors using sparse bayesian learning, IEEE J-STSP, 2011
- Z. Zhang and B. Rao, Recovery of block sparse signals using the framework of block sparse bayesian learning, ICASSP 2012
- R. Giri, B. Rao, Type I and Type II Bayesian Methods for Sparse Signal Recovery using Scale Mixtures, submitted, IEEE Trans. Sig. Proc., 2015

References

R. Prasad and C. R. Murthy, Cramér-Rao-Type Bounds for Sparse Bayesian Learning, IEEE Transactions on Sig. Proc., vol. 61, no. 3, pp. 622–632, Mar. 2013

 \widetilde{m}

- R. Prasad, C. R. Murthy and B. Rao, Joint Approximately Sparse Channel Estimation and Data Detection in OFDM Systems using Sparse Bayesian Learning, IEEE Trans. Sig. Proc., Jul. 2014
- R. Prasad and C. R. Murthy, Joint Approximately Group Sparse Channel Estimation and Data Detection in MIMO-OFDM Systems Using Sparse Bayesian Learning, NCC 2014 [best paper award!]
- S. Khanna and C. R. Murthy, Decentralized Bayesian Learning of Jointly Sparse Signals, Globecom 2014
- V. Vinuthna, R. Prasad, and C. R. Murthy, Sparse signal recovery in the presence of colored noise and rank-deficient noise covariance matrix: an SBL approach, ICASSP 2015
- R. Prasad, C. R. Murthy, and B. D. Rao, Joint Channel Estimation and Data Detection in MIMO-OFDM Systems: A Sparse Bayesian Learning Approach, IEEE Trans. on Sig. Proc., Oct. 2015
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Thank you!



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