

## Tutorial 6

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Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space and let  $X$  be an RV on  $(\Omega, \mathcal{F}, \mathbb{P})$ .

Characteristics function:  $\phi_X(u) \triangleq \mathbb{E} e^{juX}$

Example 1. Let  $X \sim \text{Exp}(\lambda)$ . Find  $\phi_X(u)$ .

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{elsewhere} \end{cases}$$

$$\begin{aligned} \therefore \phi_X(u) &= \mathbb{E} e^{juX} = \int_{\mathbb{R}} e^{jux} \cdot \lambda e^{-\lambda x} dx \\ &= \int_0^{\infty} \lambda e^{x(-\lambda + ju)} dx \\ &= \frac{\lambda}{-\lambda + ju} e^{x(-\lambda + ju)} \Big|_0^{\infty} \\ &= \frac{\lambda}{\lambda - ju} \end{aligned}$$

Example 2. Let  $X \sim \text{Pois}(\lambda)$ . Find  $\phi_X(u)$ .

$$P_X(k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k \in \mathbb{Z}_+$$

$$\begin{aligned} \therefore \mathbb{E} e^{juX} &= \sum_{k=0}^{\infty} e^{juk} e^{-\lambda} \frac{\lambda^k}{k!} = \sum_{k=0}^{\infty} e^{-\lambda} \frac{(\lambda e^{ju})^k}{k!} \\ &= e^{-\lambda} \cdot e^{\lambda e^{ju}} \\ &= e^{\lambda(e^{ju} - 1)} \end{aligned}$$

Exercise 1.  $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ ,  $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$  are uncorrelated. Show that  $X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$ .

Sol<sup>n</sup>

We know that if two Gaussian random variables

are uncorrelated, they are independent too.

We will show that if  $X_1, X_2$  are independent,

$$\phi_{X_1+X_2}(u) = \phi_{X_1}(u) \cdot \phi_{X_2}(u)$$

If  $X_1, X_2$  are independent, then  $e^{juX_1}$  and  $e^{juX_2}$  are independent. Hence  $e^{juX_1}, e^{juX_2}$  are uncorrelated, i.e.

$$\mathbb{E}[e^{juX_1} \cdot e^{juX_2}] = \mathbb{E}[e^{juX_1}] \mathbb{E}[e^{juX_2}]$$

$$\Rightarrow \phi_{X_1+X_2}(u) = \mathbb{E}[e^{ju(X_1+X_2)}] = \phi_{X_1}(u) \cdot \phi_{X_2}(u)$$

Now, as  $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ ,  $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$ ,

$$\begin{aligned} \phi_{X_1+X_2}(u) &= \exp\left(-\frac{u^2\sigma_1^2}{2} + ju\mu_1\right) \exp\left(-\frac{u^2\sigma_2^2}{2} + ju\mu_2\right) \\ &= \exp\left(-\frac{u^2}{2}(\sigma_1^2 + \sigma_2^2) + ju(\mu_1 + \mu_2)\right) \end{aligned}$$

$$\Rightarrow X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2) \quad \square$$

Probability Generating Function:  $\Psi_X(z) = \sum_{x \in \mathcal{X}} z^x P_X(x)$   
where  $\mathcal{X} \subseteq \mathbb{Z}_+$

Example 3. Let  $X \sim \text{Binomial}(n, p)$ . Find  $\Psi_X(z)$

$$P_X(k) = \begin{cases} \binom{n}{k} p^k (1-p)^{n-k} & k = 0, 1, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \therefore \Psi_X(z) &= \sum_{k=0}^n \binom{n}{k} (pz)^k (1-p)^{n-k} \\ &= (1-p + pz)^n \end{aligned}$$

### Conditional Expectation

\*  $\mathbb{E}[X|Y]$  is a random variable.

Example: Let  $X$  and  $Y$  take values on  $\{1, 2\}$  and  $\{-1, 1\}$

Example: Let  $X$  and  $Y$  take values on  $\{1, 2\}$  and  $\{-1, 1\}$  respectively with joint mass function  $P_{xy}$  s.t.

$$P_{xy}(1, -1) = \frac{1}{5}, \quad P_{xy}(1, 1) = \frac{1}{3}$$

$$P_{xy}(2, -1) = \frac{1}{5}, \quad P_{xy}(2, 1) = \frac{4}{15}$$

Find  $E[X|Y]$  and  $E[Y|X]$  and their distributions.

Sol<sup>n</sup>.  $P_{X|Y}(x=1 | y=-1) = \frac{P_{xy}(1, -1)}{P_y(-1)} = \frac{1}{2}$

$$P_{X|Y}(x=2 | y=-1) = \frac{P_{xy}(2, -1)}{P_y(-1)} = \frac{1}{2}$$

$$P_{X|Y}(x=1 | y=1) = \frac{P_{xy}(1, 1)}{P_y(1)} = \frac{\frac{1}{3}}{\frac{3}{5}} = \frac{5}{9}$$

$$P_{X|Y}(x=2 | y=1) = \frac{P_{xy}(2, 1)}{P_y(1)} = \frac{4}{9}$$

$$\begin{aligned} E[X | Y = -1] &= 1 \cdot P_{X|Y}(x=1 | y=-1) + 2 \cdot P_{X|Y}(x=2 | y=-1) \\ &= \frac{3}{2} \end{aligned}$$

$$\begin{aligned} E[X | Y = 1] &= 1 \cdot P_{X|Y}(x=1 | y=1) + 2 \cdot P_{X|Y}(x=2 | y=1) \\ &= 1 \cdot \frac{5}{9} + 2 \cdot \frac{4}{9} = \frac{13}{9} \end{aligned}$$

look at the set

$$\left\{ \omega : E[X(\omega) | Y(\omega)] = \frac{3}{2} \right\} = \left\{ \omega : Y(\omega) = -1 \right\}$$

$$\begin{aligned} \Rightarrow P\left\{ \omega : E[X(\omega) | Y(\omega)] = \frac{3}{2} \right\} &= P\left\{ \omega : Y(\omega) = -1 \right\} \\ &= P(Y = -1) = \frac{2}{5} \end{aligned}$$

$$\text{Similarly, } P\left\{ \omega : E[X(\omega) | Y(\omega)] = \frac{13}{9} \right\} = \frac{3}{5}$$

$$\therefore E[X|Y] = \begin{cases} \frac{3}{2} & \text{w.p. } \frac{2}{5} \\ \frac{13}{9} & \text{w.p. } \frac{3}{5} \end{cases}$$

Verify  $E[E[X|Y]] = E[X]$

$$\mathbb{E}[\mathbb{E}[X|Y]] = \frac{3}{2} \cdot \frac{2}{5} + \frac{13}{9} \cdot \frac{3}{5} = \frac{22}{15} = \mathbb{E}[X]$$

· Law of iterated Expectation :

\* We know  $\mathbb{E}[\mathbb{E}[X|G]] = \mathbb{E}X$ .

\* Also  $\mathbb{E}[X|Y] := \mathbb{E}[X|\sigma(Y)]$

$\therefore \mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}X$ .

Exercise 2.  $Z = \sum_{i=1}^N X_i$ .  $X_i$ 's are iid and  $N: \Omega \rightarrow \mathbb{Z}_+$  is independent of  $X_i$ 's.  $M_X$  and  $M_N$  are given. Find  $M_Z$ .

Sol<sup>n</sup>.  $M_Z(t) = \mathbb{E}\left[e^{t \sum_{i=1}^N X_i}\right]$

$$= \mathbb{E}\left[\prod_{i=1}^N e^{t X_i}\right]$$

$$= \mathbb{E}\left[\mathbb{E}\left[\prod_{i=1}^N e^{t X_i} \mid N\right]\right]$$

$$= \mathbb{E}\left[\prod_{i=1}^N \mathbb{E}[e^{t X_i}]\right]$$

$$= \mathbb{E}\left[M_X(t)^N\right]$$

$$= \mathbb{E}\left[e^{N \log M_X(t)}\right]$$

$$= M_N(\log M_X(t))$$

or,  $M_Z(t) = \mathbb{E}\left[M_X(t)^N\right] = \Psi_N[M_X(t)]$

Exercise 3. Take  $X_i \sim \text{EXP}(\mu)$ ,  $N \sim \text{Geom}(p)$  and  $Z = \sum_{i=1}^N X_i$ . Find distribution of  $Z$  given  $X_i$ 's are independent.

Sol<sup>n</sup>.  $M_X(t) = \frac{\mu}{\mu - s}$ ,  $s < \mu$

$$\Psi_N(z) = \sum_{k=1}^{\infty} z^k (1-p)^{k-1} p = \frac{pz}{1 - z(1-p)}, \quad |z| < \frac{1}{1-p}$$

$$\Psi_N(z) = \sum_{k=0}^{\mu-s} z^k (1-p)^{k-1} \cdot p = \frac{pz}{1-z(1-p)}, \quad |z| < \frac{1}{1-p}$$

$$\begin{aligned} \therefore M_Z(t) &= \Psi_N(M_X(t)) \\ &= \frac{p\left(\frac{\mu}{\mu-s}\right)}{1-\left(\frac{\mu}{\mu-s}\right)(1-p)}, \quad \left|\frac{\mu}{\mu-s}\right| < \frac{1}{1-p} \end{aligned}$$

$$\begin{aligned} \left|\frac{\mu}{\mu-s}\right| &= \frac{\mu}{\mu-s} < \frac{1}{1-p} \Rightarrow \mu - \mu p < \mu - s \\ &\Rightarrow s < \mu p \end{aligned}$$

$$M_Z(t) = \frac{\mu p}{\mu p - s}, \quad s < \mu p$$

$$\Rightarrow Z \sim \text{Exp}(\mu p)$$