## Lecture-14: Weak convergence

## 1 Convergence in distribution

Definition 1.1 (convergence in distribution). A random sequence $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ defined on a probability space $(\Omega, \mathcal{F}, P)$ converges in distribution to a random variable $X_{\infty}: \Omega^{\prime} \rightarrow \mathbb{R}$ defined on a probability space $\left(\Omega^{\prime}, \mathcal{F}^{\prime}, P^{\prime}\right)$ if

$$
\lim _{n} F_{X_{n}}(x)=F_{X_{\infty}}(x)
$$

at all continuity points $x$ of $F_{X_{\infty}}$. Convergence in distribution is denoted by $\lim _{n} X_{n}=X_{\infty}$ in distribution.
Proposition 1.2. Consider a random sequence $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ defined on a probability space $(\Omega, \mathcal{F}, P)$ and a random variable $X_{\infty}: \Omega^{\prime} \rightarrow \mathbb{R}$ defined on another probability space $\left(\Omega^{\prime}, \mathcal{F}^{\prime}, P^{\prime}\right)$. Then the following statements are equivalent.
(a) $\lim _{n} X_{n}=X_{\infty}$ in distribution.
(b) $\lim _{n} \mathbb{E}\left[g\left(X_{n}\right)\right]=\mathbb{E}\left[g\left(X_{\infty}\right)\right]$ for any bounded continuous function $g$.
(c) Characteristic functions converge point-wise, i.e. $\lim _{n} \Phi_{X_{n}}(u)=\Phi_{X_{\infty}}(u)$ for each $u \in \mathbb{R}$.

Proof. Let $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ be a sequence of random variables and let $X_{\infty}: \Omega^{\prime} \rightarrow \mathbb{R}$ be a random variable. We will show that $(a) \Longrightarrow(b) \Longrightarrow(c) \Longrightarrow(a)$.
$(a) \Longrightarrow(b):$ Let $\lim _{n} X_{n}=X_{\infty}$ in distribution, then $\lim _{n} \int_{x \in \mathbb{R}} g(x) d F_{X_{n}}(x)=\int_{x \in \mathbb{R}} g(x) \lim _{n} d F_{X_{n}}(x)$ by the bounded convergence theorem for any bounded continuous function $g$.
$(b) \Longrightarrow(c):$ Let $\lim _{n} \mathbb{E}\left[g\left(X_{n}\right)\right]=\mathbb{E}\left[g\left(X_{\infty}\right)\right]$ for any bounded continuous function $g$. Taking $g(x)=e^{j u x}$, we get the result.
$(c) \Longrightarrow(a)$ : The proof of this part is technical and is omitted.

Example 1.3 (Convergence in distribution but not in probability). Consider a sequence of nondegenerate continuous i.i.d. random variables $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ and independent random variable $Y: \Omega^{\prime} \rightarrow$ $\mathbb{R}$ with the common distribution $F_{Y}$. Then $F_{X_{n}}=F_{Y}$ for all $n \in \mathbb{N}$, and hence $\lim _{n} X_{n}=Y$ in distribution. When each $X_{n}$ is zero mean Gaussian with variance $\sigma^{2}$, it follows that $X_{n}-Y$ is zero mean Gaussian with variance $2 \sigma^{2}$. Therefore, for $\epsilon<\sigma \sqrt{\pi}$ and all $n \in \mathbb{N}$

$$
P\left\{\left|X_{n}-Y\right| \leqslant \epsilon\right\}=\frac{1}{\sqrt{4 \pi \sigma^{2}}} \int_{z \in[-\epsilon, \epsilon]} e^{-z^{2} / 4 \sigma^{2}} d z \leqslant \frac{\epsilon}{\sigma \sqrt{\pi}}<1 .
$$

Lemma 1.4 (Convergence in probability implies in distribution). Consider a sequence $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ of random variables and a random variable $X_{\infty}: \Omega \rightarrow \mathbb{R}$ defined on a probability space $(\Omega, \mathcal{F}, P)$, such that $\lim _{n} X_{n}=X_{\infty}$ in probability, then $\lim _{n} X_{n}=X_{\infty}$ in distribution.

Proof. Fix $\epsilon>0$, and consider the event

$$
E_{n}(\epsilon) \triangleq\left\{\omega \in \Omega:\left|X_{n}(\omega)-X_{\infty}(\omega)\right|>\epsilon\right\}=\left\{X_{n} \notin\left[X_{\infty}-\epsilon, X_{\infty}+\epsilon\right]\right\} \in \mathcal{F} .
$$

We further define events $A_{X_{n}}(x) \triangleq\left\{X_{n} \leqslant x\right\}$ and $A_{X_{\infty}}(x) \triangleq\left\{X_{\infty} \leqslant x\right\}$, then we can write

$$
\begin{array}{ll}
A_{X_{n}}(x) \cap A_{X_{\infty}}(x+\epsilon) \subseteq A_{X_{\infty}}(x+\epsilon), & \\
A_{X_{n}}(x) \cap A_{X_{\infty}}^{c}(x+\epsilon) \subseteq A_{n}(\epsilon) \\
A_{X_{\infty}}(x-\epsilon) \cap A_{X_{n}}(x) \subseteq A_{X_{n}}(x), &
\end{array} A_{X_{\infty}}(x-\epsilon) \cap A_{X_{n}}^{c}(x) \subseteq A_{n}(\epsilon) .
$$

From the above set relations, law of total probability, and union bound, we have

$$
F_{X_{\infty}}(x-\epsilon)-P\left(E_{n}\right) \leqslant F_{X_{n}}(x) \leqslant F_{X_{\infty}}(x+\epsilon)+P\left(E_{n}\right)
$$

From the convergence in probability, we have $\lim _{n} P\left(E_{n}\right)=0$. Therefore, we get

$$
F_{X_{\infty}}(x-\epsilon) \leqslant \liminf _{n} F_{X_{n}}(x) \leqslant \lim \sup _{n} F_{X_{n}}(x) \leqslant F_{X_{\infty}}(x+\epsilon)
$$

We get the result at the continuity points of $F_{X}$, since the choice of $\epsilon$ was arbitrary.
Theorem 1.5 (Central Limit Theorem). Consider an i.i.d. random sequence $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ defined on a probability space $(\Omega, \mathcal{F}, P)$, with $\mathbb{E} X_{n}=\mu$ and $\operatorname{Var}\left(X_{n}\right)=\sigma^{2}$ for all $n \in \mathbb{N}$. We define the $n$-sum as $S_{n} \triangleq \sum_{i=1}^{n} X_{i}$ and consider a standard normal random variable $Y: \Omega \rightarrow \mathbb{R}$ with density function $f_{Y}(y)=\frac{1}{\sqrt{2 \pi}} e^{-\frac{y^{2}}{2}}$ for all $y \in \mathbb{R}$. Then,

$$
\lim _{n} \frac{S_{n}-n \mu}{\sigma \sqrt{n}}=Y \text { in distribution. }
$$

Proof. The classical proof is using the characteristic functions. Let $Z_{i} \triangleq \frac{X_{i}-\mu}{\sigma}$ for all $i \in \mathbb{N}$, then the shifted and scaled $n$-sum is given by $\frac{S_{n}-n \mu}{\sigma \sqrt{n}}=\frac{1}{\sqrt{n}} \sum_{i=1}^{n} Z_{i}$. We use the third equivalence in Proposition 1.2 to show that the characteristic function of converges to the characteristic function of the standard normal. We define the characteristic functions

$$
\Phi_{n}(u) \triangleq \mathbb{E} \exp \left(j u \frac{\left(S_{n}-n \mu\right)}{\sigma \sqrt{n}}\right), \quad \Phi_{Z_{i}}(u) \triangleq \mathbb{E} \exp \left(j u Z_{i}\right), \quad \Phi_{Y}(u) \triangleq \mathbb{E} \exp (j u Y)
$$

We can compute the characteristic function of the standard normal as

$$
\Phi_{Y}(u)=\frac{1}{\sqrt{2 \pi}} \int_{y \in \mathbb{R}} e^{-\frac{u^{2}}{2}} \exp \left(-\frac{(y-j u)^{2}}{2}\right) d y=e^{-\frac{u^{2}}{2}} .
$$

Since the random sequence $Z: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ is a zero mean i.i.d. sequence, using the Taylor expansion of the characteristic function, we have

$$
\Phi_{n}(u)=\prod_{i=1}^{n} \mathbb{E} \exp \left(j u \frac{\left(X_{i}-\mu\right)}{\sigma \sqrt{n}}\right)=\left[\Phi_{Z_{1}}\left(\frac{u}{\sqrt{n}}\right)\right]^{n}=\left[1-\frac{u^{2}}{2 n}+o\left(\frac{u^{2}}{n}\right)\right]^{n}
$$

For any $u \in \mathbb{R}$, taking limit $n \in \mathbb{N}$, we get the result.

## 2 Strong law of large numbers

Definition 2.1. For a random sequence $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ defined on a probability space $(\Omega, \mathcal{F}, P)$ with bounded mean $\mathbb{E}\left|X_{n}\right|<\infty$ for all $n \in \mathbb{N}$, we define the $n$-sum as $S_{n} \triangleq \sum_{i=1}^{n} X_{i}$ and the empirical $n$-mean $\frac{S_{n}}{n}$ for each $n \in \mathbb{N}$. For each $n \in \mathbb{N}$, we define event

$$
E_{n} \triangleq\left\{\omega \in \Omega:\left|S_{n}-\mathbb{E} S_{n}\right|>n \epsilon\right\} \in \mathcal{F}
$$

Theorem 2.2 ( $L^{4}$ strong law of large numbers). Let $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ be a sequence of independent random variables defined on probability space $(\Omega, \mathcal{F}, P)$ with bounded mean $\mathbb{E} X_{n}$ for each $n \in \mathbb{N}$ and uniformly bounded fourth central moment $\sup _{n \in \mathbb{N}} \mathbb{E}\left(X_{n}-\mathbb{E} X_{n}\right)^{4} \leqslant B<\infty$. Then, the empirical n-mean converges to $\lim _{n} \frac{\mathbb{E} S_{n}}{n}$ almost surely.

Proof. Recall that $\mathbb{E}\left(S_{n}-\mathbb{E} S_{n}\right)^{4}=\mathbb{E}\left(\sum_{i=1}^{n}\left(X_{i}-\mathbb{E} X_{i}\right)\right)^{4}=\sum_{i=1}^{n} \mathbb{E}\left(X_{i}-\mathbb{E} X_{i}\right)^{4}+3 \sum_{i=1}^{n} \sum_{j \neq i} \mathbb{E}\left(X_{i}-\mathbb{E} X_{i}\right)^{2} \mathbb{E}\left(X_{j}-\right.$ $\left.\mathbb{E} X_{j}\right)^{2}$. Recall that when the fourth moment is bounded, then so is second moment. Hence, $\sup _{i \in \mathbb{N}} \mathbb{E}\left(X_{i}-\right.$ $\left.\mathbb{E} X_{i}\right)^{2} \leqslant C$ for some $C \in \mathbb{R}_{+}$. Therefore, from the Markov's inequality, we have

$$
P\left(E_{n}\right) \leqslant \frac{\mathbb{E}\left(S_{n}-\mathbb{E} S_{n}\right)^{4}}{n^{4} \epsilon^{4}} \leqslant \frac{n B+3 n(n-1) C}{n^{4} \epsilon^{4}}
$$

It follows that the $\sum_{n \in \mathbb{N}} P\left(E_{n}\right)<\infty$, and hence by Borel Canteli Lemma, we have

$$
P\left\{E_{n}^{c} \text { for all but finitely many } n\right\}=1
$$

Since, the choice of $\epsilon$ was arbitrary, the result follows.
Theorem 2.3 ( $L^{2}$ strong law of large numbers). Let $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ be a sequence of pair-wise uncorrelated random variables defined on a probability space $(\Omega, \mathcal{F}, P)$ with bounded mean $\mathbb{E} X_{n}$ for all $n \in \mathbb{N}$ and uniformly bounded variance $\sup _{n \in \mathbb{N}} \operatorname{Var}\left(X_{n}\right) \leqslant B<\infty$. Then, the empirical $n$-mean converges to $\lim _{n} \frac{\mathbb{E} S_{n}}{n}$ almost surely.

Proof. For each $n \in \mathbb{N}$, we define events $F_{n} \triangleq E_{n^{2}}$, and

$$
G_{n} \triangleq\left\{\max _{n^{2} \leqslant k<(n+1)^{2}}\left|S_{k}-S_{n^{2}}-\mathbb{E}\left(S_{k}-S_{n^{2}}\right)\right|>n^{2} \epsilon\right\}=\bigcup_{n^{2} \leqslant k<(n+1)^{2}}\left\{\omega \in \Omega:\left|\sum_{i=n^{2}}^{k}\left(X_{i}-\mathbb{E} X_{i}\right)\right|>n^{2} \epsilon\right\}
$$

From the Markov's inequality and union bound, we have

$$
P\left(F_{n}\right) \leqslant \frac{\sum_{i=1}^{n^{2}} \operatorname{Var}\left(X_{i}\right)}{n^{4} \epsilon^{2}} \leqslant \frac{B}{n^{2} \epsilon^{2}}, \quad P\left(G_{n}\right) \leqslant \sum_{k=n^{2}}^{(n+1)^{2}-1} \frac{\left(k-n^{2}+1\right) B}{n^{4} \epsilon^{2}} \leqslant \frac{(2 n+1)^{2} B}{n^{4} \epsilon^{2}}
$$

Therefore, $\sum_{n \in \mathbb{N}} P\left(F_{n}\right)<\infty$ and $\sum_{n \in \mathbb{N}} P\left(G_{n}\right)<\infty$, and hence by Borel Canteli Lemma, we have

$$
\lim _{n} \frac{S_{n^{2}}-\mathbb{E} S_{n^{2}}}{n^{2}}=\lim _{n} \frac{S_{k}-S_{n^{2}}-\mathbb{E}\left(S_{k}-S_{n^{2}}\right)}{n^{2}}=0 \text { a.s. }
$$

The result follows from the fact that for any $k \in \mathbb{N}$, there exists $n \in \mathbb{N}$ such that $k \in\left\{n^{2}, \ldots,(n+1)^{2}-1\right\}$ and hence

$$
\frac{\left|S_{k}-\mathbb{E} S_{k}\right|}{k} \leqslant\left(\frac{\left|S_{n^{2}}-\mathbb{E} S_{n^{2}}\right|}{n^{2}}+\frac{\left|S_{k}-S_{n^{2}}-\mathbb{E}\left(S_{k}-S_{n^{2}}\right)\right|}{n^{2}}\right)
$$

Theorem 2.4 ( $L^{1}$ strong law of large numbers). Let $X: \Omega \rightarrow \mathbb{R}^{\mathbb{N}}$ be a random sequence defined on a probability space $(\Omega, \mathcal{F}, P)$ such that $\sup _{n \in \mathbb{N}} \mathbb{E}\left|X_{n}\right| \leqslant B<\infty$. Then, the empirical n-mean converges to $\lim _{n} \frac{\mathbb{E} S_{n}}{n}$ almost surely.

