

## Tutorial 8

Monday, 19 September 2022 8:58 PM

### Recap:

1. Pointwise convergence:  $X_n \xrightarrow{P.W.} X$  if  $\lim_{n \rightarrow \infty} X_n(\omega) = X(\omega) \forall \omega \in \Omega$ .
2. Almost sure convergence:  $X_n \xrightarrow{a.s.} X$  if  $P(\{\omega \in \Omega : X_n(\omega) \rightarrow X(\omega)\}) = 1$
3. Convergence in probability:  $X_n \xrightarrow{i.p.} X$  if  $\lim_{n \rightarrow \infty} P(A_n(\varepsilon)) = 0$   
where  $A_n(\varepsilon) = \{\omega \in \Omega : |X_n(\omega) - X(\omega)| > \varepsilon\}$
4.  $P.W. \Rightarrow a.s. \Rightarrow i.p.$ ,  $i.p. \not\Rightarrow a.s.$  in general.
5.  $(A_n \text{ i.o.}) = \limsup_n A_n$  and  $(A_n \text{ a.b.f.m.}) = \liminf_n A_n$
6. Borel Cantelli Lemma:  $\sum_{n \in \mathbb{N}} P(A_n) < \infty \Rightarrow P(A_n \text{ i.o.}) = 0$
- 7 Borel 0-1 law: If  $A_n$ s are independent then
$$P(A_n \text{ i.o.}) = \begin{cases} 0 & \text{if } \sum_{n \in \mathbb{N}} P(A_n) < \infty \\ 1 & \text{if } \sum_{n \in \mathbb{N}} P(A_n) = \infty \end{cases}$$
8. If  $\forall \varepsilon > 0 \sum_n P(A_n(\varepsilon)) < \infty$ , then  $X_n \xrightarrow{a.s.} X$

Proof. For fixed  $\varepsilon > 0$ , by Borel-Cantelli lemma

$$P(A_n(\varepsilon) \text{ i.o.}) = P(\limsup_n A_n(\varepsilon)) = 0$$

$$\Rightarrow P((\limsup_n A_n(\varepsilon))^c) = 1$$

$$\Rightarrow P(\liminf_n A_n(\varepsilon)^c) = 1$$

$$\text{Let } \liminf_n A_n(\varepsilon)^c =: B_\varepsilon$$

We know that

$$N^c = \{\omega \in \Omega : X_n(\omega) \rightarrow X(\omega)\}$$

$$= \{\omega \in \Omega : \text{for every } p \in \mathbb{N} \exists N \in \mathbb{N} \text{ s.t. } \forall n \geq N \\ |X_n(\omega) - X(\omega)| < \frac{1}{p}\}$$

$$= \bigcap \{\omega \in \Omega : \omega \in (A_n(\frac{1}{p})^c \text{ a.b.f.m.})\}$$

$$= \bigcap_{p \in \mathbb{N}} \left\{ \omega \in \Omega : \omega \in \left( A_n \left( \frac{1}{p} \right) \right)^c \text{ a.b.f.m.} \right\}$$

$$= \bigcap_{p \in \mathbb{N}} B_{\frac{1}{p}}$$

$$\Rightarrow \mathbb{P}(N^c) = \mathbb{P}\left(\bigcap_{p \in \mathbb{N}} B_{\frac{1}{p}}\right) = 1 \quad \square$$

Note: for every  $\varepsilon > 0$ ,  $\sum_{n \in \mathbb{N}} \mathbb{P}(A_n(\varepsilon)) < \infty \Rightarrow \lim_{n \rightarrow \infty} A_n(\varepsilon) = 0$  for any  $\varepsilon > 0$   
 $\Rightarrow X_n \xrightarrow{i.p.} X$ .

Note: We showed that for every  $\varepsilon > 0$ ,  $\mathbb{P}(A_n(\varepsilon) \text{ i.o.}) = 0 \Rightarrow X_n \xrightarrow{a.s.} X$   
 It is also possible to show that  $X_n \xrightarrow{a.s.} X \Rightarrow$  for every  $\varepsilon > 0$   
 $\mathbb{P}(A_n(\varepsilon) \text{ i.o.}) = 0$ .

\* Verify for the examples given in the lecture.

9. Convergence in  $L^p$ :  $X_n \xrightarrow{d^p} X$  if  $\lim_{n \rightarrow \infty} \mathbb{E}|X_n - X|^p = 0$

For the particular case of  $p = 2$ . we say that  $X_n \xrightarrow{m.s.} X$ .

10.  $L^p \Rightarrow$  i.p.

11.  $L^p \not\Rightarrow$  a.s.

12.  $L^2$  weak law of large numbers:  $X_n$ 's are uncorrelated.

$\mathbb{E}X_n = \mu$ ,  $\text{Var}(X_n) = \sigma^2 \forall n \in \mathbb{N}$ .  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ . Then  
 $\bar{X}_n \xrightarrow{d^p, i.p.} \mu$

Note:  $L^2$  as  $\text{Var}(X_n) = \sigma^2$ , weak  $\Rightarrow$  not almost sure.

13.  $L^1$  weak law of large numbers:  $X_n$ 's are i.i.d,  $\mathbb{E}X_n = \mu$   
 $\forall n \in \mathbb{N}$ , then  $\bar{X}_n \xrightarrow{i.p.} \mu$ .

14. Monotone Convergence Theorem:  $X_n$ 's are non-decreasing,  $L^1$ .

If  $\lim_{n \rightarrow \infty} X_n(\omega) = X(\omega) \forall \omega \in \Omega$  then  $\lim_{n \rightarrow \infty} \mathbb{E}X_n = \mathbb{E}X$ .

15. Dominated Convergence Theorem:  $X_n$ 's are in  $L^1$ ,  $|X_n| \leq Y \in L^1$

and  $\lim_{n \rightarrow \infty} X_n(\omega) = X(\omega) \forall \omega \in \Omega$  then  $\lim_{n \rightarrow \infty} \mathbb{E}X_n = \mathbb{E}X$ .

16. Convergence in distribution:  $X_n \xrightarrow{D} X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x) \forall x \in \mathbb{R}$

16. Convergence in distribution:  $X_n \xrightarrow{D} X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x) \forall x \in \mathbb{R}$

17.  $X_n \xrightarrow{D} X \Leftrightarrow \lim_{n \rightarrow \infty} \phi_{X_n}(u) = \phi_X(u)$  for every  $u \in \mathbb{R}$ .

Exercise 1 Suppose that  $((X_1, Y_1), (X_2, Y_2), \dots)$  is a sequence of random vectors s.t.  $P(\{X_n \geq Y_n\}) = \alpha^n$  for some  $\alpha \in (0, 1)$ . Show that  $P(\{X_n \geq Y_n\} \text{ i.o.}) = 0$ .

Sol<sup>n</sup>.  $\sum_{n \in \mathbb{N}} P(\{X_n \geq Y_n\}) = \frac{1}{1-\alpha} < \infty$ . Hence

$$P(\{X_n \geq Y_n\} \text{ i.o.}) = 0.$$

Exercise 2:  $(X_n : n \in \mathbb{N})$  is a sequence of independent random variables with marginal pmfs given by

$$P(X_n = \frac{1}{2}(1 - \frac{1}{n})) = P(X_n = \frac{1}{2}(1 - \frac{1}{n})) = \frac{1}{2}$$

a. Show that the sequence converge in almost sure sense and find the limit.

b. Does  $X_n$  converge in  $d^2$ ?

Sol<sup>n</sup>. a. For fixed  $p \in \mathbb{N}$ ,

$$\begin{aligned} P(A_n(\frac{1}{p})) &= P(\{\omega \in \Omega \mid |X_n(\omega) - \frac{1}{2}| > \frac{1}{p}\}) \\ &= P(\{\omega \in \Omega \mid \frac{1}{2} - \frac{1}{p} \leq X_n(\omega) \leq \frac{1}{2} + \frac{1}{p}\}^c) \\ &= P(\{\omega \in \Omega \mid \frac{1}{2}(1 - \frac{2}{p}) \leq X_n(\omega) \leq \frac{1}{2}(1 - \frac{2}{p})\}^c) \\ &= \begin{cases} 0 & \text{if } n \geq \frac{p}{2} \\ 1 & \text{if } n < \frac{p}{2} \end{cases} \end{aligned}$$

We have two observations

1.  $\lim_{n \rightarrow \infty} P(A_n(\frac{1}{p})) = 0$  for every  $p \in \mathbb{N}$

2.  $\lim \sum_n P(A_n(\frac{1}{p})) < \frac{p}{2} < \infty$  for every  $p \in \mathbb{N}$

1.  $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \mathbb{P}(A_i(\frac{1}{p})) = 0$  for every  $p \in \mathbb{N}$

2.  $\lim_{n \rightarrow \infty} \sum_{i=1}^n \mathbb{P}(A_i(\frac{1}{p})) < \frac{p}{2} < \infty$  for every  $p \in \mathbb{N}$

$\Rightarrow X_n \xrightarrow{\text{a.s.}} \frac{1}{2}$ .

$$\begin{aligned} \text{b. } \mathbb{E} \left[ \left( X_n - \frac{1}{2} \right)^2 \right] &= \left( \frac{1}{2} \left( 1 - \frac{1}{n} \right) - \frac{1}{2} \right)^2 \cdot \frac{1}{2} + \left( \frac{1}{2} \left( 1 + \frac{1}{n} \right) - \frac{1}{2} \right)^2 \\ &= \frac{1}{4n^2} \end{aligned}$$

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[ \left( X_n - \frac{1}{2} \right)^2 \right] = \lim_{n \rightarrow \infty} \frac{1}{4n^2} = 0$$

Hence,  $X_n \xrightarrow{\mathcal{L}^2} 0$

Exercise 3: Let  $U \sim \text{Unif}([0,1])$  and let  $X_n = \frac{(-1)^n U}{n}$  for  $n \geq 1$ .

a. Show that  $X_n$  converges in almost sure sense. Find the limit.

b. Show that  $X_n$  converges in mean-squared sense. Find the limit.

Sol<sup>n</sup>: a. Claim:  $X_n \xrightarrow{\text{a.s.}} 0$

Fix  $\varepsilon > 0$  arbitrarily.

Let  $A_n(\varepsilon) = \{ \omega \in \Omega \mid |X_n(\omega)| > \varepsilon \}$ . If  $\mathbb{P}(A_n(\varepsilon) \text{ i.o.}) = 0$

then,  $X_n \xrightarrow{\text{a.s.}} 0$ .

We will show that  $\mathbb{P}(\liminf_n A_n(\varepsilon)^c) = 1$ .

$$\liminf_n A_n(\varepsilon)^c = \bigcup_{n \in \mathbb{N}} \bigcap_{m \geq n} \{ \omega \in \Omega \mid |X_m(\omega)| \leq \varepsilon \}$$

$$= \bigcup_{n \in \mathbb{N}} \bigcap_{m \geq n} \{ \omega \in \Omega \mid U(\omega) \leq m\varepsilon \}$$

$$= \bigcup_{n \in \mathbb{N}} \{ \omega \in \Omega \mid U(\omega) \leq n\varepsilon \}$$

$$= \{ \omega \in \Omega \mid U(\omega) \in \mathbb{R} \}$$

$$= \Omega$$

$$\Rightarrow P\left(\liminf_n A_n(\varepsilon)^c\right) = 1$$

It is to see that,  $\liminf_n A_n^c = (\limsup_n A_n)^c$

$$\therefore P(\limsup_n A_n(\varepsilon)) = P(A_n(\varepsilon) \text{ i.o.}) = 0$$

$$\Rightarrow X_n \xrightarrow{\text{a.s.}} X.$$

$$b. E[|X_n|^2] = \frac{1}{n^2} E[U^2] = \frac{1}{3n^2}$$

$$\Rightarrow \lim_{n \rightarrow \infty} E[|X_n|^2] = 0$$

$$\therefore X_n \xrightarrow{\text{m.s.}} 0.$$

Exercise 4: Let  $\Theta \sim \text{Unif}([0, 2\pi])$ . Show that  $X_n = \cos(n\Theta)$  does not converge in almost sure sense.

Sol<sup>n</sup>: Define  $N^c = \{\omega \in \Omega \mid \cos(n\Theta(\omega)) \text{ converges}\}$ .

Let for some  $\theta \in [0, 2\pi]$ ,  $\cos(n\theta)$  converges to  $c \in [-1, 1]$

We know that

$$\cos((n+1)\theta) + \cos((n-1)\theta) = 2\cos(n\theta)\cos\theta$$

Taking limit  $n \rightarrow \infty$ ,

$$c + c = 2c\cos\theta$$

$$\Rightarrow \cos\theta = 1 \quad \Rightarrow \theta = 0, 2\pi.$$

$$\therefore N^c = \{\omega \in \Omega \mid \Theta(\omega) = 0 \text{ or } \Theta(\omega) = 2\pi\}$$

$\Rightarrow P(N^c) = 0$ , hence  $(X_n)$  does not converge in almost sure sense.

Exercise 5: Let  $\Theta \sim \text{Unif}([0, 2\pi])$ . Show that  $|1 - \frac{\Theta}{\pi}|^n$  converges to 0 in distribution.

Sol<sup>n</sup>.

$$\begin{aligned}
\mathbb{P}\left(\left|1 - \frac{\theta}{\pi} - 0\right| > \varepsilon\right) &= 1 - \mathbb{P}\left(\left|1 - \frac{\theta}{\pi}\right| \leq \varepsilon^{1/n}\right) \\
&= 1 - \mathbb{P}\left(1 - \varepsilon^{1/n} < \frac{\theta}{\pi} < 1 + \varepsilon^{1/n}\right) \\
&= 1 - \mathbb{P}\left(\pi(1 - \varepsilon^{1/n}) < \theta < \pi(1 + \varepsilon^{1/n})\right) \\
&= 1 - \varepsilon^{1/n}
\end{aligned}$$

$$\lim_{n \rightarrow \infty} \mathbb{P}\left(\left|1 - \frac{\theta}{\pi} - 0\right| > \varepsilon\right) = 0 \Rightarrow \left|1 - \frac{\theta}{\pi}\right|^n \xrightarrow{i.p.} 0$$

$$\text{Hence, } \left|1 - \frac{\theta}{\pi}\right|^n \xrightarrow{D} 0$$

Exercise 6.  $X_n \sim \text{Geom}\left(\frac{\lambda}{n}\right)$ ,  $\lambda > 0$ .  $Y_n = \frac{X_n}{n}$ . Show that  $Y_n$  converges to  $Y \sim \text{Exp}(\lambda)$  in distribution.

Sol<sup>n</sup>:  $X_n \sim \text{Geom}\left(\frac{\lambda}{n}\right)$ . Distribution of  $X_n$ ,

$$\begin{aligned}
\mathbb{P}(X_n \leq m) &= \sum_{i=1}^m \left(1 - \frac{\lambda}{n}\right)^{i-1} \frac{\lambda}{n} \\
&= \frac{\lambda}{n} \cdot \frac{1 - \left(1 - \frac{\lambda}{n}\right)^m}{\frac{\lambda}{n}} = 1 - \left(1 - \frac{\lambda}{n}\right)^m
\end{aligned}$$

Let us find distribution of  $Y_n$ ,

$$\mathbb{P}(Y_n \leq y) = \mathbb{P}(X_n \leq ny) = 1 - \left(1 - \frac{\lambda}{n}\right)^{\lfloor ny \rfloor}$$

$$\lim_{n \rightarrow \infty} \mathbb{P}(Y_n \leq y) = 1 - \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{\lfloor ny \rfloor} = 1 - e^{-\lambda y}$$

$$\Rightarrow Y_n \xrightarrow{D} Y \sim \text{Exp}(\lambda). \quad \square$$