

# Lecture 21: Exchangeability

## 1 Random Walk

**Definition 1.1.** Let  $\{X_i : i \in \mathbb{N}\}$  be *iid* random variables with finite  $\mathbb{E}[|X_1|]$ . Let

$$S_n = \sum_{k=1}^n X_k, \quad n \in \mathbb{N}_0.$$

Then the process  $\{S_n : n \in \mathbb{N}_0\}$  is called a **random walk**.

**Definition 1.2.** A random walk is called a **simple random walk** if

$$\Pr\{X_1 = 1\} = 1 - \Pr\{X_1 = -1\}.$$

*Remark 1.3.* A simple random walk has the interpretation of the winnings of a gambler who plays a simple coin toss game and wins Rupee 1 if heads and loses Rupee 1 if tails.

*Remark 1.4.* Random walks are useful in analyzing GI/GI/1 Queues, Ruin systems and even stock prices.

**Definition 1.5.** Let  $X_i$  belong to probability space  $(S, \mathcal{S}, \mu)$ . Consider the probability space  $(\Omega, \mathcal{F}, P)$  for process  $\{X_i : i \in \mathbb{N}\}$  where

$$\Omega = \prod_{i \in \mathbb{N}} S, \quad \mathcal{F} = \prod_{i \in \mathbb{N}} \mathcal{S}, \quad P = \prod_{i \in \mathbb{N}} \mu.$$

A **finite permutation** of  $\mathbb{N}$  is a map  $\pi : \mathbb{N} \rightarrow \mathbb{N}$  such that  $\pi(i) \neq i$  for only finitely many  $i$ .

**Definition 1.6.** For a finite permutation  $\pi$ , we define  $(\pi\omega)_i = \omega_{\pi(i)}$  for all  $i \in \mathbb{N}$ .

**Definition 1.7.** An event  $A$  is **permutable** if  $A = \pi^{-1}A = \{\omega \in \Omega : \pi\omega \in A\}$  for any finite permutation  $\pi$ .

**Definition 1.8.** The collection of permutable events is a  $\sigma$ -field called the **exchangeable**  $\sigma$ -field and denoted by  $\mathcal{E}$ .

**Definition 1.9.** A sequence  $X$  of random variables is called **exchangeable** if for each  $n$  and permutation  $\pi : [n] \rightarrow [n]$ , joint distribution of  $(X_1, X_2, \dots, X_n)$  and  $(X_{\pi(1)}, X_{\pi(2)}, \dots, X_{\pi(n)})$  are same.

**Example 1.10.** Suppose balls are selected randomly, without replacement, from an urn consisting of  $n$  balls of which  $k$  are white. For  $i \in [n]$ , let

$$X_i = 1_{\{i^{\text{th}} \text{ selection is white}\}},$$

then  $(X_1, \dots, X_n)$  will be exchangeable but not independent. In particular, let  $A = \{i \in [n] : X_i = 1\}$ . Then, we know that  $|A| = k$ , and we can write

$$\Pr\{X_i = 1, i \in A, X_j = 0, j \in A^c\} = \Pr\{A = (i_1, i_2, \dots, i_k)\} = \frac{(n-k)!k!}{n!} = \frac{1}{\binom{n}{k}}.$$

This joint distribution is independent of set of exact locations  $A$ , and hence exchangeable. Further, we can show that all  $X_i$  are identically distributed, since

$$\Pr\{X_1 = 1, X_2, \dots, X_n\} = \Pr\{X_i = 1, X_1, \dots, X_{i-1}, X_i, \dots, X_n\}.$$

Further, it can be seen that

$$\Pr\{X_2 = 1 | X_1 = 1\} = \frac{k-1}{n-1} \neq \frac{k}{n-1} = \Pr\{X_2 = 1 | X_1 = 0\}.$$

**Example 1.11.** Let  $\Lambda$  denote a random variable having distribution  $G$ . Let  $X$  be a sequence of dependent random variables, where each of these random variables are conditionally *iid* with distribution  $F_\lambda$  given  $\Lambda = \lambda$ . Then, these random variables are exchangeable since

$$\Pr\{X_1 \leq x_1, \dots, X_n \leq x_n\} = \int_\lambda \prod_{i=1}^n F_\lambda(x_i) dG(\lambda),$$

which is symmetric in  $(x_1, \dots, x_n)$ .

**Theorem 1.12 (de Finetti's Theorem).** *If  $X$  is an exchangeable sequence of random variables then conditioned on  $\mathcal{E}$ , each random variable  $X_i$  in the sequence is iid.*

*Proof.* Let  $I_{n,k} = \{i \subseteq [n]^k : i_j \text{ distinct}\}$ . Then, for a function  $\phi : \mathbb{R}^k \rightarrow \mathbb{R}$ , we can define

$$A_n(\phi) = \frac{1}{\binom{n}{k}} \sum_{i \in I_{n,k}} \phi(X_{i_1}, X_{i_2}, \dots, X_{i_k}),$$

where  $\binom{n}{k} = n(n-1)\dots(n-k+1)$ . Clearly,  $A_n(\phi) \in \mathcal{E}_n$  measurable and hence,  $\mathbb{E}[A_n(\phi) | \mathcal{E}_n] = A_n(\phi)$ . Since,  $X$  is exchangeable, we have

$$A_n(\phi) = \frac{1}{\binom{n}{k}} \sum_{i \in I_{n,k}} \mathbb{E}[\phi(X_{i_1}, X_{i_2}, \dots, X_{i_k}) | \mathcal{E}_n] = \mathbb{E}[\phi(X_1, X_2, \dots, X_k) | \mathcal{E}_n].$$

Since  $\mathcal{E}_n \rightarrow \mathcal{E}$ , we have

$$\lim_{n \in \mathbb{N}} A_n(\phi) = \lim_{n \in \mathbb{N}} \mathbb{E}[\phi(X_1, X_2, \dots, X_k) | \mathcal{E}_n] = \mathbb{E}[\phi(X_1, X_2, \dots, X_k) | \mathcal{E}].$$

Let  $f$  and  $g$  be bounded functions on  $\mathbb{R}^{k-1}$  and  $\mathbb{R}$  respectively, such that  $\phi(x_1, \dots, x_k) = f(x_1, \dots, x_{k-1})g(x_k)$ . We also define  $\phi_j(x_1, \dots, x_{k-1}) = f(x_1, \dots, x_{k-1})g(x_j)$ , to write

$$\begin{aligned} \binom{n}{k-1} A_n(f) n A_n(g) &= \sum_{i \in I_{n,k-1}} f(X_{i_1}, \dots, X_{i_{k-1}}) \sum_m g(X_m) \\ &= \sum_{i \in I_{n,k}} f(X_{i_1}, \dots, X_{i_{k-1}}) g(X_{i_k}) + \sum_{i \in I_{n,k-1}} \sum_{j=1}^{k-1} f(X_{i_1}, \dots, X_{i_{k-1}}) g(X_{i_j}) \\ &= \binom{n}{k} A_n(\phi) + \sum_{j=1}^k \binom{n}{k-1} A_n(\phi_j). \end{aligned}$$

Dividing both sides by  $(n)_k$  and rearranging terms, we get

$$A_n(\phi) = \frac{n}{n-k+1} A_n(f) A_n(g) - \frac{1}{n-k+1} \sum_{j=1}^k A_n(\phi_j),$$

Using above two results, theorem follows by induction. □