Lecture 19 : Martingales

1 Martingales

Definition 1.1. A stochastic process $\{Z_n, n \in \mathbb{N}\}$ is said to be a **martingale** if

- 1. $\mathbb{E}[|Z_n|] < \infty$, for all n.
- 2. $\mathbb{E}[Z_{n+1}|Z_1, Z_2, \dots, Z_n] = Z_n.$

If the equality in second condition is replaced by \leq or \geq , then the process is called **super-martingale** or **submartingale**, respectively.

Remark 1.2. Taking expectation on both sides of part 2 of the above definition, we get $\mathbb{E}[Z_{n+1}] = \mathbb{E}[Z_n]$, and hence $\mathbb{E}[Z_{n+1}] = \mathbb{E}[Z_1]$, for all n.

Example 1.3 (Simple random walk). Let $\{X_i\}$ be a sequence of independent random variables with mean 0. Let $Z_n = \sum_{i=1}^n X_i$. Then, $\{Z_n, n \in \mathbb{N}\}$ is a martingale. This is so because, $\mathbb{E}[Z_n] = 0$ and

$$\mathbb{E}[Z_{n+1}|Z_1, Z_2 \dots Z_n] = \mathbb{E}[Z_n + X_{n+1}|Z_1, Z_2 \dots Z_n] = Z_n.$$

Example 1.4. Let $\{X_i\}$ be a sequence of independent random variables with mean 1. Let $Z_n = \prod_{i=1}^n X_i$. Then, $\{Z_n, n \in \mathbb{N}\}$ is a martingale. This is so because, $\mathbb{E}[Z_n] = 1$ and

$$\mathbb{E}[Z_{n+1}|Z_1, Z_2 \dots Z_n] = \mathbb{E}[Z_n X_{n+1}|Z_1, Z_2 \dots Z_n] = Z_n.$$

Example 1.5 (Branching Process). Let $\{X_n\}$ be a branching process. Let $X_0 = 1$. Then,

$$X_n = \sum_{i=1}^{X_{n-1}} Z_i,$$

where Z_i represents the number of offspring of the *i*th individual of the $(n-1)^{\text{st}}$ generation. conditioning on X_{n-1} yields, $\mathbb{E}[X_n] = \mu^n$ where μ is the mean number of offspring per individual. Then $\{Y_n = X_n/\mu^n : n \in \mathbb{N}\}$ is a martingale because $\mathbb{E}[Y_n] = 1$ and

$$\mathbb{E}[Y_{n+1}|Y_1, \dots Y_n] = \frac{1}{\mu^{n+1}} \mathbb{E}[\sum_{i=1}^{X_n} Z_i|Y_1, \dots Y_n] = \frac{X_n}{\mu^n} = Y_n.$$

Example 1.6 (Doob's Martingale). Let $X, Y_1, Y_2...$ be arbitrary random variables such that $\mathbb{E}[|X|] < \infty$. Then

$$Z_n = \mathbb{E}[X|Y_1, Y_2, \dots Y_n]$$

is a martingale. The integrability condition can be directly verified, and

$$\mathbb{E}[Z_{n+1}|Y_1, Y_2, \dots, Y_n] = \mathbb{E}[\mathbb{E}[X|Y_1, \dots, Y_{n+1}]|Y_1, \dots, Y_n] = \mathbb{E}[X|Y_1, \dots, Y_n]] = Z_n.$$

Example 1.7. For any sequence of random variables $X_1, X_2...$, the random variables $X_i - \mathbb{E}[X_i|X_1...X_{i-1}]$ have zero mean. Define

$$Z_n = \sum_{i=1}^n X_i - \mathbb{E}[X_i | X_1, X_2, \dots X_{i-1}]$$

is a martingale provided $\mathbb{E}[|Z_n|] < \infty$. To verify the same,

$$\mathbb{E}[Z_{n+1}|Z_1\dots Z_n] = \mathbb{E}[Z_n + X_n - \mathbb{E}[X_n|X_1\dots X_{n-1}]]$$
$$= Z_n + \mathbb{E}[X_n - \mathbb{E}[X_n|X_1\dots X_{n-1}]] = Z_n.$$

1.1 Stopping Times

Definition 1.8. The positive integer values, possibly infinite, random variable N is said to be a **random time** for the process $\{Z_n\}$ if the event $\{N = n\}$ is determined by the random variables $Z_1 \ldots Z_n$. If $\Pr\{N < \infty\} = 1$, then the random time N is said to be a **stopping time**.

Definition 1.9. A predictable sequence $\{H_n : n \in \mathbb{N}\}$ for process $\{X_n\}$ is the one where H_n is completely determined by $X_1, X_2, ..., X_{n-1}$

$$\sum_{m=1}^{n} H_m (X_m - Xm - 1) = (H \cdot X)_n$$

Theorem 1.10. Let $\{X_n, n \ge 0\}$ be a super martingale if $\{H_n \ge 0 : n \in \mathbb{N}\}$ is predictable and each H_n is bounded then $(H \cdot X)_n$ is super martingale

Proof.

$$\mathbb{E}[(H \cdot X)_{n+1} | (H \cdot X)_1 ... (H \cdot X)_n] = \mathbb{E}[H_{n+1}(X_{n+1} - X_n) + (H \cdot X)_n | (H \cdot X)_1 ... (H \cdot X)_n]$$

= $H_{n+1}(\mathbb{E}[X_{n+1} | (H \cdot X)_1 ... (H \cdot X)_n] - X_n) + (H \cdot X)_n$
 $\leq (H \cdot X)_n$

Definition 1.11. Let T be a random time for the process $\{X_n : n \in \mathbb{N}\}$, then stopping process $\{X_{T \wedge n}\}$ is defined as

$$X_{T \wedge n} = X_n \mathbb{1}_{\{n \le T\}} + X_T \mathbb{1}_{\{n > T\}}.$$

Proposition 1.12. If T is a random time for the martingale $\{X_n : n \in \mathbb{N}\}$, then the stopping process $\{X_{T \wedge n}\}$ is a martingale.

Proof.

$$X_{T \wedge n} = (H \cdot X)_n \text{ when } H_n = 1_{\{n \le T\}}$$
$$X_{T \wedge n} = X_{T \wedge n-1} + 1_{\{n \le T\}} (X_n - X_{n-1})$$

 $\begin{array}{ll} n \leq T \colon & X_{T \wedge n} = X_n \\ n > T \colon \text{Since } n > T \text{ gives } n-1 \geq T \text{ therefore } T \wedge n-1 \geq T \text{ which implies } X_{T \wedge n} = X_T \end{array}$

$$X_{T \wedge n} = X_0 + \sum_{m=1}^n \mathbb{1}_{\{m \le T\}} (X_m - X_{m-1})$$

It is suffice to show $\{1_{n < T}\}$ is a predictable sequence which is true since

$${n \le T} = {T > n-1} = {T < n-1}^c$$

Therefore from the previous theorem we have

$$\mathbb{E}[X_{T \wedge n}] = \mathbb{E}[X_{T \wedge 1}] = \mathbb{E}[X_1]$$

Remark 1.13. For any martingale $\{X_n : n \in \mathbb{N}\}$, we have $\mathbb{E}[X_{T \wedge n}] = \mathbb{E}[X_1]$, for all n. Now assume that T is a stopping time. It is immediate that

$$\Pr\left\{\lim_{n\in\mathbb{N}}X_{T\wedge n}=X_N\right\}=1.$$

But is it true that

$$\lim_{n \in \mathbb{N}} \mathbb{E}[X_{T \wedge n}] = \mathbb{E}[X_N]?$$

It so turns out that the above is true under some additional regularity constraints only.

Theorem 1.14 (Martingale Stopping Theorem). If T is a stopping time for a martingale $\{X_n : n \in \mathbb{N}\}$ such that either of the following conditions is true:

- (i) T is bounded,
- (ii) $X_{T \wedge n}$ is uniformly bounded,
- (iii) $\mathbb{E}[T] < \infty$, and for some real positive K, we have $\sup_{n \in \mathbb{N}} \mathbb{E}[|X_{n+1} X_n||X_1 \dots X_n] < K$,
- then X_T is integrable and $\lim_{n \in \mathbb{N}} \mathbb{E}[X_{T \wedge n}] = \mathbb{E}[X_T] = \mathbb{E}[X_1].$

Proof. We show this is true for all three cases.

(i) Let K be the bound on T then for all $n \ge K$, we have $X_{T \wedge n} = X_T$, and hence it follows that

$$\mathbb{E}X_1 = \mathbb{E}X_{T \wedge n} = \mathbb{E}X_T, \ \forall n \ge K.$$

- (ii) Dominated convergence theorem implies the result.
- (iii) Since T is integrable and

$$X_{T \wedge n} \le |X_1| + KT_2$$

we observe that $X_{T \wedge n}$ is bounded by an integrable random variable, and hence result follows from dominated convergence theorem.

Corollary 1.15 (Wald's Equation). If T is a stopping time for $\{X_i, i \in \mathbb{N}\}$ iid with $\mathbb{E}[|X|] < \infty$ and $\mathbb{E}[T] < \infty$, then

$$\mathbb{E}[\sum_{i=1}^{T} X_i] = \mathbb{E}[T]\mathbb{E}[X].$$

Proof. Let $\mu = \mathbb{E}[X]$. Then $\{Z_n = \sum_{i=1}^n (X_i - \mu) : n \in \mathbb{N}\}$ is a martingale and hence from the Martingale stopping theorem, we have $\mathbb{E}[Z_T] = \mathbb{E}[Z_1] = 0$. But

$$\mathbb{E}[Z_T] = \mathbb{E}\sum_{i=1}^N X_i - \mu \mathbb{E}N.$$

Observe that condition 3 for Martingale stopping theorem to hold can be directly verified. Hence the result follows. $\hfill \Box$

Before we state and prove martingale convergence theorem, we state some results which will be used in the proof of the theorem.

Lemma 1.16. If $\{Z_i, i \in \mathbb{N}\}$ is a submartingale and T is a stopping time such that $\Pr\{T \leq n\} = 1$ then

$$\mathbb{E}Z_1 \leq \mathbb{E}Z_T \leq \mathbb{E}Z_n.$$

Proof. It follows from Theorem 1.14 that since T is bounded, $\mathbb{E}[Z_T] \ge \mathbb{E}[Z_1]$. Now, since T is a stopping time, we see that for $\{T = k\}$

$$\mathbb{E}[Z_n|Z_1,\ldots,Z_T,T=k] = \mathbb{E}[Z_n|Z_1\ldots Z_k,T=k] = \mathbb{E}[Z_n|Z_1\ldots Z_k] \ge Z_k = Z_T.$$

Result follows by taking expectation on both sides.

Lemma 1.17. If $\{Z_n, n \in \mathbb{N}\}$ is a martingale and f is a convex function, then $\{f(Z_n), n \in \mathbb{N}\}$ is a submartigale.

Proof. The result is a direct consequence of Jensen's inequality.

$$\mathbb{E}[f(Z_n)|Z_1,\ldots,Z_n] \ge f(\mathbb{E}[Z_{n+1}|Z_1,\ldots,Z_n]) = f(Z_n).$$

Construction 1.18. Let $X_n, n \ge 0$ be a sub martingale. Let a < b and $N_0 = -1$.

$$N_{2k-1} = \inf\{m > N_{2k-2} : X_m \le a\}.$$

$$N_{2k} = \inf\{m > N_{2k-1} : X_m \le b\}.$$

The above quantities N_{2k-1}, N_{2k} are stopping times and the set containing values of m in the transition from a to b can be defined as

$$\{N_{2k-1} < m \le N_{2k}\} = \{N_{2k-1} < m-1\} \cap \{m > N_{2k}\}^c$$

= $\{m-1 \ge N_{2k-1}\} \cap \{m-1 \ge N_{2k}\}^c$

Since the above set depends on $\{m-1\}$ values instead of $\{m\}$ values, So

$$H_m = \mathbb{1}_{\{N_{2k-1} < m \le N_{2k}\}}$$
$$U_n = \sup\{k : N_{2k} \le n\}$$

 H_m defines a predictable sequence and U_n is the number of up crossings completed in time n. Lemma 1.19 (Upcrossing inequality). If $\{X_m : m \ge 0\}$ is a sub martingale, then

$$(b-a)\mathbb{E}U_n \leq \mathbb{E}Y_n - \mathbb{E}Y_0$$

where $Y_n := a + (X_n - a)^+$

Proof. Since X_n is a submartingale so is Y_n , as it is a convex function of X_n . Since each up crossing has a gain slightly more than b - a the following inequality exists

$$(b-a)U_n \leq (H \cdot Y)_n$$

= $\sum_{m=1}^n 1_{\{N_{2k-1} < m \le N_{2k}\}} (Y_{m+1} - Y_m)$
= $\sum_{k=1}^{U_n} (Y_{N_{2k+1}} - Y_{N_{2k+1}})$

Now let $K_m = 1 - H_m$ then clearly,

$$Y_n - Y_0 = (H \cdot Y)_n + (K \cdot Y)_n$$

we have by the submartingale property of Y

$$\mathbb{E}[(K \cdot Y)_n] \ge \mathbb{E}[(K \cdot Y)_0] = 0$$

Therefore,

$$\mathbb{E}[(H \cdot Y)_n] + \mathbb{E}[(K \cdot Y)_n] = \mathbb{E}[Y_n - Y_0]$$
$$\mathbb{E}[(H \cdot Y)_n] \leq \mathbb{E}[Y_n - Y_0]$$
$$(b - a)\mathbb{E}U_n \leq \mathbb{E}(Y_n - Y_0)$$

Theorem 1.20 (Martingale Convergence Theorem). If X_n is a submartingale with $\sup_{a \in A} \mathbb{E}[X_n^+] \leq \infty$ then $\lim_{n \in \mathbb{N}} X_n = X$ a.s with $\mathbb{E}[X] < \infty$.

Proof.

$$(X-a)^+ \leq X^+ + |a|$$
$$\mathbb{E}[U_n] \leq \frac{\mathbb{E}[X_n^+] + |a|}{b-a}$$

 $\lim_{n\to\infty} U_n = U$ since, $\mathbb{E}[X_n^+] < \infty$ gives $U < \infty$ a.s. This conclusion leads to

$$Pr\{\bigcup_{a,b\in\mathbb{Q}} \{\liminf_{n\in\mathbb{N}} X_n < a < b < \limsup_{n\in\mathbb{N}} X_n\}\} = 0;$$

From the above probability we have a.s

$$\underset{n\in\mathbb{N}}{limsup}X_n = \underset{n\in\mathbb{N}}{limin}fX_n$$

Now the Fatou's lemma in measure theory guarantees

$$\mathbb{E}[X^+] \le \liminf_{n \in \mathbb{N}} f\mathbb{E}[X_n^+] < \infty$$

which implies $X < \infty$ almost sure. To see $X > -\infty$, we observe that

$$\mathbb{E}[X_n^-] = \mathbb{E}[X_n^+] - \mathbb{E}[X_n]$$
$$\leq \mathbb{E}[X_n^+] - \mathbb{E}[X_0]$$

The above inequality comes from the submartingale property of X_n . Now from another application of Fatou's lemma gives,

$$\mathbb{E}[X^{-}] \leq \liminf_{n \in \mathbb{N}} f\mathbb{E}[X_{n}^{-}] \leq \sup_{n \in \mathbb{N}} \mathbb{E}[X_{n}^{+}] - \mathbb{E}[X_{0}] < \infty.$$