

# Lecture-03: Independence

## 1 Independence

**Definition 1.1 (Independence of events).** For a probability space  $(\Omega, \mathcal{F}, P)$ , a family of events  $(A_i \in \mathcal{F} : i \in I)$  is said to be independent, if for any finite set  $F \subseteq I$ , we have

$$P(\cap_{i \in F} A_i) = \prod_{i \in F} P(A_i).$$

*Remark 1.* The certain event  $\Omega$  and the impossible event  $\emptyset$  are always independent to every event  $A \in \mathcal{F}$ .

**Example 1.2 (Two coin tosses).** Consider two coin tosses, such that the sample space is  $\Omega = \{HH, HT, TH, TT\}$ , and the event space is  $\mathcal{F} = 2^\Omega$ . It suffices to define a probability function  $P : \mathcal{F} \rightarrow [0, 1]$  on the sample space. We define one such probability function  $P$ , such that

$$P(\{HH\}) = P(\{HT\}) = P(\{TH\}) = P(\{TT\}) = \frac{1}{4}.$$

Let event  $A_1 \triangleq \{HH, HT\}$  and  $B_2 \triangleq \{HH, TH\}$  correspond to getting a head on the first or the second toss respectively.

From the defined probability function, we obtain the probability of getting a tail on the first or the second toss is  $\frac{1}{2}$ , and identical to the probability of getting a head on the first or the second toss. That is,  $P(A_1) = P(A_2) = \frac{1}{2}$  and the intersecting event  $A_1 \cap A_2 = \{HH\}$  with the probability  $P(A_1 \cap A_2) = \frac{1}{4}$ . That is, for events  $A_1, A_2 \in \mathcal{F}$ , we have

$$P(A_1 \cap A_2) = P(A_1)P(A_2).$$

That is, events  $A_1$  and  $A_2$  are independent.

**Example 1.3 (Countably infinite coin tosses).** Consider a sequence of coin tosses, such that the sample space is  $\Omega = \{H, T\}^{\mathbb{N}}$ . For set of outcomes  $E_n \triangleq \{\omega \in \Omega : \omega_n = H\}$ , we consider an event space generated by  $\mathcal{F} \triangleq \sigma(\{E_n : n \in \mathbb{N}\})$ . We define a probability function  $P : \mathcal{F} \rightarrow [0, 1]$  by  $P(\cap_{i \in F} E_i) = p^{|F|}$  for any finite subset  $F \subseteq \mathbb{N}$ . By definition,  $(E_n : n \in \mathbb{N})$  is a sequence of independent events.

We observe that the set of outcomes corresponding to at least one head in first  $n$  outcomes

$$A_n \triangleq \{\omega \in \Omega : \omega_i = H \text{ for some } i \in [n]\} = \cup_{i=1}^n E_i \in \mathcal{F},$$

and set of outcomes corresponding to first head at the  $n$ th outcome

$$B_n \triangleq \{\omega \in \Omega : \omega_1 = \dots = \omega_{n-1} = T, \omega_n = H\} = \cap_{i=1}^{n-1} E_i^c \cap E_n \in \mathcal{F}.$$

In particular this implies that  $\sigma(\{A_n : n \in \mathbb{N}\}) \subseteq \mathcal{F}$  and  $\sigma(\{B_n : n \in \mathbb{N}\}) \subseteq \mathcal{F}$ . We can show that  $P(A_n) = 1 - (1 - p)^n$  and  $P(B_n) = p(1 - p)^{n-1}$  for  $n \in \mathbb{N}$ .

Let  $\mathcal{F}_n$  be the event space generated by the first  $n$  coin tosses, i.e.  $\mathcal{F}_n \triangleq \sigma(\{E_i : i \in [n]\})$ . Then, we can show that  $\mathcal{F} = \sigma(\{\mathcal{F}_n : n \in \mathbb{N}\})$ . For any  $\omega \in \Omega$ , we can define the number of heads in first  $n$  trials by  $k_n \triangleq \sum_{i=1}^n \mathbb{1}_{\{\omega_i=H\}}$ . Then, we observe that any event  $A \in \mathcal{F}_n$  can be written as union of  $\cap_{i=1}^n C_i$  where  $C_i = E_i$  or  $E_i^c$ . That is, we can specify the first  $n$  outcomes for each  $\omega \in A$ . Since the probability  $P(\cap_{i=1}^n C_i) = \prod_{i=1}^n P(C_i)$ , we have

$$P(A) = \sum_{\omega \in A} p^{k_n(\omega)} (1-p)^{n-k_n(\omega)}.$$

**Example 1.4 (Counter example).** Consider a probability space  $(\Omega, \mathcal{F}, P)$  and the events  $A_1, A_2, A_3 \in \mathcal{F}$ . The condition  $P(A_1 \cap A_2 \cap A_3) = P(A_1)P(A_2)P(A_3)$  is not sufficient to guarantee independence of the three events. In particular, we see that if

$$P(A_1 \cap A_2 \cap A_3) = P(A_1)P(A_2)P(A_3), \quad P(A_1 \cap A_2 \cap A_3^c) \neq P(A_1)P(A_2)P(A_3^c),$$

then  $P(A_1 \cap A_2) = P(A_1 \cap A_2 \cap A_3) + P(A_1 \cap A_2 \cap A_3^c) \neq P(A_1)P(A_2)$ .

**Definition 1.5.** A family of collections of events  $(\mathcal{A}_i \subseteq \mathcal{F} : i \in I)$  is called independent, if for any finite set  $F \subseteq I$  and  $A_i \in \mathcal{A}_i$  for all  $i \in F$ , we have

$$P(\cap_{i \in F} A_i) = \prod_{i \in F} P(A_i).$$

## 2 Law of Total Probability

**Theorem 2.1 (Law of total probability).** For a probability space  $(\Omega, \mathcal{F}, P)$ , consider a sequence of events  $B = (B_n \in \mathcal{F} : n \in \mathbb{N})$  that partitions the sample space  $\Omega$ , i.e.  $B_m \cap B_n = \emptyset$  for all  $m \neq n$ , and  $\cup_{n \in \mathbb{N}} B_n = \Omega$ . Then, for any event  $A \in \mathcal{F}$ , we have

$$P(A) = \sum_{n \in \mathbb{N}} P(A \cap B_n).$$

*Proof.* We can expand any event  $A \in \mathcal{F}$  in terms of any partition  $B$  of the sample space  $\Omega$  as

$$A = A \cap \Omega = A \cap (\cup_{n \in \mathbb{N}} B_n) = \cup_{n \in \mathbb{N}} (A \cap B_n).$$

From the mutual disjointness of the events  $(B_n \in \mathcal{F} : n \in \mathbb{N})$ , it follows that the sequence  $(A \cap B_n \in \mathcal{F} : n \in \mathbb{N})$  is mutually disjoint. The result follows from the countable additivity of probability of disjoint events.  $\square$

## 3 Conditional Probability

Consider  $N$  trials of a random experiment over an outcome space  $\Omega$  and an event space  $\mathcal{F}$ . Let  $\omega_n \in \Omega$  denote the outcome of the experiment of the  $n$ th trial. Consider two events  $A, B \in \mathcal{F}$  and denote the number of times event  $A$  and event  $B$  occurs by  $N(A)$  and  $N(B)$  respectively. We denote the number of times both events  $A$  and  $B$  occurred by  $N(A \cap B)$ . Then, we can write these numbers in terms of indicator functions as

$$N(A) = \sum_{n=1}^N \mathbb{1}_{\{\omega_n \in A\}}, \quad N(B) = \sum_{n=1}^N \mathbb{1}_{\{\omega_n \in B\}}, \quad N(A \cap B) = \sum_{n=1}^N \mathbb{1}_{\{\omega_n \in A \cap B\}}.$$

We denote the relative frequency of events  $A, B, A \cap B$  in  $N$  trials by  $\frac{N(A)}{N}, \frac{N(B)}{N}, \frac{N(A \cap B)}{N}$  respectively. We can find the relative frequency of events  $A$ , on the trials where  $B$  occurred as

$$\frac{\frac{N(A \cap B)}{N}}{\frac{N(B)}{N}} = \frac{N(A \cap B)}{N(B)}.$$

Inspired by the relative frequency, we define the conditional probability function conditioned on events.

**Definition 3.1.** Fix an event  $B \in \mathcal{F}$  such that  $P(B) > 0$ , we can define the conditional probability  $P(\cdot|B) : \mathcal{F} \rightarrow [0,1]$  of any event  $A \in \mathcal{F}$  conditioned on the event  $B$  as

$$P(A|B) = \frac{P(A \cap B)}{P(B)}.$$

**Lemma 3.2 (Conditional probability).** For any event  $B \in \mathcal{F}$  such that  $P(B) > 0$ , the conditional probability  $P(\cdot|B) : \mathcal{F} \rightarrow [0,1]$  is a probability measure on space  $(\Omega, \mathcal{F})$ .

*Proof.* We will show that the conditional probability satisfies all four axioms of a probability measure.

**Non-negativity:** For all events  $A \in \mathcal{F}$ , we have  $P(A|B) \geq 0$  since  $P(A \cap B) \geq 0$ .

**$\sigma$ -additivity:** For an infinite sequence of mutually disjoint events  $(A_i \in \mathcal{F} : i \in \mathbb{N})$  such that  $A_i \cap A_j = \emptyset$  for all  $i \neq j$ , we have  $P(\cup_{i \in \mathbb{N}} A_i | B) = \sum_{i \in \mathbb{N}} P(A_i | B)$ . This follows from disjointness of the sequence  $(A_i \cap B \in \mathcal{F} : i \in \mathbb{N})$ .

**Certainty:** Since  $\Omega \cap B = B$ , we have  $P(\Omega|B) = 1$ .

□

*Remark 2.* For two independent events  $A, B \in \mathcal{F}$  such that  $P(A \cap B) > 0$ , we have  $P(A|B) = P(A)$  and  $P(B|A) = P(B)$ . If either  $P(A) = 0$  or  $P(B) = 0$ , then  $P(A \cap B) = 0$ .

*Remark 3.* For any partition  $B$  of the sample space  $\Omega$ , if  $P(B_n) > 0$  for all  $n \in \mathbb{N}$ , then from the law of total probability and the definition of conditional probability, we have

$$P(A) = \sum_{n \in \mathbb{N}} P(A|B_n)P(B_n).$$

## 4 Conditional Independence

**Definition 4.1 (Conditional independence of events).** For a probability space  $(\Omega, \mathcal{F}, P)$ , a family of events  $(A_i \in \mathcal{F} : i \in I)$  is said to be conditionally independent given an event  $C \in \mathcal{F}$  such that  $P(C) > 0$ , if for any finite set  $F \subseteq I$ , we have

$$P(\cap_{i \in F} A_i | C) = \prod_{i \in F} P(A_i | C).$$

*Remark 4.* Let  $C \in \mathcal{F}$  be an event such that  $P(C) > 0$ . Two events  $A, B \in \mathcal{F}$  are said to be conditionally independent given event  $C$ , if

$$P(A \cap B | C) = P(A | C)P(B | C).$$

If the event  $C = \Omega$ , it implies that  $A, B$  are independent events.

*Remark 5.* Two events may be independent, but not conditionally independent and vice versa.

**Example 4.2.** Consider two independent events  $A, B \in \mathcal{F}$  such that  $P(A \cap B) > 0$  and  $P(A \cup B) < 1$ . Then the events  $A$  and  $B$  are not conditionally independent given  $A \cup B$ . To see this, we observe that

$$P(A \cap B | A \cup B) = \frac{P((A \cap B) \cap (A \cup B))}{P(A \cup B)} = \frac{P(A \cap B)}{P(A \cup B)} = \frac{P(A)P(B)}{P(A \cup B)} = P(A|A \cup B)P(B).$$

We further observe that  $P(B|A \cup B) = \frac{P(B)}{P(A \cup B)} \neq P(B)$  and hence  $P(A \cap B | A \cup B) \neq P(A|A \cup B)P(B|A \cup B)$ .

**Example 4.3.** Consider two non-independent events  $A, B \in \mathcal{F}$  such that  $P(A) > 0$ . Then the events  $A$  and  $B$  are conditionally independent given  $A$ . To see this, we observe that

$$P(A \cap B | A) = \frac{P(A \cap B)}{P(A)} = P(B|A)P(A|A).$$