

# Lecture-02: Embedded Markov Chain and Sojourn Times

## 1 State Evolution

For a homogeneous Markov process  $X : \Omega \rightarrow \mathcal{X}_+^{\mathbb{R}}$  on countable state space  $\mathcal{X} \subseteq \mathbb{R}$  with sample paths that are right continuous with left limits (rcll), we wish to characterize the transition kernel  $P : \mathbb{R}_+ \rightarrow [0, 1]^{\mathcal{X} \times \mathcal{X}}$ , where  $P_{xy}(t) \triangleq P(\{X_{s+t} = y\} \mid \{X_s = x\})$  for all  $s, t \in \mathbb{R}_+$ . To this end, we define the sojourn time in any state, the jump times, and the jump transition probabilities.

### 1.1 Transition instants and counting process

**Definition 1.1.** Let  $S_0 \triangleq 0$ . The  $n$ th **jump time** of a right continuous countable state stochastic process  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}_+}$  is defined inductively as  $S_n \triangleq \inf \{t > S_{n-1} : X_t \neq X_{S_{n-1}}\}$ .

**Definition 1.2.** The counting process associated with jump times sequence  $S : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  is denoted by  $N : \Omega \rightarrow \mathbb{Z}_+^{\mathbb{R}_+}$ , where the number of jumps in duration  $(0, t]$  is denoted by  $N_t \triangleq \sum_{n \in \mathbb{N}} \mathbb{1}_{\{S_n \leq t\}}$ .

**Lemma 1.3.** Each term of the jump time sequence  $S : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  is adapted to the natural filtration  $\mathcal{F}_\bullet$  of the process  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}_+}$ .

*Proof.* Since  $X$  is rcll, it is progressively measurable, and hence the event  $\{S_n \leq t\} \in \mathcal{F}_t$ .  $\square$

**Definition 1.4 (Age of a state).** For the counting process  $N$  associated with the CTMC  $X$ , we can define age process  $A : \Omega \rightarrow \mathbb{R}_+^{\mathbb{R}_+}$  where the age of last transition at time  $t$  is denoted by  $A_t \triangleq t - S_{N_t}$  for all times  $t \in \mathbb{R}_+$ . For the CTMC  $X$ , we can write the age of last transition at time  $t \in \mathbb{R}_+$  as

$$A_t \triangleq \inf \{s > 0 : X_{t-s} \neq X_t\}.$$

**Definition 1.5 (Excess time in a state).** For the counting process  $N$  associated with the CTMC  $X$ , we can define the excess time process  $Y : \Omega \rightarrow \mathbb{R}_+^{\mathbb{R}_+}$  where  $Y_t \triangleq S_{N_t+1} - t$  is the time until next transition at time  $t$ . We can write the excess time at time  $t \in \mathbb{R}_+$  for the CTMC  $X$  as

$$Y_t \triangleq \inf \{s > 0 : X_{t+s} \neq X_t\}.$$

*Remark 1.* For a homogeneous CTMC  $X$ , the distribution of excess time  $Y_t$  conditioned on the current state  $X_t$ , doesn't depend on time  $t$ . Hence, we can define the following conditional complementary distribution of excess time as  $\bar{F}_x(u) \triangleq P(\{Y_t > u\} \mid \{X_t = x\}) = P_x \{Y_0 > u\}$ .

**Lemma 1.6.** For a homogeneous CTMC  $X$ , there exists a positive sequence  $\nu \in \mathbb{R}_+^{\mathcal{X}}$ , such that

$$\bar{F}_x(u) \triangleq P(\{Y_t > u\} \mid \{X_t = x\}) = e^{-u\nu_x}, \quad x \in \mathcal{X}.$$

*Proof.* We fix a state  $x \in \mathcal{X}$ , and observe that the function  $\bar{F}_x : \mathbb{R}_+ \rightarrow [0, 1]$  is non-negative, non-increasing, and right-continuous. Using the Markov property and the time-homogeneity, we can show that  $\bar{F}_x$  satisfies the semigroup property. In particular,

$$\bar{F}_x(u+v) = P(\{Y_t > u+v\} \mid \{X_t = x\}) = P(\{Y_t > u, Y_{t+u} > v\} \mid \{X_t = x\}) = \bar{F}_x(u)\bar{F}_x(v).$$

The only continuous function  $\bar{F}_x : \mathbb{R}_+ \rightarrow [0, 1]$  that satisfies the semigroup property is an exponential function with a negative exponent.  $\square$

**Example 1.7 (Poisson process).** Consider the counting process  $N : \Omega \rightarrow \mathbb{Z}_+^{\mathbb{R}_+}$  for a Poisson point process with homogeneous rate  $\lambda$ . Using the stationary independent increment property, we have for all  $u \geq 0$

$$\bar{F}_i(u) = P(\{Y_t > u\} \mid \{N_t = i\}) = P(\{N_{t+u} = i\} \mid \{N_t = i\}) = P\{N_{t+u} - N_t = 0\} = P\{Y_t > u\} = e^{-\lambda u}.$$

## 1.2 Sojourn time in a state

**Definition 1.8.** The **jump process**  $Z : \Omega \rightarrow \mathcal{X}^{\mathbb{N}}$  is a discrete time process, derived from the continuous time stochastic process  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}_+}$  by sampling  $X$  at the jump times  $S : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$ . The state of the process  $X$  at the  $n$ th jump time  $S_n$  is the  $n$ th state  $Z_n \triangleq X_{S_n}$  of the jump process  $Z$ .

**Definition 1.9.** The **sojourn time** in the state  $Z_{n-1}$  for the process  $X$  is defined as  $T_n \triangleq S_n - S_{n-1}$ .

*Remark 2.* From the definition of jump instants, it follows that the history until time  $t$  is

$$\mathcal{F}_t = \sigma(S_0, (Z_0, T_1), (Z_1, T_2), \dots, (Z_{N_t}, A_t)).$$

We can verify that  $\mathcal{F}_{S_n} = \sigma(S_0, (Z_0, T_1), \dots, (Z_{n-1}, T_n), Z_n)$ .

**Lemma 1.10.** For a homogeneous CTMC, each sojourn time  $T_n : \Omega \rightarrow \mathbb{R}_+$  is a continuous memoryless random variable, and the sequence of sojourn times  $(T_j : j \geq n)$  is independent of the past  $\mathcal{F}_{S_{n-1}}$  conditioned on  $Z_{n-1}$ .

*Proof.* We observe that the sojourn time  $T_n$  equals the excess time  $Y_{S_{n-1}}$ , where the process remains in state  $Z_{n-1} = X_{S_{n-1}}$  in the duration  $S_{n-1} + [0, T_n)$ . Using the strong Markov property, we can write the conditional complementary distribution of  $T_n$  given any historical event  $H \in \mathcal{F}_{S_{n-1}}$  and  $u \geq 0$  as

$$P(\{T_n > u\} | \{Z_{n-1} = x\} \cap H) = P(\{Y_{S_{n-1}} > u\} | \{X_{S_{n-1}} = x\} \cap H) = \exp(-u\nu_x) = \bar{F}_x(u).$$

□

**Corollary 1.11.** If  $X_n = x$ , then the holding time  $T_{n+1}$  is an exponential random variable with rate  $\nu_x$ .

**Definition 1.12.** For a homogeneous CTMC  $X$ , the exponential rate for the random holding time in a state  $x$  is called the **transition rate** out of state  $x$  denoted by  $\nu_x$ .

**Definition 1.13.** For a CTMC  $X$ , a state  $x \in \mathcal{X}$  is called **instantaneous** if  $\nu_x = \infty$ , **stable** if  $\nu_x \in (0, \infty)$ , and **absorbing** if  $\nu_x = 0$ .

*Remark 3.* Transition rate out of a state  $x$  is the inverse of mean holding time in this state  $x$ , i.e.  $\nu_x = \frac{1}{\mathbb{E}_x T_1}$ . Therefore, the mean holding time  $\mathbb{E}_x T_1$  in state  $x$  is  $\infty$  in an absorbing state, zero in an instantaneous state, and almost surely finite and non-zero in a stable state.

**Definition 1.14.** A homogeneous CTMC with no instantaneous states is called a **pure jump CTMC**.

**Definition 1.15.** A pure jump CTMC with

- (i) all stable states and  $\inf_{x \in \mathcal{X}} \nu_x \geq \nu > 0$  is called **stable**, and
- (ii)  $\sup_{x \in \mathcal{X}} \nu_x \leq \nu < \infty$  is called **regular**.

**Example 1.16 (Non-regular CTMC).** For the countable state space  $\mathbb{N}$ , consider the probability transition matrix  $P$  such that  $p_{n,n+1} = 1$  and the exponential holding times with rate  $\nu_n = n^2$  for each state  $n \in \mathbb{N}$ . Clearly,  $\sup_{n \in \mathbb{N}} \nu_n = \infty$ , and hence it is not regular.

*Remark 4.* Pure jump homogeneous CTMC with finite stable states are stable and regular. We will focus on pure jump homogeneous CTMC over countably infinite states, that are stable and regular.

## 1.3 Jump process

**Proposition 1.17.** For a stable CTMC, the jump times are stopping times.

*Proof.* For a stable CTMC  $X$ , we let  $0 < \nu \leq \inf_{x \in \mathcal{X}} \nu_x$ . Then, by coupling in Appendix B, we have a sequence of *i.i.d.* random variables  $\bar{T} : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$ , such that  $T_n \leq \bar{T}_n$  almost surely and  $\mathbb{E} \bar{T}_n = \frac{1}{\nu}$  for each  $n \in \mathbb{N}$ . Defining  $\bar{S}_n \triangleq \sum_{i=1}^n \bar{T}_i$ , it follows that  $S_n \leq \bar{S}_n$  for each  $n \in \mathbb{N}$ . Since  $\bar{S}_n$  is sum of  $n$  almost surely finite random variables, it is finite almost surely. It follows that  $S_n$  is finite almost surely. □

**Proposition 1.18.** For a regular CTMC,  $N_t$  is almost surely finite for all finite times  $t \in \mathbb{R}_+$ .

*Proof.* Let  $X$  be a regular CTMC and  $\sup_{x \in \mathcal{X}} \nu_x \leq \nu < \infty$ . Then, by coupling in Appendix B, we have a sequence of *i.i.d.* random variables  $\underline{T} : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$ , such that  $T_n \geq \underline{T}_n$  almost surely and  $\mathbb{E}\underline{T}_n = \frac{1}{\nu}$  for each  $n \in \mathbb{N}$ . Defining  $\underline{S}_n \triangleq \sum_{i=1}^n \underline{T}_i$  and  $\underline{N}_t \triangleq \sum_{n \in \mathbb{N}} \mathbb{1}_{\{\underline{S}_n \leq t\}}$ , it follows that  $S_n \geq \underline{S}_n$  for each  $n \in \mathbb{N}$  and  $N_t \leq \underline{N}_t$  for all  $t \in \mathbb{R}_+$ . Since  $\underline{N}$  is a Poisson counting process with finite rate  $\nu$ , it is almost surely finite at all  $t \in \mathbb{R}_+$  and the result follows.  $\square$

*Remark 5.* From the strong Markov property and the time-homogeneity of the CTMC  $X$ , we see that

$$P(\{Z_n = y\} \mid \{Z_{n-1} = x\}) = P_{xy}(S_{n-1}, S_n) = P_{xy}(0, T_n).$$

*Remark 6.* From the law of total probability, it follows that for any rcll stochastic process  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}_+}$  with countable state space  $\mathcal{X}$ , the sum of jump transition probabilities  $\sum_{y \neq x} P_{xy}(S_{n-1}, S_n) = 1$  for all states  $X_{S_{n-1}} = x \in \mathcal{X}$ .

**Lemma 1.19.** *For a homogeneous CTMC  $X$ , the jump probability from state  $Z_{n-1}$  to state  $Z_n$  depends solely on  $Z_{n-1}$  and is independent of jump instants.*

*Proof.* Fix states  $x, y \in \mathcal{X}$  and a historical event  $H \in \mathcal{F}_{S_{n-1}}$ . From the definition of conditional probability, we write

$$P(\{T_n > u, Z_n = y\} \mid \{Z_{n-1} = x\} \cap H) = P(\{X_{S_n} = y\} \mid \{T_n > u, X_{S_{n-1}} = x\} \cap H) P(\{T_n > u\} \mid \{Z_{n-1} = x\} \cap H).$$

From the strong Markov property of  $X$ , we get  $P(\{T_n > u\} \mid \{Z_{n-1} = x\} \cap H) = \bar{F}_x(u)$ . We further observe that  $\{T_n > u, X_{S_{n-1}} = x\} \cap H = \{X_t = x, t \in S_{n-1} + [0, u]\} \cap H \in \mathcal{F}_{S_{n-1}+u}$ . From the definition of excess time, we can write  $S_n = S_{n-1} + u + Y_{S_{n-1}+u}$  for any  $u \in [0, T_n]$ . Further, from the strong Markov and the time-homogeneity of CTMC  $X$ , and the memoryless property of excess time  $Y$ , we obtain

$$P(\{X_{S_n} = y\} \mid \{T_n > u, X_{S_{n-1}} = x\} \cap H) = P(\{X_{S_{n-1}+u+Y_{S_{n-1}+u}} = y\} \mid \{X_{S_{n-1}+u} = x\}) = P_{xy}(0, Y_0).$$

This implies that sojourn times and jump instant probabilities are independent.  $\square$

**Definition 1.20.** The jump process  $Z$  is also sometimes referred to as the **embedded DTMC** of the pure jump CTMC  $X$ . The corresponding **jump transition probabilities** are defined

$$p_{xy} \triangleq P_{xy}(S_{n-1}, S_n) = P(\{X_{S_n} = y\} \mid \{X_{S_{n-1}} = x\}), \quad x, y \in \mathcal{X}.$$

*Remark 7.* If  $\nu_x = 0$ , then for any  $u \geq 0$ , we have  $P(\{Y_0 > u\} \mid \{X_0 = x\}) = 1$ , and hence  $S_1 = \infty$  almost surely whenever  $X_0 = x$ . By convention, we set  $p_{xx} = 1$  and  $p_{xy} = 0$  for all states  $y \neq x$ .

**Corollary 1.21.** *The matrix  $p \triangleq (p_{xy} : x, y \in \mathcal{X})$  is stochastic, and if  $\nu_x > 0$  then  $p_{xx} = 0$ .*

*Proof.* Recall  $p_{xy} = P_{xy}(S_1)$ . If  $\nu_x > 0$ , then  $\lim_{u \rightarrow \infty} P(\{Y_0 > u\} \mid \{X_0 = x\}) = 0$ , and hence  $S_1$  is finite almost surely. By definition  $X_{S_1} \neq X_0 = x$ , and hence  $p_{xx} = 0$ .  $\square$

**Proposition 1.22.** *Consider a stable CTMC  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}_+}$ . Then for all states  $x, y \in \mathcal{X}$  and duration  $u \in \mathbb{R}_+$ ,*

$$P(\{T_{n+1} > u, Z_{n+1} = y\} \mid \{X_0 = x_0, \dots, Z_n = x, S_0 \leq s_0, \dots, S_n \leq s_n\}) = p_{xy} e^{-u\nu_x}.$$

*Proof.* The history of the process until stopping time  $S_n$  is given by  $\mathcal{F}_{S_n} = \sigma(S_0, (Z_0, T_1), \dots, (Z_{n-1}, T_n), Z_n)$ . Therefore  $H \triangleq \{S_0 \leq s_0\} \cap \bigcap_{i=1}^n \{Z_{i-1} = x_{i-1}, S_i \leq s_i\} \in \mathcal{F}_{S_n}$  and  $\{Z_n = x\} \cap H \in \mathcal{F}_{S_n}$ . Using strong Markov property and time-homogeneity of the CTMC  $X$ , we have

$$P(\{T_{n+1} > u, Z_{n+1} = y\} \mid \{Z_n = x\} \cap H) = P_x \{S_1 > u, Z_1 = y\}.$$

The result follows from the previous Lemma 1.19.  $\square$

**Corollary 1.23.** *For a time-homogeneous CTMC, the transition probabilities  $(p_{xy} : x, y \in \mathcal{X})$  and holding times  $T : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  are independent. The jump process  $Z : \Omega \rightarrow \mathcal{X}^{\mathbb{Z}_+}$  is a homogeneous Markov chain with countable state space  $\mathcal{X}$ . The holding time sequence  $T : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  is independent and  $T_n$  is distributed exponentially with rate  $\nu_{Z_{n-1}}$  for each  $n \in \mathbb{N}$ .*

**Example 1.24 (Poisson process).** For a Poisson counting process  $N : \Omega \rightarrow \mathbb{Z}_+^{\mathbb{R}^+}$  with time-homogeneous rate  $\lambda$ , the countable state space is  $\mathbb{Z}_+$ , and transition rate  $\nu_n = \lambda$  for each state  $n \in \mathbb{Z}_+$ . It follows from the memoryless property of exponential random variables, that

$$\bar{F}_n(t) = P(\{Y_u > t\} \mid \{N_u = n\}) = P\{S_1 > t\} = e^{-\lambda t}.$$

Further, the embedded Markov chain or the jump process is given by the initial state  $N_0 = 0$  and the transition probability matrix  $P = (p_{n,m} : n, m \in \mathbb{Z}_+)$  where  $p_{n,n+1} = 1$  and  $p_{n,m} = 0$  for  $m \neq n + 1$ . This follows from the definition of  $T_1$ , since  $p_{n,m} = P(\{N_{T_1} = m\} \mid \{N_0 = n\}) = \mathbb{1}_{\{m=n+1\}}$ .

**Theorem 1.25.** *A pure-jump homogeneous CTMC whose embedded DTMC is recurrent is regular.*

*Proof.* Let  $X_0 = x \in \mathcal{X}$  be the initial state. Let  $N_x(n) = \sum_{k=1}^n \mathbb{1}_{\{Z_k=x\}}$  be the number of visits to a state  $x \in \mathcal{X}$  in the first  $n$  transitions and  $T_i^x$  be the  $i$ th sojourn time in the state  $x$ . From the recurrence of the embedded chain, the state  $x$  occurs infinitely often, i.e.  $\lim_{n \in \mathbb{N}} N_x(n) = \infty$  almost surely. It follows that the sojourn time sequence  $T^x : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  is *i.i.d.* and exponentially distributed with mean  $\mathbb{E}T_i^x = \frac{1}{\nu_x} < \infty$ . Since  $S_n \geq \sum_{i=1}^{N_x(n)} T_i^x$ , we get that

$$m_t = \sum_{n \in \mathbb{N}} P\{S_n \leq t\} \leq \sum_{n \in \mathbb{N}} P\left\{\sum_{i=1}^{N_x(n)} T_i^x \leq t\right\} = \nu_x t.$$

It follows that  $N_t$  is almost surely finite for any finite time  $t \in \mathbb{R}_+$ .  $\square$

## A Exponential random variables

**Lemma A.1.** *Let  $X$  be an exponential random variable, and  $S$  be any positive random variable, independent of  $X$ . Then, for all  $u \geq 0$*

$$P(\{X > S + u\} \mid \{X > S\}) = P\{X > u\}.$$

*Proof.* Let the distribution of  $S$  be  $F$  and  $X$  be memoryless with rate  $\mu$ . From the definition of conditional probability, we can write  $P(\{X > S + u\} \mid \{X > S\}) = \frac{P\{X > S + u\}}{P\{X > S\}}$ . Since a probability is an expectation of an indicator, we can write

$$P\{X > S + u\} = \mathbb{E}[\mathbb{E}[\mathbb{1}_{\{X > S + u\}} \mid \sigma(S)]] = \mathbb{E}[e^{-\mu(S+u)}] = e^{-\mu u} \mathbb{E}[e^{-\mu S}].$$

It follows that  $P(\{X > S + u\} \mid \{X > S\}) = P\{X > u\} = e^{-\mu u}$  for all  $u \in \mathbb{R}_+$ .  $\square$

## B Coupling

For a homogeneous regular and stable CTMC  $X : \Omega \rightarrow \mathcal{X}^{\mathbb{R}^+}$ , we denote the embedded Markov chain by  $Z : \Omega \rightarrow \mathcal{X}^{\mathbb{Z}^+}$  and the independent inter-jump time sequence by  $T : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  where  $T_n$  is an exponential random variable with rate  $\nu_{Z_{n-1}}$  for all  $n \in \mathbb{N}$ . From the regularity and stability of process  $X$ , we have

$$0 < \bar{\nu} \leq \inf_{x \in \mathcal{X}} \nu_x \leq \sup_{x \in \mathcal{X}} \nu_x \leq \underline{\nu} < \infty.$$

Consider an *i.i.d.* uniform random sequence  $U : \Omega \rightarrow [0, 1]^{\mathbb{N}}$  and define three dependent random sequences  $\underline{T}, T, \bar{T} : \Omega \rightarrow \mathbb{R}_+^{\mathbb{N}}$  such that for each  $n \in \mathbb{N}$ , we have

$$\bar{T}_n \triangleq -\frac{1}{\bar{\nu}} \log U_n, \quad \underline{T}_n \triangleq -\frac{1}{\underline{\nu}} \log U_n, \quad T_n \triangleq -\frac{1}{\nu_{Z_{n-1}}} \log U_n.$$

We observe that  $\underline{T}$  and  $\bar{T}$  are *i.i.d.* exponential random sequences with rates  $\underline{\nu}$  and  $\bar{\nu}$  respectively. Further,  $T$  is an independent exponential random sequence with the rate  $\nu_{Z_{n-1}}$  for  $T_n$ . In addition, by construction, we have  $\underline{T}_n \leq T_n \leq \bar{T}_n$  for each  $n \in \mathbb{N}$ .